Research Article

A Generalization of the Havrda-Charvat and Tsallis Entropy and Its Axiomatic Characterization

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In this communication, we characterize a measure of information of types \(\alpha, \beta, \gamma\) by taking certain axioms parallel to those considered earlier by Havrda and Charvat along with the recursive relation

\[
H_n(p_1, \ldots, p_n; \alpha, \beta, \gamma) = H_{n-1}(p_1 + p_2, p_3, \ldots, p_n; \alpha, \beta, \gamma) + (p_1 + p_2)(p_1 + p_2; \alpha, \beta, \gamma), \quad \alpha \neq \gamma \neq \beta, \quad \alpha, \beta, \gamma > 0.
\]

Some properties of this measure are also studied. This measure includes Shannon’s information measure as a special case.

1. Introduction

Shannon’s measure of entropy for a discrete probability distribution

\[
P = (p_1, \ldots, p_n), \quad p_i \geq 0, \quad \sum_{i=1}^{n} p_i = 1,
\]

given by

\[
H(P) = -\sum_{i=1}^{n} p_i \log p_i,
\]

has been characterized in several ways (see Aczéľ and Daróczy [1]). Out of the many ways of characterization the two elegant approaches are to be found in the work of (i) Faddeev [2], who uses branching property namely,

\[
H_n(p_1, \ldots, p_n) = H_{n-1}(p_1 + p_2, p_3, \ldots, p_n)
\]

\[
+ (p_1 + p_2)H_2\left(\frac{p_1}{p_1 + p_2}, \frac{p_2}{p_1 + p_2}\right),
\]

\(n = 3, 4, \ldots\) for the above distribution \(P\), as the basic postulate, and (ii) Chaundy and McLeod [3], who studied the functional equation

\[
\sum_{i=1}^{n} \sum_{j=1}^{m} f(p_i q_j) = \sum_{i=1}^{n} f(p_i) + \sum_{j=1}^{m} f(q_j),
\]

for \(p_i \geq 0, q_j \geq 0\).

Both of the above-mentioned approaches have been extensively exploited and generalized. The most general form of (4) has been studied by Sharma and Taneja [4], who considered the functional equation

\[
\sum_{i=1}^{n} \sum_{j=1}^{m} f(p_i q_j) = \sum_{i=1}^{n} \sum_{j=1}^{m} f(p_i) g(q_j)
\]

\[
+ \sum_{i=1}^{n} \sum_{j=1}^{m} g(p_i) f(q_j),
\]

\(\sum_{j=1}^{m} q_j = 1, \quad p_i \geq 0, q_j \geq 0\).
We define the information measure as

\[ H_n(p_1, \ldots, p_n; \alpha, \beta, \gamma) = \left( 2^{\frac{\gamma - \beta}{\gamma}} - 2^{\frac{\gamma - \beta}{\gamma}} \right)^{-1} \sum_{i=1}^{n} \left( p_i^{\alpha/\gamma} - p_i^{\beta/\gamma} \right), \quad (6) \]

for a complete probability distribution \( P = (p_1, \ldots, p_n), \ p_i \geq 0, \ \sum_{i=1}^{n} p_i = 1. \)

Measure (6) reduces to entropy of type \( \beta \) (or \( \alpha \)) when \( \alpha = \gamma = 1 \) (or \( \beta = \gamma = 1 \)) given by

\[ H_n(p_1, \ldots, p_n; \beta) = \left( 2^{1 - \beta} - 1 \right) \left( \sum_{i=1}^{n} p_i^\beta - 1 \right), \quad (7) \]

\[ \beta \neq 1, \ \beta > 0. \]

When \( \beta \to 1 \), measure (7) reduces to Shannon’s entropy [5], namely,

\[ H_n(p_1, \ldots, p_2) = -\sum_{i=1}^{n} p_i \log_2 p_i, \quad (8) \]

The measure (7) was characterized by many authors by different approaches. Havrda and Charvát [6] characterized (7) by an axiomatic approach. Daróczy [7] studied (7) by a functional equation. A joint characterization of the measures (7) and (8) has been done by author in two different ways. Firstly by a generalized functional equation having four different functions and secondly by an axiomatic approach. Later on Tsallis [8] gave the applications of (7) in Physics.

To characterize strongly interacting statistical systems within a thermodynamical framework—complex system in particular—it might be necessary to introduce generalized entropies. A series of such entropies have been proposed in the past, mainly to accommodate important empirical distribution functions to a maximum ignorance principle. The understanding of the fundamental origin of these entropies and its deeper relations to complex systems is limited. Authors [9] explore this question from first principle. Authors [9] observed that the 4th Khinchin axiom is violated by strongly interacting system in general and by assuming the first three Khinchin axioms derived a unique entropy and also classified the known entropies with equivalence classes.

For statistical system that violates the four Shannon-Khinchin axioms, entropy takes a more general form than the Boltzmann-Gibbs entropy. The framework of superstatistics allows one to formulate a maximum entropy principle with these generalized entropies, making them useful for understanding distribution functions of non-Markovian or nonergodic complex systems. For such systems where the composability axiom is violated there exist only two ways to implement the maximum entropy principle; one is using the escort probabilities and the other is not. The two ways are connected through a duality. Authors [10] showed that this duality fixes a unique escort probability and derived a complete theory of the generalized logarithms and also gave the relationship between the functional forms of generalized logarithms and the asymptotic scaling behavior of the entropy.

