Quantification of Margins and Uncertainties Approach for Structure Analysis Based on Evidence Theory

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Quantification of Margins and Uncertainties (QMU) is a decision-support methodology for complex technical decisions centering on performance thresholds and associated margins for engineering systems. Uncertainty propagation is a key element in QMU process for structure reliability analysis at the presence of both aleatory uncertainty and epistemic uncertainty. In order to reduce the computational cost of Monte Carlo method, a mixed uncertainty propagation approach is proposed by integrated Kriging surrogate model under the framework of evidence theory for QMU analysis in this paper. The approach is demonstrated by a numerical example to show the effectiveness of the mixed uncertainty propagation method.

1. Introduction

The uncertainties of material properties, environment loads, and design models are inevitable in engineering. Uncertainty is usually classified into aleatory and epistemic types \cite{1, 2}, and the presence of uncertain factors introduces uncertainty in the reliability of the structure. Probabilistic approaches that deal with aleatory parameter uncertainty have been vastly investigated in typical structure reliability analysis. When sufficient data are not available or there is lack of information due to ignorance, the concept of subjective probability is well established for quantifying epistemic uncertainty. However, when the probabilities of rare events are very difficult to assess or events occur only once, the classical probability methodology may not be suitable because there is not enough statistic data \cite{3, 4}. In order to overcome the lack of probabilistic method, the nonprobabilistic methods have been proposed and are more suitable to handle the epistemic uncertainty based on the Fuzzy theory, interval theory, evidence theory, and so forth. The theoretical concept and the application of the Quantification of Margins and Uncertainties (QMU) methodology were reported in the certification of reliability, safety of nuclear weapons stockpile, and risk-informed decision making process under restriction of test data in the last decade. Recently, the QMU methodology has been applied to more general complex systems, such as commercial nuclear power plants, reactor safety, and missile reliability \cite{5, 6}. Eardley \cite{7} described the main components (performance gates, margins, and uncertainties) of the QMU methodology. Under the key ideas and application procedures of QMU methodology, the uncertainty propagation that determines the output uncertainty from input uncertainty is a broad research area for QMU process. Reference \cite{8} shows that evidence theory is a more general theory that can handle both types of uncertainty but it requires much more computational cost.

The objective of this paper is to propose an implementation framework of QMU under mixed uncertainty based on the evidence theory. To alleviate the computational costs, a stochastic surrogate model based on Kriging model and adaptive sampling method has been applied for uncertainty propagation for structure performance response. The rest of this paper is organized as follows. Section 2 briefly introduces the basic concept and metric of QMU. Section 3 details the mixed uncertainty propagation using evidence theory in QMU implementation. The new calculation scheme of mixed uncertainty analysis by integrating Kriging model and
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Performance gate upper bound

Performance gate lower bound

(failure threshold)

Margin (M)

Performance uncertainty (U)

Figure 1: Notional illustration of the QMU.

The Bel and Pl of a given event set A and B can be derived from the basic probability assignment by

\[ Bel(B) = \sum_{A \subseteq B} m(A), \]

\[ Pl(B) = \sum_{B \subseteq A} m(A). \]  

Resulting from (2), the belief function, Bel(\cdot), is calculated by summing the BPAs that totally agree with event B, while the plausibility function, Pl(\cdot), is calculated by summing BPAs that agree with event B totally and partially. Both Bel and Pl play roles similar to distribution functions in the standard probability and they give the lower and upper bounds of the event set.

3.2. Mixed Uncertainty Propagation Using Evidence Theory

The mixed uncertainty propagated from input parameters to systems output needs to be quantified for structure reliability analysis when aleatory and epistemic uncertainties exist. The performance function with aleatory and epistemic uncertainties can be given by

\[ G(Z) = g(X, Y), \]

where \( X = [X_1, X_2, \ldots, X_{n_X}], Y = [Y_1, Y_2, \ldots, Y_{n_Y}] \) represents the aleatory uncertainty variables described by probability distributions. For easy demonstration, we assume that the elements of \( X \) are independent. \( Y = [Y_1, Y_2, \ldots, Y_{n_Y}] \) is the vector of parameters with epistemic uncertainty described by evidence specification \( (\mathcal{Y}_1, Y_1, m_1), (\mathcal{Y}_2, Y_2, m_2), \ldots, (\mathcal{Y}_{n_Y}, Y_{n_Y}, m_{n_Y}) \). The uncertainty associated with the model inputs \( X \) and \( Y \) are propagated through the model \( g(\cdot) \) to the model output \( G \). The joint evidence specification can be expressed by \( (\mathcal{Y}, \mathcal{V}, m) \) and the joint BPA is defined by

\[ m_\mathcal{Y}(\mathcal{U}) = \begin{cases} \prod_{k=1}^{n_Y} m_\mathcal{V}(\mathcal{U}_k), & \mathcal{U} = \mathcal{U}_1 \times \mathcal{U}_2 \times \cdots \times \mathcal{U}_{n_Y} \\ 0, & \text{Otherwise}. \end{cases} \]  

