Comparative precision of laboratory methods

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Introduction

Different quantitation techniques or methods of measurement are sometimes employed to measure the same substance in a laboratory. Higher precision, i.e. lower variability, is one of the criteria for selecting a better measurement method. This could be the case when one wants to compare a new (test) method to a standard (reference) method. When independent samples are used between methods, then it is well known that one can use the F-test [1] to compare their variances. However, when the same samples are employed for both quantitative methods, the F-test is inappropriate for comparing the between sample variation and could fail to detect the more precise method.

The appropriate statistical test was proposed by Pitman [2], [1] in 1939. This note illustrates the use of Pitman's test for comparative precision in the situation where the same samples are used for two laboratory quantitative methods.

Method

Denote by SD_i the standard deviation of the quantitation method i ($i = 1, 2$) from N samples, and the variance ratio between the two methods:

$$F = (SD_1/SD_2)^2 \quad (1)$$

If the N samples for method 1 are separate (independent) samples from those for method 2, then the F value in equation (1) behaves like an F -distribution with $(N - 1)$ and $(N - 1)$ degrees of freedom. One can therefore use this statistic for comparing their precision. However, if the same N samples are used for both methods, SD_1 and SD_2 are correlated and the F value is no longer distributed as an F -distribution. A modified test by Pitman is as follows:

$$t = (F - 1)/[(N - 2)/4F(1 - r^2)]^{1/2} \quad (2)$$

is a t -distribution with $(N - 2)$ degrees of freedom, where r is the correlation coefficient between the two methods and F is defined in equation (1).

Example

To illustrate the Pitman's test an example has been taken from Griffiths *et al.* [3]: the sodium levels of 21 patient serum specimens were analysed by Beckman Astra-8 and flame photometry methods. The raw data are reproduced in table 1. From equation (1) the variance ratio:

$$F = (6.725/5.963)^2 = 1.272$$

Table 1. Raw data for Na by two quantitative methods (see Griffiths *et al.*, 1986 [3]).

Sample	Astra	Flame
1	129	130
2	140	139
3	135	137
4	139	138
5	132	131
6	140	139
7	138	137
8	136	137
9	135	135
10	144	145
11	142	142
12	140	139
13	119	121
14	134	135
15	151	149
16	139	138
17	134	133
18	142	141
19	146	143
20	145	143
21	141	142
Mean	138.1	137.8
SD	6.725	5.963
Correlation coefficient (r)	0.9844	
No.	21	

Had we used the F-test here, this statistic would not be statistically significant (probability > 0.50), i.e. the two methods were equally precise. (The critical value of F at 0.05 level with 20 and 20 degrees of freedom = 2.46.) However, from equation (2):

$$t = (1.272 - 1)/[(21 - 2)/4 \times 1.272 (1 - 0.9844^2)]^{1/2} = 2.987 \quad \text{with 19 degrees of freedom}$$

is highly significant (probability < 0.008), i.e. the flame photometry was more precise than the Astra method. (The critical value of t at 0.05 level with 19 degrees of freedom = 2.093.) Here one can see that by ignoring the information that the data come from the same samples, one can fail to detect a superior method with respect to the precision by employing the F-test.

Discussion

The problem of comparing variabilities discussed so far has been concerned with the situation where only a single measurement is available per sample per method. Where there are equal replications for each sample, both the between sample and within sample variations should be examined.

For the between sample variation (the same variability discussed in this note so far), one can take the average values over replicates and apply the Pitman's test on the sample averages. The individual replications do not enter into the statistical test directly.

The within sample variation measures the repeatability of the quantitation method on the same sample. Unlike the between sample variations, the within sample variations are not correlated between methods. In this instance one can use the F-test again for comparing the within sample variations. The statistical model with replication within sample is given in the Appendix.

Appendix

The model:

$$y_{ijk} = \mu_{jk} + \eta_i + \tau_{ij} + \varepsilon_{ijk}$$

where y_{ijk} is the determination from sample i ($i = 1, \dots, N$), quantitation method j ($j = 1, 2$) and replicate k ($k = 1, \dots, R$). μ_{jk} is the mean response of method j and replicate k , and η_i , τ_{ij} and ε_{ijk} are the residual error terms due to sample i , method j and replicate k respectively. Further assumptions are made on the residual error terms such that they are mutually independent and normally distributed:

$$\begin{aligned} \eta_i &\sim N(0, \sigma_S^2) \\ \tau_{ij} &\sim N(0, \sigma_{Mj}^2) \\ \varepsilon_{ijk} &\sim N(0, \sigma_{Wj}^2) \end{aligned} \quad \text{and}$$

Denote by $\bar{y}_{ij}$ the mean of sample i and method j over the R replicates (i.e. $\bar{y}_{ij} = \sum_k y_{ijk}/R$) etc., then the between sample variance for method j :

$$\text{var}(\bar{y}_{ij}) = \sigma_S^2 + \sigma_{Mj}^2 + \sigma_{Wj}^2/R$$

and can be compared by the Pitman's test (using $\bar{y}_{ij}$).

σ_{Wj}^2 is the within sample variance, i.e. the variation between determinations due to replication. It is estimated by:

$$S_{Wj}^2 = \sum_{i,k} (y_{ijk} - \bar{y}_{ij} - \bar{y}_{jk} + \bar{y}_j)^2 / (N-1)(R-1)$$

with $(N-1)(R-1)$ degrees of freedom. The F-test can be used here to compare S_{Wj}^2 's:

$$F = S_{W1}^2 / S_{W2}^2$$

is distributed as an F-distribution with $(N-1)(R-1)$ and $(N-1)(R-1)$ degrees of freedom.

References

1. SNEDECOR, G. W. and COCHRAN, W. G. *Statistical Methods* (The Iowa State Uni. Press, Iowa, 1976), p. 116, and p. 195.
2. PITMAN, E. J. G., *Biometrika*, **31** (1939), 9.
3. GRIFFITHS, W. C., CAMARA, P., DIAMOND, I. and PEZZULLO, J. C., *Journal of Automatic Chemistry*, **8** (1986), 147.

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- Biosensors (C. R. Lowe, University of Cambridge, UK).
- Physiological sensors (D. Parker, University College Hospital, London, UK).
- Sensors in industrial metrology (B. E. Jones, Brunel University, UK).

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