Suyari [11] has proposed a generalization of Shannon-Khinchin axioms, which determines a class of entropies containing the well-known Tsallis and Havrda-Charvat entropies. Authors [12] showed that the class of entropy functions determined by Suyari’s axioms is wider than the one proposed by Suyari and generalized Suyari’s axioms characterizing recently introduced class of entropies obtained by averaging pseudoadditive information content.

In this communication, we characterized the measure (6) by taking certain axioms parallel to those considered earlier by Havrda and Charvát [6] along with the recursive relation (9). Some properties of this measure are also studied.

The measure (6) satisfies a recursive relation as follows:

\[ H_n(p_1, \ldots, p_n; \alpha, \beta, \gamma) = H_{n-1}(p_1 + p_2, \ldots, p_n; \alpha, \beta, \gamma) + A_{(\alpha, \gamma)}(\alpha, \gamma) H_n(p_1, \ldots, p_n; \alpha, \beta, \gamma), \]

\[ + A_{(\beta, \gamma)}(\beta, \gamma) H_n(p_1, \ldots, p_n; \alpha, \beta, \gamma), \]

\[ \alpha \neq \gamma \neq \beta, \ \alpha, \beta, \gamma > 0, \quad (9) \]

where \( p_1 + p_2 > 0, \ A_{(\alpha, \gamma)} = (2^{\gamma - \alpha})^{-1}, \) and \( A_{(\beta, \gamma)} = (2^{\gamma - \beta})^{-1}. \)

Consider

\[ H(p_1, p_2, \ldots, p_n; \alpha, \gamma) = A_{(\alpha, \gamma)}^{-1} \left[ \sum_{i=1}^{n} p_i^{\alpha/\gamma} - 1 \right], \]

\[ \alpha \neq \gamma, \ \alpha, \gamma > 0 \neq 1, \quad (10) \]

\[ H(p_1, p_2, \ldots, p_n; \beta, \gamma) = A_{(\beta, \gamma)}^{-1} \left[ 1 - \sum_{i=1}^{n} p_i^{\beta/\gamma} \right], \]

\[ \beta \neq \gamma, \ \beta, \gamma > 0 \neq 1. \]

**Proof.**

\[ H_n(p_1, \ldots, p_n; \alpha, \beta, \gamma) - H_{n-1}(p_1 + p_2, \ldots, p_n; \alpha, \beta, \gamma) = \left( 2^{\frac{\gamma - \beta}{\gamma}} - 2^{\frac{\gamma - \beta}{\gamma}} \right)^{-1} \left[ \left( p_1^{\alpha/\gamma} - p_1^{\beta/\gamma} \right) + \left( p_2^{\alpha/\gamma} - p_2^{\beta/\gamma} \right) + \ldots \right] + \left( p_n^{\alpha/\gamma} - p_n^{\beta/\gamma} \right) \]

\[ - \left[ \left( 2^{\frac{\gamma - \alpha}{\gamma}} - 2^{\frac{\gamma - \alpha}{\gamma}} \right)^{-1} \right] \left[ \left( p_1 + p_2 \right)^{\alpha/\gamma} - \left( p_1 + p_2 \right)^{\beta/\gamma} + \left( p_3^{\alpha/\gamma} - p_3^{\beta/\gamma} \right) + \ldots \right] + \left( p_n^{\alpha/\gamma} - p_n^{\beta/\gamma} \right) \]
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\[ = (2^{(\gamma - \alpha)/\gamma} - 2^{(\gamma - \beta)/\gamma})^{-1} \left\{ p_1^{\alpha/\gamma} - p_1^{\beta/\gamma} + p_2^{\alpha/\gamma} - p_2^{\beta/\gamma} 
\right\} - (p_1 + p_2)^{\alpha/\gamma} + (p_1 + p_2)^{\beta/\gamma} \]

\[ = (2^{(\gamma - \alpha)/\gamma} - 2^{(\gamma - \beta)/\gamma})^{-1} \left\{ p_1^{\alpha/\gamma} + p_2^{\alpha/\gamma} - (p_1 + p_2)^{\alpha/\gamma} \right\} \]

\[ + (2^{(\gamma - \alpha)/\gamma} - 2^{(\gamma - \beta)/\gamma})^{-1} \left\{ (p_1 + p_2)^{\beta/\gamma} - p_1^{\beta/\gamma} - p_2^{\beta/\gamma} \right\} \]

\[ = (2^{(\gamma - \alpha)/\gamma} - 2^{(\gamma - \beta)/\gamma})^{-1} (p_1 + p_2)^{\alpha/\gamma} \times \left[ \frac{p_1^{\alpha/\gamma}}{(p_1 + p_2)^{\alpha/\gamma}} + \frac{p_2^{\alpha/\gamma}}{(p_1 + p_2)^{\alpha/\gamma}} - 1 \right] \]