Let the number of the subsets of \( Y \) in the joint space be \( n \). The probability of the system output value \( G \) is less than the threshold \( c \) which can be given by the following based on the total probability formula:

\[ P = \sum_{i=1}^{n} \Pr[G(X, Y) < c \mid Y_i \in \mathcal{U}_i] m_\mathcal{Y}(\mathcal{U}_i), \]  

where \( P \) is the probability of \( G(Z) < c \), which is denoted as event \( F \). \( \Pr[Y_i \in \mathcal{U}_i] = m_\mathcal{V}(\mathcal{U}_i) \) means the probability of the focal element \( \mathcal{U}_i \) in joint space \( \mathcal{V} \) equals the joint BPA value. Because of the intervals in \( Y \), the minimum value (belief) and maximum value (plausibility) of \( P \) can be expressed by [11]

\[ Bel(F) = P_{\min} = \sum_{i=1}^{n} m_\mathcal{Y}(\mathcal{U}_i) \Pr[G_{\max}(X, Y) < c \mid Y_i \in \mathcal{U}_i], \]  

\[ Pl(F) = P_{\max} = \sum_{i=1}^{n} m_\mathcal{Y}(\mathcal{U}_i) \Pr[G_{\max}(X, Y) < c \mid Y_i \in \mathcal{U}_i]. \]
\[ P_l(F) = P_{\text{max}} = \sum_{i=1}^{n} m_Y(U_i) \Pr\{G_{\text{min}}(X,Y) < c \mid Y_i \in U_i\}. \]  

(7)

3.3. The Solution Framework of Mixed Uncertainty Propagation. Equations (6)-(7) can be calculated using sampling methods such as MCS with large computational cost. The Kriging surrogate model [12, 13] can be employed to reduce the cost for uncertainty propagation. With training observations, the response and predicted mean square error for any given new point \( x' \) can be expressed as

\[ \tilde{G}(x') = \mu + r^T R^{-1} (G - A \mu), \]

\[ \tilde{e}(x') = \sigma^2 \left[ 1 - r^T R^{-1} r + \frac{(1 - A^T R^{-1} r)^2}{A^T R^{-1} A} \right], \]  

where \( R \) represents a correlation matrix, \( r \) is the correlation vector between \( x' \) and the observed samples, and \( A \) is an \( n \times 1 \) unit vector. This surrogate model is used to evaluate the uncertainty distribution of system output based on the traditional MCS method without calling the original performance function. The Maximum Confidence Enhancement adaptive sampling [13] is employed to ensure the surrogate model accuracy by adding new training sample.

The procedure of the kriging-based method for solving (7) is simply introduced as follows:

1. Calculate the joint BPA of epistemic uncertainties variables \( Y \).
2. Use the Latin hypercube sampling (LHS) to generate \( N = (n + 1)(n + 2)/2 \) sample points \( x_t = (x_{1t}, \ldots, x_{nt}) \) \((t = 1, 2, \ldots, N)\) of the aleatory variables \( X \).
3. For each \( x_t \), calculate the minimum output value of \( G_{\text{min}}(x_t) \) in focal element \( U_t \) of joint space.
4. Build Kriging model based on training data \( x_t \) and \( G_{\text{min}}(x_t) \).
5. Calculate the maximum Confidence point \( (x^*) \) from the MC sampling and update Kriging model by adding new training sample \((x^*, G_{\text{min}}(x^*))\).
6. Repeat steps (4)-(5) until the surrogate model approximate accuracy is satisfied.
7. Calculate the probability boundary using the updated Kriging model and MC samples.
8. Repeat steps (3)-(7) for each focal element of \( Y \).
9. Compute the \( P_l(\cdot) \) based on the joint BPA and probability of each focal element.

For the Bel(\cdot) solution, the minimum output value \( G_{\text{min}}(x_t) \) needs to change to the maximum output value \( G_{\text{max}}(x_t) \).

3.4. Calculation of CF under the Evidence Theory Framework. As described in Section 2, the safety or reliability of a structural system will be measured by the Confidence Factor, CF. When the system output uncertainty is represented in terms of evidence theory measures with belief and plausibility, the demonstration of uncertainties and margins with distance between system output and performance gate boundary is shown in Figure 2. The performance gate boundary is a random variable.

