Research Article

Preliminary Compositional Evidence of Provenance of Ceramics from Hatahara Archaeological Site, Central Amazonia

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One hundred twenty-four ceramic fragments and six clay samples from the Hatahara archaeological site in Amazonas state, Brazil, were analyzed using instrumental neutron activation analysis, INAA, to determine the concentration of twenty chemical elements: Ce, Co, Cr, Cs, Eu, Fe, Hf, K, La, Lu, Na, Nd, Rb, Sc, Ta, Tb, Th, U, Yb, and Zn. The dataset was submitted to multivariate statistical analysis. The classification was done by cluster analysis and discriminant analysis. The results demonstrated the occurrence of four different groups of ceramics, which represent three archaeological phases: Paredão, Manacapuru, and Guarita. This data is consistent with previous traditional petrographic examination of the ceramic samples. Based on probability measures, the great majority of the ceramics are considered to be local in origin.

1. Introduction

Archaeological ceramics are the most abundant tracers in archaeological studies because they maintain their characteristics for many years, constitute a historic record, and can reflect the behavior patterns of ancient communities. For this reason, the chemical analysis of this kind of sample has attracted the attention of many authors worldwide [1].

The production technology, provenance and use of the ceramics by ancient communities have been the main issues approached since 1950 when the first studies of these materials began to appear [2]. The determination of provenance probably forms the primary role of geochemical analyses in archaeological ceramic studies.

In provenance studies, the elementary concentrations in ceramics and clays are chemically and mineralogically correlated. By determining the trace elements in both ceramics and clays, the raw material used for their manufacture can be identified. The group differentiation depends on the discriminant element concentration, which indicates whether there existed one or more clay deposits. Although it is a complex exercise, especially with purely chemical studies, there are cases in which this has been successful [3, 4]. The main reason for this complexity is that the chemical and mineral compositions of ceramics do not correspond in a simple way to any one distinct raw material. A number of compositional changes may be introduced in the manipulation, and mixing of the raw material, as well as the history of the ceramic in use, the nature of the production process, and the post-depositional alteration can lead to compositional variations between the ceramic fabric and the raw material(s) used to produce it.

Due to these problems, it is common in ceramic provenance studies to compare archaeological pottery with reference material, comprised of ceramic samples with a known or assumed origin. Although successful in many cases, this approach relies on the opinion and assumptions of what comprises local pottery for a given archaeological site. Overall, raw materials used in ceramics continue to be the primary sources for interpreting the chemical and mineralogical diversity of pottery, in order to differentiate production units and determine their location.
From this point of view, physical and chemical techniques are important tools for these kinds of studies. Among the various techniques, INAA employing γ-ray spectrometry seems to be most suitable analytical technique because it enables several elements to be simultaneously determined with high sensitivity, precision, and accuracy. Moreover, sample preparation is relatively easy and fast [5–7].

In this study, the concentration data for the elements Ce, Co, Cr, Cs, Eu, Fe, Hf, K, La, Lu, Na, Nd, Rb, Sc, Ta, Tb, Th, U, Yb, and Zn was obtained via INAA on the ceramics and clay samples to determine the different production units and locations. In order to form chemical groups, the data was studied using cluster and discriminant analysis to find the source of the raw material used in the ceramics.

1.1. Study Area. The Hatahara site belongs to the district of Iranduba, 30 km southwest of Manaus, in the region located on the left margin of the Solimões river, next to where it merges with the Negro river [8]. The site demonstrates the history of four distinct phases of occupation: Paredão, Manacupuru, Guarita, and Açutuba [8]. The phases were established based on the decoration, form of the vessel, and mostly by the kind of the temper used in the paste. Such aspects characterize the site as complex in relation to its pre-colonial composition, and archaeometric studies are important in understanding the development of these communities.