\[ + (2^{(\gamma - \alpha)/\gamma} - 2^{(\gamma - \beta)/\gamma})^{-1} (p_1 + p_2)^{\beta/\gamma} \times \left[ 1 - \frac{p_1^{\beta/\gamma}}{(p_1 + p_2)^{\alpha/\gamma}} - \frac{p_2^{\beta/\gamma}}{(p_1 + p_2)^{\alpha/\gamma}} \right] \]

\[ = \frac{A_{(\alpha,\gamma)}}{A_{(\alpha,\gamma)} - A_{(\beta,\gamma)}} (p_1 + p_2)^{\alpha/\gamma} H_2 \left( \frac{p_1}{p_1 + p_2}, \frac{p_2}{p_1 + p_2}; \alpha, \gamma \right) \]

\[ + \frac{A_{(\beta,\gamma)}}{A_{(\beta,\gamma)} - A_{(\alpha,\gamma)}} (p_1 + p_2)^{\beta/\gamma} \times H_2 \left( \frac{p_1}{p_1 + p_2}, \frac{p_2}{p_1 + p_2}; \gamma, \beta \right), \]

(11)

which proves (9). □

2. Set of Axioms

For characterizing a measure of information of types \(\alpha, \beta,\) and \(\gamma\) associated with a probability distribution \(P = (p_1, \ldots, p_n),\) \(p_i \geq 0, \sum_{i=1}^n p_i = 1,\) we introduce the following axioms:

(1) \(H_n(p_1, \ldots, p_n; \alpha, \beta, \gamma)\) is continuous in the region

\[ p_i \geq 0, \sum_{i=1}^n p_i = 1, \alpha, \beta, \gamma > 0; \]

(12)

(2) \(H_2(1, 0; \alpha, \beta, \gamma) = 0;\)

(3) \(H_2(1/2, 1/2; \alpha, \beta, \gamma) = 1, \alpha, \beta, \gamma > 0;\)

(4)

\[ H_n(p_1, \ldots, p_i, 0, p_{i+1}, \ldots, p_n; \alpha, \beta, \gamma) = H_{n-1}(p_1, \ldots, p_{i-1}, p_{i+1}, \ldots, p_n; \alpha, \beta, \gamma), \]

for every \(i = 1, 2, \ldots, n;\)

\[ H_{n+1}(p_1, \ldots, p_{i-1}, v_i, v_{i+1}, \ldots, p_n; \alpha, \beta, \gamma) \]

\[ = H_n(p_1, \ldots, p_{i-1}, 0, p_{i+1}, \ldots, p_n; \alpha, \beta, \gamma) \]

\[ \times H_{n-2}(v_i, v_{i+1}, \ldots, p_n; \alpha, \beta, \gamma), \]

for every \(i = 1, 2, \ldots, n;\)

\[ H_{n+1}(p_1, \ldots, p_{i-1}, v_i, v_{i+1}, \ldots, p_n; \alpha, \beta, \gamma) \]

\[ = H_n(p_1, \ldots, p_{i-1}, v_i, v_{i+1}, \ldots, p_n; \alpha, \beta, \gamma) \]

\[ \times H_{n-2}(p_1, \ldots, p_{i-1}, v_i, v_{i+1}, \ldots, p_n; \alpha, \beta, \gamma), \]

(13)

\[ (5) \]

\[ H_{n+1}(p_1, \ldots, p_{i-1}, v_i, v_{i+1}, \ldots, p_n; \alpha, \beta, \gamma) \]

\[ = H_n(p_1, \ldots, p_{i-1}, v_i, v_{i+1}, \ldots, p_n; \alpha, \beta, \gamma) \]

\[ \times H_{n-2}(p_1, \ldots, p_{i-1}, v_i, v_{i+1}, \ldots, p_n; \alpha, \beta, \gamma), \]

(14)

\[ \alpha \neq \gamma \neq \beta, \alpha, \beta, \gamma > 0, \]

for every \(v_i + v_{i+1} = p_i > 0, i = 1, 2, \ldots, n,\) where \(A_{(\alpha,\gamma)} = (2^{(\gamma - \alpha)/\gamma} - 1)\) and \(A_{(\beta,\gamma)} = (2^{(\gamma - \beta)/\gamma} - 1), \alpha \neq \gamma \neq \beta.\)

Theorem 1. If \(\alpha \neq \gamma \neq \beta; \alpha, \beta, \gamma > 0,\) then the axioms (1)–(5)

determine a measure given by

\[ H_n(p_1, \ldots, p_n; \alpha, \beta, \gamma) \]

\[ = (A_{(\alpha,\gamma)} - A_{(\beta,\gamma)})^{-1} \sum_{i=1}^n \left( p_i^{\alpha/\gamma} - p_i^{\beta/\gamma} \right), \]

(15)

\[ \alpha \neq \gamma \neq \beta, \alpha, \beta, \gamma > 0, \]

where \(A_{(\alpha,\gamma)} = (2^{(\gamma - \alpha)/\gamma} - 1)\) and \(A_{(\beta,\gamma)} = (2^{(\gamma - \beta)/\gamma} - 1).\)

Before proving the theorem we prove some intermediate results based on the above axioms.