The calculations of \( M \) and \( U \) to perform upper boundary are given by the following equations:

\[ M = \left| \text{Bel}_{P=0.5} - F_{P=0.5} \right|, \]

\[ U = U_{\text{function}} + U_T \]

\[ = \left| \text{Bel}_{P=(1+\gamma)/2} - P_{l=0.5} \right| + \left| F_{P=0.5} - F_{P=(1-\gamma)/2} \right|, \]  

where the subscript \( P \) corresponds to the quantity of belief/plausibility/probability level and \( \gamma \) is the specified confidence level.

4. Application Example for Structure Analysis

Figure 3 shows a crank-slider mechanism [11]. The length of the crank \( a \), the length of the coupler \( b \), the external force \( P \), Young’s modulus of the material of the coupler \( E \), and the yield strength of the coupler \( S \) are random variables. The coefficient of friction \( \mu \) between the ground and the slider and the offset \( e \) are epistemic variables. The random variables and epistemic variables with BPA are provided in Tables 1 and 2.
Table 1: Random variables.

<table>
<thead>
<tr>
<th>Variable</th>
<th>Parameters symbol</th>
<th>Mean</th>
<th>Standard deviation</th>
<th>Distribution</th>
</tr>
</thead>
<tbody>
<tr>
<td>X_1</td>
<td>a (mm)</td>
<td>100</td>
<td>0.01</td>
<td>Normal</td>
</tr>
<tr>
<td>X_2</td>
<td>b (mm)</td>
<td>400</td>
<td>0.01</td>
<td>Normal</td>
</tr>
<tr>
<td>X_3</td>
<td>P (KN)</td>
<td>280</td>
<td>28</td>
<td>Normal</td>
</tr>
<tr>
<td>X_4</td>
<td>E (GPa)</td>
<td>200</td>
<td>10</td>
<td>Normal</td>
</tr>
<tr>
<td>X_5</td>
<td>S (MPa)</td>
<td>290</td>
<td>29</td>
<td>Normal</td>
</tr>
<tr>
<td>X_6</td>
<td>d_1 (mm)</td>
<td>60</td>
<td>3</td>
<td>Normal</td>
</tr>
<tr>
<td>X_7</td>
<td>d_2 (mm)</td>
<td>25</td>
<td>2.5</td>
<td>Normal</td>
</tr>
</tbody>
</table>

Table 2: Uncertain variables with epistemic uncertainty.

<table>
<thead>
<tr>
<th>Variable</th>
<th>Symbol</th>
<th>Intervals</th>
<th>BPA</th>
</tr>
</thead>
<tbody>
<tr>
<td>Y_1</td>
<td>ε</td>
<td>[100, 120]</td>
<td>0.2</td>
</tr>
<tr>
<td></td>
<td></td>
<td>[120, 140]</td>
<td>0.4</td>
</tr>
<tr>
<td></td>
<td></td>
<td>[140, 150]</td>
<td>0.4</td>
</tr>
<tr>
<td>Y_2</td>
<td>µ</td>
<td>[0.15, 0.18]</td>
<td>0.3</td>
</tr>
<tr>
<td></td>
<td></td>
<td>[0.18, 0.23]</td>
<td>0.3</td>
</tr>
<tr>
<td></td>
<td></td>
<td>[0.23, 0.25]</td>
<td>0.4</td>
</tr>
</tbody>
</table>

Table 3: Calculated CF vs different γ.

<table>
<thead>
<tr>
<th>γ</th>
<th>M</th>
<th>U</th>
<th>CF</th>
</tr>
</thead>
<tbody>
<tr>
<td>0.95</td>
<td>133.5</td>
<td>138</td>
<td>0.9674</td>
</tr>
<tr>
<td>0.9</td>
<td>133.5</td>
<td>117.5</td>
<td>1.136</td>
</tr>
<tr>
<td>0.8</td>
<td>133.5</td>
<td>93.5</td>
<td>1.428</td>
</tr>
</tbody>
</table>

Figure 4: The uncertainty distribution of system performance and threshold.

The performance function is defined by the maximum stress as (10) and the boundary is the material strength:

\[ Y = \frac{4P(b-a)}{\pi \left( \sqrt{(b-a)^2 - e^2} - \mu_e \right) (d_1^2 - d_2^2)} \]  \hspace{1cm} (10)

The belief and plausibility measures of the maximum stress are calculated by the proposed adaptive sampling Kriging model approach, and these results are shown in Figure 4. The QMU analysis with different γ is summarized in Table 3. Results show that the confidence level is very sensitive for QMU analysis in risk-informed decision-making.

5. Conclusion

The mixed uncertainty propagation approach is proposed by integrated adaptive sampling method and Kriging model for QMU analysis in this paper. The technique is demonstrated by a numerical example to account for the QMU analysis process and the approach for mixed uncertainty propagation. The results indicate the potential effectiveness of the proposed QMU approach for the evaluation of structure reliability.

Competing Interests

The authors declare that there is no conflict of interests regarding the publication of this paper.

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