The site is formed by a farm that encompasses 160,000 m² in a fertile valley in the center of the Amazon. This region contains both dry and flooded land (swamp). The dry land, which is not inundated by the annual river's flooding, is sustained by sedimentary rocks formations Alter do Chão and Novo Remanso from the Cretaceous and Miocene periods, respectively [9–11]. The swamp area, inundated by annual floods, corresponds to the Quaternary sedimentary deposits from the Solimões and Ariaú river flood plains.

The site is characterized by an enormous extension along the Solimões river as well as by the presence of an enormous amount of ceramic material and black soil, which is a type of soil rich in organic material. This soil is present throughout the site's entire surface.

Artificially structured mounds were found at the site made from layers of oxisol, black organic soil, fauna, flora, and human burials remains [10].

2. Experimental

2.1. Sample Collection, Preparation, and Description of the Method. Care was taken during the collection of the samples to ensure, to the greatest extent possible, that duplicates from the same item of ceramic were not included in the sample set. The ceramics were tentatively identified as representing all archaeological phases.

Ceramic powder samples were obtained by cleaning the outer surface and drilling with a variable speed drill to a depth of 1-2 cm, using a tungsten carbide rotary file attached to the end of a flexible shaft. Depending on the thickness, 3 or 5 holes were drilled as deep into the core of the fragment as possible without drilling through the walls. Clay samples were ground in an agate mortar until a granulometry of 100 mesh was achieved. Finally, the powered samples were dried in an oven at 105°C for 24 h and stored in desiccators [5, 7].

Constituent elements in Coal Fly Ash - NIST-SRM-1633b were used as the standard in every analysis. The standard reference material IAEA Soil 7 (Trace Elements in Soil) was used to check the analytical quality of the results. The standards were dried in an oven at 105°C for 2 h and stored in a desiccator to be later weighed.

The bulk samples were analyzed using instrumental neutron activation analysis, following the routine measurement procedure applied to ceramics and clays [5]. Approximately 100 mg of ceramics samples, NIST-SRM-1633b and Soil 7, was weighed in polyethylene bags and wrapped in aluminum foil. Groups of 8 ceramics samples and two reference materials were packed in aluminum foil and irradiated in the research reactor pool, IEA-R1 m (IPEN/CNEN-SP), at a thermal neutron flux of about $5 \times 10^{12}$ cm$^{-2}$ s$^{-1}$ for 8 h.

Two measurement series were carried out using Ge (hyperpure) detector, model GX 2519 from Canberra, with a resolution of 1.90 keV at the 1332.49 keV gamma peak of $^{60}$Co. Spectra were collected with a Canberra S-100 multichannel analyzer with 8192 channels. K, La, Lu, Na, Nd, Sm, U, and Yb were measured after 7 days of cooling time, and Ce, Co, Cr, Cs, Eu, Fe, Hf, Rb, Sb, Sc, Ta, Tb, Th, and Zn were measured after 25–30 days. Gamma-ray spectra analysis was carried out using the Genie 2000 NAA Procedure software from Canberra [5].

3. Results and Discussion

The analytical quality control of the analysis was tested using 28 independent determinations of the reference material IAEA Soil-7. The results observed were compared to the certified values. Some statistical studies were applied to the dataset, such as mean, mean confidence interval, and relative standard deviation. The results showed that most elements had a precision of ≤10%. This precision is considered by several authors as appropriate for the choice of chemical elements for provenance studies using multivariate statistical methods [12]. The clays in the different sampling loci of the region may not differ greatly in composition; therefore, the method of analysis must be sensitive enough to cope with this problem. Since the Co determination showed a precision of less than 10%, it was eliminated in the dataset due to contamination by tungsten carbide during the sample preparation [13].