Lemma 2. If \(v_k \geq 0, k = 1, 2, \ldots, m\) and \(\sum_{k=1}^m v_k = p_i > 0,\) then

\[ H_{n+m-1}(p_1, \ldots, p_{i-1}, v_1, \ldots, v_m, p_{i+1}, \ldots, p_n; \alpha, \beta, \gamma) \]

\[ = H_n(p_1, \ldots, p_{i-1}, p_i, \ldots, p_n; \alpha, \beta, \gamma) \]

\[ + \frac{A_{(\alpha,\gamma)}}{A_{(\alpha,\gamma)} - A_{(\beta,\gamma)}} p_i^{\alpha/\gamma} H_m \left( \frac{v_1}{p_i}, \ldots, \frac{v_m}{p_i}; \alpha, \gamma \right) \]

\[ + \frac{A_{(\beta,\gamma)}}{A_{(\beta,\gamma)} - A_{(\alpha,\gamma)}} p_i^{\beta/\gamma} H_m \left( \frac{v_1}{p_i}, \ldots, \frac{v_m}{p_i}; \gamma, \beta \right). \]

(16)

Proof. To prove the lemma, we proceed by induction. For \(m = 2,\) the desired statement holds (cf. axiom (4)). Let us suppose
that the result is true for numbers less than or equal to \( m \), we will prove it for \( m + 1 \). We have

\[
H_{m+1} \left( p_1, \ldots, p_{i-1}, v_1, \ldots, v_{m+1}, p_{i+1}, \ldots, p_n; \alpha, \beta, \gamma \right) = H_n \left( p_1, \ldots, p_{i-1}, v_1, L, p_{i+1}, \ldots, p_n; \alpha, \beta, \gamma \right) + \frac{A(\alpha,\gamma)}{A(\alpha,\gamma) - A(\beta,\gamma)} L^{\alpha/\gamma} H_m \left( \frac{v_2}{L}, \ldots, \frac{v_{m+1}}{L}; \alpha, \gamma \right) + \frac{A(\beta,\gamma)}{A(\beta,\gamma) - A(\alpha,\gamma)} L^{\beta/\gamma} H_m \left( \frac{v_2}{L}, \ldots, \frac{v_{m+1}}{L}; \gamma, \beta \right).
\]

(21)

where \( L = v_2 + \cdots + v_{m+1} \).

For \( \beta = \gamma \), (18) reduces to

\[
H_{m+1} \left( \frac{v_1}{p_i}, \ldots, \frac{v_{m+1}}{p_i}; \alpha, \gamma \right) = H_2 \left( \frac{v_1}{p_i}, L; \alpha, \gamma \right) + \left( \frac{L}{p_i} \right)^{\alpha/\gamma} H_m \left( \frac{v_2}{L}, \ldots, \frac{v_{m+1}}{L}; \alpha, \gamma \right).
\]

(19)

Similarly for \( \alpha = \gamma \), (18) reduces to

\[
H_{m+1} \left( \frac{v_1}{p_i}, \ldots, \frac{v_{m+1}}{p_i}; \alpha, \beta \right) = H_2 \left( \frac{v_1}{p_i}, \frac{L}{p_i}; \gamma, \beta \right) + \left( \frac{L}{p_i} \right)^{\beta/\gamma} H_m \left( \frac{v_2}{L}, \ldots, \frac{v_{m+1}}{L}; \gamma, \beta \right).
\]

(20)

Expression (17) together with (19) and (20) gives the desired result.

\[\square\]

**Lemma 3.** If \( v_{ij} \geq 0 \), \( j = 1, 2, \ldots, m_j, \sum_{j=1}^{m_j} v_{ij} = p_i > 0 \), \( i = 1, 2, \ldots, n \), and \( \sum_{i=1}^{n} p_i = 1 \), then

\[
H_{m_1, \ldots, m_n} \left( v_{11}, v_{12}, \ldots, v_{1m_1}; \ldots; v_{n1}, v_{n2}, \ldots, v_{nm_n}; \alpha, \beta, \gamma \right) = H_n \left( p_1, p_2, \ldots, p_n; \alpha, \beta, \gamma \right) + \frac{A(\alpha,\gamma)}{A(\alpha,\gamma) - A(\beta,\gamma)} \sum_{i=1}^{n} p_i^{\alpha/\gamma} H_m \left( \frac{v_{1i}}{p_i}, \ldots, \frac{v_{imi}}{p_i}; \alpha, \gamma \right) + \frac{A(\beta,\gamma)}{A(\beta,\gamma) - A(\alpha,\gamma)} \sum_{i=1}^{n} p_i^{\beta/\gamma} H_m \left( \frac{v_{1i}}{p_i}, \ldots, \frac{v_{imi}}{p_i}; \gamma, \beta \right).
\]

(21)

**Proof.** Proof of this lemma directly follows from Lemma 2. \[\square\]

**Lemma 4.** If \( F(\alpha, \beta, \gamma) = H_n(1/n, \ldots, 1/n; \alpha, \beta, \gamma) \), then

\[
F(\alpha, \beta, \gamma) = \frac{A(\alpha,\gamma)}{A(\alpha,\gamma) - A(\beta,\gamma)} F(\alpha, \alpha, \gamma) + \frac{A(\beta,\gamma)}{A(\beta,\gamma) - A(\alpha,\gamma)} F(\alpha, \gamma, \beta),
\]

