

Research Article **Optimal Bounds for the Variance of Self-Intersection Local Times**

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For a \mathbb{Z}^d -valued random walk $(S_n)_{n \in \mathbb{N}_0}$, let l(n, x) be its local time at the site $x \in \mathbb{Z}^d$. For $\alpha \in \mathbb{N}$, define the α -fold selfintersection local time as $L_n(\alpha) := \sum_x l(n, x)^{\alpha}$. Also let $L_n^{\mathsf{SRW}}(\alpha)$ be the corresponding quantities for the simple random walk in \mathbb{Z}^d . Without imposing any moment conditions, we show that the variance of the self-intersection local time of any genuinely *d*-dimensional random walk is bounded above by the corresponding quantity for the simple symmetric random walk; that is, $\operatorname{var}(L_n(\alpha)) = O(\operatorname{var}(L_n^{\mathsf{SRW}}(\alpha)))$. In particular, for any genuinely *d*-dimensional random walk, with $d \ge 4$, we have $\operatorname{var}(L_n(\alpha)) = O(n)$. On the other hand, in dimensions $d \le 3$ we show that if the behaviour resembles that of simple random walk, in the sense that $\liminf_{n\to\infty} \operatorname{var}(L_n(\alpha)) / \operatorname{var}(L_n^{\mathsf{SRW}}(\alpha)) > 0$, then the increments of the random walk must have zero mean and finite second moment.

1. Introduction and Main Results

Let $X, X_1, X_2, ...$ be independent, identically distributed, \mathbb{Z}^d -valued random variables, and define the random walk $S_0 := 0$, $S_n = \sum_{j=1}^n X_j$, for $n \ge 1$. The special case with $\mathbb{P}(X_i = e) = 1/(2d)$, for all $e \in \mathbb{Z}^d$ with |e| = 1, is known as the *simple random walk* in \mathbb{Z}^d and will be denoted by $(SRW_n)_{n \in \mathbb{N}_0}$.

Let $l(n, x) = \sum_{j=1}^{n} \mathbb{1}(S_j = x)$ be the local time of $(S_n)_{n \in \mathbb{N}_0}$ at the site $x \in \mathbb{Z}^d$, and define for a positive integer α the α -fold *self-intersection local time*

$$L_n = L_n(\alpha) = \sum_{x \in \mathbb{Z}^d} l(n, x)^{\alpha}$$

=
$$\sum_{i_1, \dots, i_{\alpha}=0}^n \mathbb{1}\left(S_{i_1} = \dots = S_{i_{\alpha}}\right).$$
 (1)

We will denote the corresponding quantities for simple random walk in \mathbb{Z}^d by $L_n^{\text{SRW}}(\alpha, d)$ or simply $L_n^{\text{SRW}}(\alpha)$ when the dimension is clear from the context.

Let R^+ and R^- be, respectively, the semigroup and the group generated by the support of *X*,

$$R^{+} \coloneqq \left\{ x \in \mathbb{Z}^{d} \mid \mathbb{P}\left(S_{n} = x\right) > 0 \text{ for some } n \ge 0 \right\},$$

$$\overline{R} \coloneqq \left\{ x \in \mathbb{Z}^{d} \mid x = y - z \text{ for some } x, y \in R^{+} \right\}.$$
(2)

Following Spitzer [1], we call the random variable X and the random walk it generates *genuinely d-dimensional* if the group \overline{R} is *d*-dimensional.