The determination of Zn is not reliable as a consequence of a strong γ-ray interference by $^{40}$Sc and $^{182}$Ta. The interference by the $^{235}$U fission in determining the La, Ce, and Nd was negligible because the U concentration did not exceed 5 ppm and the rare earth elements were not extremely low [14]. Nd, Rb, and Sb showed a good precision; however, previous studies have shown that there are no reliable elements to include in the database due to the natural heterogeneity [15]. Therefore, the elements used in the subsequent studies were Ce, Cr, Cs, Eu, Fe, Hf, K, La, Lu, Na, Sc, Tb, Th, U, and Yb. None of these elements contained missing values. The mean, standard deviation, minimum, and maximum are presented in Table 1.
In geochemistry, concentration data is often assumed to follow a lognormal distribution after being \( \log_{10} \) transformed, as suggested by Ahrens [16]; however, in geochemistry, this assumption rarely holds true. For the majority of the variables, a log base 10 transformation does not result in a normal distribution [17]. This fact may have serious consequences to the later statistical treatment of datasets because the vast majority of advanced statistical methods require not only that each variable shows a normal distribution but also that the variables show a multivariate normal distribution. In addition, although the dataset does not present the total composition of the samples, that is, the variables measured are <100%, this type of data frequently displays a curvature. In this paper, the transformation proposed by Aitchison [18] was used, so each sample \( x_{ij} \) (\( i = 1, \ldots, n \) and \( j = 1, \ldots, p \)) was transformed into \( y_{ij} \) by taking the natural log transformation and subtracting the mean of the transformed variables; that is,

\[
y_{ij} = \ln x_{ij}, \quad \overline{y}_i = \frac{1}{p} \sum_{j=1}^{p} y_{ij}, \quad z_{ij} = y_{ij} - \overline{y}_i.
\]  

In addition, the data was standardized to offset the large differences in magnitude between the elements measured at trace levels and the more prevalent elements [19]. The most universally used method is \( z \)-transformation, in which the raw data is subtracted by the mean and then divided by the standard deviation of the data. When working with geochemical data, a robustified version, using the median and the median absolute deviation (MAD) is preferred instead of the mean and standard deviation [20]. Therefore, this procedure was used as follows [19]:

\[
z\text{-transformation} = \frac{z_{ij} - \text{median}(z_i)}{\text{MAD}(z_i)}.
\]

Table 1: Mean, standard deviation, and range for the ceramic samples, \( n = 124 \), and clay, \( n = 6 \), in \( \mu g/g \) unless otherwise indicated.