(22)

where \( F(\alpha, \gamma) = A^{-1}(\alpha,\gamma)(n^{(\gamma-\alpha)/\gamma} - 1), \alpha \neq \gamma, \) and

\[
F(\gamma, \beta) = A^{-1}(\beta,\gamma)(n^{(\gamma-\beta)/\gamma} - 1), \beta \neq \gamma.
\]

(23)
Proof. Replacing in Lemma 3 $m_i$ by $m$ and putting $v_{ij} = 1/mn$, $i = 1, 2, n$, $j = 1, 2, m$, where $m$ and $n$ are positive integer, we have

$$F(mn; \alpha, \beta, \gamma) = F(m; \alpha, \beta, \gamma) + \frac{A_{(\alpha, \gamma)} - A_{(\beta, \gamma)}}{A_{(\alpha, \gamma)} - A_{(\beta, \gamma)}} \left( \frac{1}{m} \right)^{(\alpha - \gamma)/\gamma} F(m; \alpha, \gamma)$$

$$+ \frac{A_{(\beta, \gamma)} - A_{(\alpha, \gamma)}}{A_{(\beta, \gamma)} - A_{(\alpha, \gamma)}} \left( \frac{1}{n} \right)^{(\beta - \gamma)/\gamma} F(m; \gamma, \beta)$$

Putting $m = 1$ in (24) and using $F(1; \alpha, \beta, \gamma) = 0$ (by axiom (2)), we get

$$F(n; \alpha, \beta, \gamma) = \frac{A_{(\alpha, \gamma)} - A_{(\beta, \gamma)}}{A_{(\alpha, \gamma)} - A_{(\beta, \gamma)}} F(n; \alpha, \gamma) + \frac{A_{(\beta, \gamma)} - A_{(\alpha, \gamma)}}{A_{(\beta, \gamma)} - A_{(\alpha, \gamma)}} F(n; \gamma, \beta).$$

(25)

which is (22).

Comparing the right hand sides of (24) and (25), we get

$$F(m; \alpha, \beta, \gamma) + \frac{A_{(\alpha, \gamma)} - A_{(\beta, \gamma)}}{A_{(\alpha, \gamma)} - A_{(\beta, \gamma)}} \left( \frac{1}{m} \right)^{(\alpha - \gamma)/\gamma} F(m; \alpha, \gamma)$$

$$+ \frac{A_{(\beta, \gamma)} - A_{(\alpha, \gamma)}}{A_{(\beta, \gamma)} - A_{(\alpha, \gamma)}} \left( \frac{1}{n} \right)^{(\beta - \gamma)/\gamma} F(m; \gamma, \beta)$$

$$= F(n; \alpha, \beta, \gamma) + \frac{A_{(\alpha, \gamma)} - A_{(\beta, \gamma)}}{A_{(\alpha, \gamma)} - A_{(\beta, \gamma)}} \left( \frac{1}{n} \right)^{(\alpha - \gamma)/\gamma} F(m; \alpha, \gamma)$$

$$+ \frac{A_{(\beta, \gamma)} - A_{(\alpha, \gamma)}}{A_{(\beta, \gamma)} - A_{(\alpha, \gamma)}} \left( \frac{1}{n} \right)^{(\beta - \gamma)/\gamma} F(m; \gamma, \beta).$$

(27)

Equation (27) together with (22) gives

$$A_{(\alpha, \gamma)} \left[ 1 - \left( \frac{1}{n} \right)^{(\alpha - \gamma)/\gamma} - 1 \right] F(m; \alpha, \gamma)$$

$$+ \left[ \left( \frac{1}{m} \right)^{(\alpha - \gamma)/\gamma} - 1 \right] F(n; \alpha, \gamma)$$

$$= A_{(\beta, \gamma)} \left[ 1 - \left( \frac{1}{n} \right)^{(\beta - \gamma)/\gamma} - 1 \right] F(m; \gamma, \beta)$$

$$+ \left[ \left( \frac{1}{m} \right)^{(\beta - \gamma)/\gamma} - 1 \right] F(n; \gamma, \beta).$$

(28)

Putting $n = 2$ in (28) and using $F(2; \alpha, \beta, \gamma) = H_2(1/2, 1/2; \alpha, \beta, \gamma) = 1$, we get

$$A_{(\alpha, \gamma)} \left[ 1 - 2^{1-\alpha/\gamma} \right] F(m; \alpha, \gamma) - \left( 1 - \left( \frac{1}{m} \right)^{\alpha/\gamma-1} \right) = A_{(\beta, \gamma)} \left[ 1 - 2^{1-\beta/\gamma} \right] F(m; \gamma, \beta) - \left( 1 - \left( \frac{1}{m} \right)^{\beta/\gamma-1} \right)$$

$$= C \quad \text{(say)}.$$

That is, $A_{(\alpha, \gamma)} \{ 1 - 2^{1-\alpha/\gamma} F(m; \alpha, \gamma) - \left( 1 - (1/m)^{\alpha/\gamma-1} \right) \} = C$, where $C$ is an arbitrary constant.

For $m = 1$, we get $C = 0$.

Thus, we have

$$F(m; \alpha, \gamma) = \frac{1 - m^{1-\alpha/\gamma}}{1 - 2^{1-\alpha/\gamma}} = A_{(\gamma, \alpha)} \left( m^{1-\alpha/\gamma} - 1 \right), \quad \alpha \neq \gamma.$$

(30)

Similarly,

$$F(m; \gamma, \beta) = \frac{1 - m^{1-\beta/\gamma}}{1 - 2^{1-\beta/\gamma}} = A_{(\gamma, \beta)} \left( m^{1-\beta/\gamma} - 1 \right), \quad \beta \neq \gamma.$$

(31)

which is (23).

Now (22) together with (23) gives

$$F(n; \alpha, \beta, \gamma) = \frac{A_{(\alpha, \gamma)} - A_{(\beta, \gamma)}}{A_{(\alpha, \gamma)} - A_{(\beta, \gamma)}} F(n; \alpha, \gamma)$$

$$+ \frac{A_{(\beta, \gamma)} - A_{(\alpha, \gamma)}}{A_{(\beta, \gamma)} - A_{(\alpha, \gamma)}} F(n; \gamma, \beta).$$

(32)

Proof of the Theorem. We prove the theorem for rationals and then the continuity axiom (1) extends the result for reals. For this let $m$ and $r_i$'s be positive integers such that $\sum_{i=1}^{n} r_i = m$ and if we put $p_i = r_i/m, i = 1, 2, \ldots, n$ then an application of Lemma 3 gives

$$H_m \left( \frac{1}{m}, \ldots, \frac{1}{m}, \ldots, \frac{1}{m}; \alpha, \beta, \gamma \right)$$

$$= H_n \left( p_1, p_2, \ldots, p_n; \alpha, \beta, \gamma \right)$$

$$+ \frac{A_{(\alpha, \gamma)} - A_{(\beta, \gamma)}}{A_{(\alpha, \gamma)} - A_{(\beta, \gamma)}} \sum_{i=1}^{n} p_i^{\beta/\gamma} H_{r_i} \left( \frac{1}{r_i}, \ldots, \frac{1}{r_i}; \alpha, \beta, \gamma \right)$$

$$+ \frac{A_{(\beta, \gamma)} - A_{(\alpha, \gamma)}}{A_{(\beta, \gamma)} - A_{(\alpha, \gamma)}} \sum_{i=1}^{n} p_i^{\alpha/\gamma} H_{r_i} \left( \frac{1}{r_i}, \ldots, \frac{1}{r_i}; \alpha, \beta, \gamma \right).$$

(33)
That is,

\[ H_n(p_1, \ldots, p_n; \alpha, \beta, \gamma) = F(m; \alpha, \beta, \gamma) - A(\alpha, \gamma) \frac{1}{A(\alpha, \gamma) - A(\beta, \gamma)} \sum_{i=1}^{n} p_i^{\alpha/\gamma} F(r_i; \alpha, \gamma) \]

Equation (34) together with (23) and (32) gives

\[ H_n(p_1, \ldots, p_n; \alpha, \beta, \gamma) = \frac{1}{A(\alpha, \gamma) - A(\beta, \gamma)} \sum_{i=1}^{n} (p_i^{\alpha/\gamma} - p_i^{\beta/\gamma}), \]

\[ \alpha \neq \gamma \neq \beta, \alpha, \beta, \gamma > 0. \]

(35)

which is (15).

This completes the proof of the theorem.

3. Properties of Entropy of Types $\alpha$, $\beta$, and $\gamma$

The measure $H_n(P; \alpha, \beta, \gamma)$, where $P = (p_1, \ldots, p_n)$, $p_i \geq 0$, $\sum_{i=1}^{n} p_i = 1$, is a probability distribution, as characterized in the preceding section and satisfies certain properties, which are given in the following theorems:

**Theorem 5.** The measure $H_n(P; \alpha, \beta, \gamma)$ is nonnegative for $\alpha \neq \gamma \neq \beta$, $\alpha, \beta, \gamma > 0$.

**Proof.**

Case 1. $\alpha > \gamma$; $\beta < \gamma \Rightarrow \alpha/\gamma > 1$, $\beta/\lambda < 1$;

\[ \sum_{i=1}^{n} p_i^{\alpha/\gamma} < 1, \quad \sum_{i=1}^{n} p_i^{\beta/\gamma} > 1, \]

\[ \Rightarrow \sum_{i=1}^{n} (p_i^{\alpha/\gamma} - p_i^{\beta/\gamma}) < 0. \]

(36)

Since, $\alpha > \gamma$ and $\beta < \gamma$, we get

\[ (2^{1-\alpha/\gamma} - 2^{1-\beta/\gamma})^{-1} \sum_{i=1}^{n} (p_i^{\alpha/\gamma} - p_i^{\beta/\gamma}) > 0. \]

(37)

Case 2. Similarly for $\alpha < \gamma$ and $\beta > \gamma$, we get

\[ (2^{1-\alpha/\gamma} - 2^{1-\beta/\gamma})^{-1} \sum_{i=1}^{n} (p_i^{\alpha/\gamma} - p_i^{\beta/\gamma}) > 0. \]

(38)

Therefore from Case 1, Case 2, and axiom (2), we get

\[ H_n(P; \alpha, \beta, \gamma) \geq 0. \]

(39)

This completes the proof of theorem.