<table>
<thead>
<tr>
<th>Element</th>
<th>Mean ± SD</th>
<th>Range</th>
<th>Clay 1</th>
<th>Clay 2</th>
<th>Clay 3</th>
<th>Clay 4</th>
<th>Clay 5</th>
<th>Clay 6</th>
</tr>
</thead>
<tbody>
<tr>
<td>Na, %</td>
<td>0.3 ± 0.1</td>
<td>0.1–0.7</td>
<td>0.8</td>
<td>0.1</td>
<td>0.7</td>
<td>1.2</td>
<td>1.2</td>
<td>0.6</td>
</tr>
<tr>
<td>K, %</td>
<td>1.2 ± 0.4</td>
<td>0.4–2.6</td>
<td>1.6</td>
<td>0.9</td>
<td>1.5</td>
<td>2.3</td>
<td>2.3</td>
<td>2.5</td>
</tr>
<tr>
<td>La</td>
<td>38.9 ± 7.9</td>
<td>20.1–93.9</td>
<td>43.6</td>
<td>24.4</td>
<td>36.8</td>
<td>36.4</td>
<td>39.7</td>
<td>56.8</td>
</tr>
<tr>
<td>Yb</td>
<td>2.8 ± 0.6</td>
<td>1.6–7.7</td>
<td>4.1</td>
<td>3.3</td>
<td>2.8</td>
<td>3.3</td>
<td>3.1</td>
<td>3.6</td>
</tr>
<tr>
<td>Lu</td>
<td>0.4 ± 0.1</td>
<td>0.3–1.2</td>
<td>0.5</td>
<td>0.5</td>
<td>0.4</td>
<td>0.4</td>
<td>0.5</td>
<td>0.6</td>
</tr>
<tr>
<td>U</td>
<td>3.2 ± 0.7</td>
<td>0.5–7.5</td>
<td>3.6</td>
<td>3.9</td>
<td>3.3</td>
<td>3.2</td>
<td>3.2</td>
<td>3.5</td>
</tr>
<tr>
<td>Ce</td>
<td>73.2 ± 11.1</td>
<td>39.5–96.8</td>
<td>107.9</td>
<td>40.9</td>
<td>73.0</td>
<td>76.3</td>
<td>81.3</td>
<td>105.7</td>
</tr>
<tr>
<td>Cr</td>
<td>61.7 ± 8.2</td>
<td>40.8–89.0</td>
<td>60.1</td>
<td>91.5</td>
<td>59.1</td>
<td>60.6</td>
<td>63.0</td>
<td>92.7</td>
</tr>
<tr>
<td>Cs</td>
<td>6.4 ± 1.7</td>
<td>2.6–10.6</td>
<td>7.3</td>
<td>8.3</td>
<td>7.1</td>
<td>4.6</td>
<td>5.3</td>
<td>13.7</td>
</tr>
<tr>
<td>Eu</td>
<td>1.2 ± 0.2</td>
<td>0.7–1.9</td>
<td>1.7</td>
<td>0.5</td>
<td>1.3</td>
<td>1.5</td>
<td>1.7</td>
<td>1.8</td>
</tr>
<tr>
<td>Fe, %</td>
<td>3.3 ± 0.7</td>
<td>1.5–4.7</td>
<td>4.0</td>
<td>6.3</td>
<td>3.0</td>
<td>4.1</td>
<td>4.4</td>
<td>3.9</td>
</tr>
<tr>
<td>Hf</td>
<td>5.8 ± 2.4</td>
<td>3.1–17.6</td>
<td>8.1</td>
<td>11.4</td>
<td>9.4</td>
<td>8.9</td>
<td>9.5</td>
<td>4.7</td>
</tr>
<tr>
<td>Sc</td>
<td>14.5 ± 1.9</td>
<td>9.1–19.3</td>
<td>14.7</td>
<td>14.6</td>
<td>13.6</td>
<td>14.8</td>
<td>15.6</td>
<td>21.4</td>
</tr>
<tr>
<td>Tb</td>
<td>0.7 ± 0.2</td>
<td>0.3–1.4</td>
<td>1.2</td>
<td>0.8</td>
<td>1.2</td>
<td>0.8</td>
<td>0.8</td>
<td>0.6</td>
</tr>
<tr>
<td>Th</td>
<td>13.3 ± 2.0</td>
<td>8.7–20.0</td>
<td>12.0</td>
<td>17.6</td>
<td>11.7</td>
<td>12.0</td>
<td>12.9</td>
<td>17.2</td>
</tr>
</tbody>
</table>

In turn, the dataset was submitted to outlying tests using the Mahalanobis distance. Outliers can have a considerable influence on multivariate statistical methods because they can disturb homogeneous groups.

The Mahalanobis distance is an important measure in statistics and has been suggested by many authors as the best method for detecting outliers in multivariate data. For each of the \( n \) samples and \( p \) variables, the Mahalanobis distance (\( D_i \)) was taken from the sample to the centroid, as calculated by the expression [20]

\[
D_i = \sqrt{(x_i - \overline{x})' S^{-1} (x_i - \overline{x})},
\]

where ‘ is the transpose matrix, \( S = \sum_{i=1}^{n} (x_i - \overline{x})' (x_i - \overline{x}) \) is the variance-covariance sampling matrix, and \( (x_i - \overline{x}) \) is the vector of difference between the concentrations measured in one group and the concentrations measured in the other group. Each one of these values is compared with the critical value, \( c_v \), which can be calculated using the lambda Wilks criteria [21], calculated as follows:

\[
\frac{p(n-1)^2 F_{p,n-p-1,\alpha/n}}{n(n-p-1 + p^2 F_{p,n-p-1,\alpha/n})},
\]
lower or higher concentrations of Fe, Cr, K, U, Tb, Th, La, and Ce were found compared to other samples. This result is not altogether surprising because correlations exist between elements that are geochemically related and that are neighbors or nearly neighbors in the same group, subgroups, or series in the periodic table. For example, members of the lanthanide series (La, Ce) demonstrate strong positive correlations, as do members of the actinides series (U and Th) and the elements of the first transition series (Fe and Co).