**Definition 6.** We will use the following definition of a convex function.

A function $f(\cdot)$ over the points in a convex set $R$ is convex if for all $r_1, r_2 \in R$ and $\mu \in (0, 1)$

\[ \mu f(r_1) + (1 - \mu) f(r_2) \leq f(\mu r_1 + (1 - \mu) r_2). \]

(40)

The function $f(\cdot)$ is convex if (40) holds with $\geq$ in place of $\leq$.

**Theorem 7.** The measure $H_n(P; \alpha, \beta, \gamma)$ is convex function of the probability distribution $P = (p_1, \ldots, p_n)$, $p_i \geq 0$, $\sum_{i=1}^{n} p_i = 1$, when either $\alpha > \gamma$ and $\beta \leq \gamma$ or $\beta > \gamma$ and $\alpha \leq \gamma$.

**Proof.** Let there be $r$ distributions

\[ P_k(X) = \{p_k(x_1), \ldots, p_k(x_n)\}, \quad \sum_{k=1}^{r} p_k(x_i) = 1, \]

\[ k = 1, 2, \ldots, r, \]

associated with the random variable $X = (x_1, \ldots, x_n)$.

Consider $r$ numbers $(a_1, \ldots, a_r)$ such that $a_k \geq 0$ and $\sum_{k=1}^{r} a_k = 1$ and define

\[ P_o(X) = \{p_o(x_1), \ldots, p_o(x_n)\}, \]

where

\[ p_o(x_i) = \sum_{k=1}^{r} a_k p_k(x_i), \quad i = 1, 2, \ldots, n. \]

(43)

Obviously, $\sum_{i=1}^{n} p_o(x_i) = 1$ and thus $P_o(x)$ is a bonafide distribution of $X$.

Let $\alpha > \gamma$ and $0 < \beta \leq \gamma$, then we have

\[ \sum_{k=1}^{r} a_k H_n(p_k; \alpha, \beta, \gamma) - H_n(P_o(\alpha, \beta, \gamma)) \]

\[ = \sum_{k=1}^{r} a_k H_n(p_k; \alpha, \beta, \gamma) - (A(\alpha, \gamma) - A(\beta, \gamma))^{-1} \left\{ \left[ \sum_{j=1}^{r} a_j p_j \right]^{\alpha/\gamma} - \left[ \sum_{j=1}^{r} a_j p_j \right]^{\beta/\gamma} \right\} \]

\[ \leq \sum_{k=1}^{r} a_k H_n(p_k; \alpha, \beta, \gamma) - \left( A(\alpha, \gamma) - A(\beta, \gamma) \right)^{-1} \left( \sum_{j=1}^{r} a_j p_j^{\alpha/\gamma} - \sum_{j=1}^{r} a_j p_j^{\beta/\gamma} \right) = 0, \]

(by Jensen’s inequality).

(44)

\[ \Rightarrow \sum_{k=1}^{r} a_k H_n(p_k; \alpha, \beta, \gamma) - H_n(P_o(\alpha, \beta, \gamma)) \leq 0, \]

that is, $\sum_{k=1}^{r} a_k H_n(p_k; \alpha, \beta, \gamma) \leq H_n(p_o; \alpha, \beta, \gamma)$, for $\alpha > \gamma$, $0 < \beta \leq \gamma$.

By symmetry in $\alpha, \beta,$ and $\gamma$ the above result is true for $\beta > \gamma$ and $0 < \alpha \leq \gamma$. 

□
Theorem 8. The measure \( H_n(p, \alpha, \beta, \gamma) \) satisfies the following relations:

(i) Generalized-Additive:

\[
H_{nm} (P \ast Q; \alpha, \beta, \gamma) = G_n (P; \alpha, \beta, \gamma) H_m (Q; \alpha, \beta, \gamma) + G_m (Q; \alpha, \beta, \gamma) H_n (P; \alpha, \beta, \gamma),
\]

\[\alpha, \beta, \gamma > 0,\]

(45)

where

\[
G_n (P; \alpha, \beta, \gamma) = \frac{1}{n} \sum_{i=1}^{n} (p_i^{\alpha/\gamma} + p_i^{\beta/\gamma}),
\]

(46)

\[\alpha, \beta, \gamma > 0.\]

(ii) Subadditive: for \( \alpha, \beta > \gamma \), the measure \( H_n(p; \alpha, \beta, \gamma) \) is subadditive; that is,

\[
H_{nm} (P \ast Q; \alpha, \beta, \gamma) \leq H_n (P; \alpha, \beta, \gamma) + H_m (Q; \alpha, \beta, \gamma),
\]

(47)

where \( P = (p_1, \ldots, p_n) \), \( Q = (q_1, \ldots, q_m) \) and

\[P \ast Q = (p_1q_1, \ldots, p_mq_m, \ldots, p_nq_1, \ldots, p_nq_m)\]

(48)

are complete probability distributions.

Proof of (i). We have

\[
H_{nm} (P \ast Q; \alpha, \beta, \gamma) = \left( A_{(\alpha,\gamma)} - A_{(\beta,\gamma)} \right)^{-1} \sum_{i=1}^{n} \sum_{j=1}^{m} \left[ (p_iq_j)^{\alpha/\gamma} - (p_iq_j)^{\beta/\gamma} \right]
\]

\[
= \left( A_{(\alpha,\gamma)} - A_{(\beta,\gamma)} \right)^{-1} \sum_{i=1}^{n} \sum_{j=1}^{m} \left[ (p_iq_j)^{\alpha/\gamma} - (p_iq_j)^{\beta/\gamma} \right. \\
+ \left. p_i^{\alpha/\gamma} q_j^{\beta/\gamma} - p_i^{\beta/\gamma} q_j^{\alpha/\gamma} \right]
\]

\[
= \left( A_{(\alpha,\gamma)} - A_{(\beta,\gamma)} \right)^{-1} \sum_{i=1}^{n} \sum_{j=1}^{m} \left[ p_i^{\alpha/\gamma} q_j^{\beta/\gamma} - p_i^{\beta/\gamma} q_j^{\alpha/\gamma} \right. \\
+ \left. p_i^{\alpha/\gamma} q_j^{\beta/\gamma} - p_i^{\beta/\gamma} q_j^{\alpha/\gamma} \right]
\]

Adding (49) and (50), we get

\[
2H_{nm} (P \ast Q; \alpha, \beta, \gamma) = \left( A_{(\alpha,\gamma)} - A_{(\beta,\gamma)} \right)^{-1} \sum_{i=1}^{n} \sum_{j=1}^{m} \left[ q_j^{\alpha/\gamma} q_j^{\beta/\gamma} + q_j^{\alpha/\gamma} + q_j^{\beta/\gamma} \right]
\]

(49)
\[ + (A_{(\alpha,\gamma)} - A_{(\beta,\gamma)})^{-1} \left[ \sum_{j=1}^{m} q_j^{\alpha/\gamma} \sum_{i=1}^{n} (p_i^{\alpha/\gamma} + p_i^{\beta/\gamma}) \right. \\
\left. - \sum_{i=1}^{n} p_i^{\beta/\gamma} \sum_{j=1}^{n} (q_j^{\alpha/\gamma} + q_j^{\beta/\gamma}) \right] \\
= \sum_{i=1}^{n} (p_i^{\alpha/\gamma} + p_i^{\beta/\gamma}) (A_{(\alpha,\gamma)} - A_{(\beta,\gamma)})^{-1} \\
\times \sum_{j=1}^{m} (q_j^{\alpha/\gamma} - q_j^{\beta/\gamma}) \\
+ \sum_{j=1}^{m} (q_j^{\alpha/\gamma} + q_j^{\beta/\gamma}) (A_{(\alpha,\gamma)} - A_{(\beta,\gamma)})^{-1} \\
\times \sum_{i=1}^{n} (p_i^{\alpha/\gamma} - p_i^{\beta/\gamma}). \tag{51} \]

Using (46)

\[ H_{nm}(P \ast Q; \alpha, \beta, s) = G_n(P; \alpha, \beta, \gamma) H_m(Q; \alpha, \beta, \gamma) \\
+ G_m(Q; \alpha, \beta, \gamma) H_n(P; \alpha, \beta, \gamma), \tag{52} \]

which is (45). This completes the proof of part (i).

Proof of (ii). From part (i), we have

\[ H_{nm}(P \ast Q; \alpha, \beta, \gamma) = G_n(P; \alpha, \beta, \gamma) H_m(Q; \alpha, \beta, \gamma) \\
+ G_m(Q; \alpha, \beta, \gamma) H_n(P; \alpha, \beta, \gamma). \tag{53} \]

As \( G_n(P; \alpha, \beta, \gamma) = (1/2) \sum_{i=1}^{n} (p_i^{\alpha/\gamma} + p_i^{\beta/\gamma}) \leq 1 \), for \( \alpha, \beta \geq \gamma \),

\[ H_{nm}(P \ast Q; \alpha, \beta, \gamma) \leq H_m(Q; \alpha, \beta, \gamma) + H_n(P; \alpha, \beta, \gamma). \tag{54} \]

This proves the subadditivity.

4. Conclusion

In addition to well-known information measure of Shannon, Renyi’s, Havrda-Charvat, Vajda [13], Daroczy, we have characterized a measure which we call \( \alpha, \beta \), and \( \gamma \) information measure. We have given some basic axioms and properties with recursive relation. The Shannon’s [5] measure included in the \( \alpha, \beta \), and \( \gamma \) information measure for the limiting case \( \alpha = \gamma = 1 \) and \( \beta \to 1 \); \( \beta = \gamma = 1 \) and \( \alpha \to 1 \). This measure is generalization of Havrda-Charvat entropy.

Conflict of Interests

The authors declare that there is no conflict of interests regarding the publication of this paper.

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