Cluster analysis is routinely applied to analytical data with the purpose of identifying groups of chemically similar shards, which can be interpreted as representing vessels made from the same raw materials or mixtures of raw materials and therefore presumably made in the same place although not necessarily at the same time.

The data from the remaining 121 ceramics samples were submitted to cluster analysis on a $121 \times 15$ matrix, in which the columns represented the analyzed elements and the rows the samples using Ward’s method and square Euclidean distance. Ward’s method was employed because it tends to form groups with high internal homogeneity and takes account of the cluster structure. In all statistical studies, the software used was R version 2.4.1.

In Figure 1, the dendrogram shows that the samples were classified into five primary groups, linked at different levels of similarity. With such a large number of samples, this result is not altogether surprising because correlations exist between elements that are geochemically related and that are neighbors or nearly neighbors in the same group, subgroups, or series in the periodic table. For example, members of the lanthanide series (La, Ce) demonstrate strong positive correlations, as do members of the actinides series (U and Th) and the elements of the first transition series (Fe and Co).

Cluster analysis is routinely applied to analytical data with the purpose of identifying groups of chemically similar shards, which can be interpreted as representing vessels made from the same raw materials or mixtures of raw materials and therefore presumably made in the same place although not necessarily at the same time.

Figure 2: Discriminant function 1 versus discriminant function 2 for the ceramic samples, $n = 121$. The ellipses represent 95% confidence level for sample inclusion in the cluster.
belonging to each group. It is possible to see that no variation occurred in relation to the number of groups formed by the dendrogram. Groups 1 and 3 are represented by ceramics belonging to the Manacapuru phase. These groups revealed only the presence of ceramics related to this phase, but they used a different source of clay for each group.

It is possible to also see that larger amounts of ceramics exist in Groups 2, 4, and 5 where the presence of the Paredão phase ceramics is predominant. Although in these groups the concentrations of some elements were similar, the concentrations of some elements were clear in the alkali metals and the rare earth elements, and, particularly Eu (Table 1). The high Na concentration may be accounted for by the plagioclase content, and Eu may exist in the divalent state (Eu²⁺) and it substitutes Sr in Ca plagioclase [24].

The clay samples were analyzed with the purpose of studying the raw material sources used in the production of the ceramic artifacts. The geographical domain of the provenence studies for the clays collected was confined to a relatively small region. In order to accomplish this, six clay samples were collected and analyzed by INAA. Three of them were collected from the Solimões river flood plains in front of the city of Iranduba, which is 6 km from the site. The other three clay samples were taken from the lowlands of the Ariau river. The results are presented in Table 1. The elementary concentrations of the clays were submitted to discriminant analysis together with ceramic samples in order to verify the distribution of the samples in the ceramics groups. Figure 3 shows discriminant function 1 versus discriminant function 2 for the clay and ceramics samples. The six clays show patterns that probably have a chemical composition that is somewhat similar to the ceramics of the Paredão phase and other undefined ceramic groups (clusters 2, 4, and 5). Considering the abundance of criteria, this suggests that probably clay could be used as raw material for the Paredão ceramic phase.

4. Conclusion

Through the analysis of Ce, Cr, Cs, Eu, Fe, Hf, K, La, Lu, Na, Sc, Tb, Th, U, and Yb in 121 ceramics samples and 6 clays from the Hatahara archaeological site determined through the INAA, it was possible to confirm the three different and concurrent phases of occupation as well as the probable source of the raw material used in the production of the ceramic artifacts. The results suggest that no single clay source was used at this site to manufacture the ceramics analyzed in this study. However, our conclusions are based on a small sample size, and further studies with the analysis of more clay samples are needed to confirm the provenance.

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