The quantity $L_n(\alpha)$ has received considerable attention in the literature due to its relation to *self-avoiding walks* and *random walks in random scenery*. In particular let the *random scenery* $\{\xi_x, x \in \mathbb{Z}^d\}$ be a collection of i.i.d. random variables, independent of $(S_n)_n$, and define the process $Z_0 =$ $0, Z_n = \sum_{i=1}^n \xi_{S_i}$. Then $(Z_n)_n$ is commonly referred to as *random walk in random scenery* and was introduced in Kesten and Spitzer [2], where functional limit theorems were obtained for $Z_{[nt]}$ under appropriate normalization for the case d = 1. The case d = 2, with X_i centered with nonsingular covariance matrix, was treated in [3] where it was shown that $Z_{[nt]}/\sqrt{n \log n}$ converges weakly to Brownian motion. As is obvious from the identities $Z_n = \sum_{x \in \mathbb{Z}^d} l(n, x)\xi_x$ and $\operatorname{var}(Z_n) = \operatorname{var}[L_n(2)] \operatorname{var}(\xi_x)$, limit theorems for $(Z_n)_n$ usually require asymptotic results for the local times of the random walk $(S_n)_n$.

Such asymptotic results are usually obtained from Fourier techniques applied to the characteristic function $f(t) = \mathbb{E}[\exp(it \cdot X)]$, under the additional assumption of a Taylor expansion of the form $f(t) = 1 - \langle \Sigma t, t \rangle + o(|t|^2)$, where Σ is a positive definite covariance matrix [3–7], which further requires that $\mathbb{E}|X|^2 < \infty$ and $\mathbb{E}X = 0$. Similar restrictions are also required for the application of local limit theorems such as in [8, 9].

In this paper, motivated by the results of Spitzer [1] for genuinely *d*-dimensional random walks and the approach of Becker and König [10], we will study the asymptotic behavior of var($L_n(\alpha)$) without imposing any moment assumptions on the random walk. The central idea behind our approach is to compare the self-intersection local times $L_n(\alpha)$ of a general *d*-dimensional walk with those of its symmetrised version. In addition we will compare the self-intersection local times of a general *d*-dimensional random walk with those of the *d*-dimensional simple symmetric random walk, $(SRW_n)_{n \in \mathbb{N}_0}$. It is well known that, for some positive constants $K_{\alpha,d}$, $var(L_n^{SRW}(\alpha, d)) \sim K_{\alpha,d}v_{d,\alpha}(n)$ as $n \to \infty$, for

$$v_{1,\alpha}(n) \coloneqq n^{1+\alpha},$$

$$v_{2,\alpha}(n) \coloneqq n^2 \log(n)^{2\alpha-4},$$

$$v_{3,\alpha}(n) \coloneqq n \log(n),$$

$$v_{d,\alpha}(n) \coloneqq n, \quad d \ge 4.$$
(3)

Several other cases have been treated in the literature, using a variety of methods.

A careful look at the literature reveals that the most difficult case in d = 2 is the *near transient recurrent* case, where $\mathbb{P}(S_n = 0) \sim C/n$, which corresponds to genuinely 2-dimensional symmetric recurrent random walks, which will be referred to as a critical case. Surprisingly enough, the variance of the self-intersection local times in the critical case is asymptotically the largest.

Theorem 1. Let $X, X_1, X_2, ...$ be independent, identically distributed, and genuinely d-dimensional \mathbb{Z}^d -valued random variables, for any $d \ge 1$. Then, there exist positive constants $C_{\alpha,X} > c_{\alpha,X} > 0$, depending on α and the distribution of X, such that for all n large enough

$$\operatorname{var}\left(L_{n}\left(\alpha\right)\right) \leq c_{\alpha,X}\operatorname{var}\left(L_{n}^{SRW}\left(\alpha,d\right)\right) \leq C_{\alpha,X}v_{d,\alpha}\left(n\right). \quad (4)$$

The result was motivated by [1, 10] and improves related results of Becker and König for d = 3 and d = 4. Several cases treated in [3, 4, 10–13] can then be obtained as particular cases.

Moreover, we also show the surprising converse. More precisely, we show that the right asymptotic behaviour of $var(L_n)$ implies that the jumps must have zero mean and finite second moment.

Theorem 2. Let $X, X_1, X_2, ...,$ be independent, identically distributed, and genuinely *d*-dimensional with $d \le 3$. If

$$\liminf_{n \to \infty} \frac{\operatorname{var} \left(L_n(\alpha) \right)}{\operatorname{var} \left(L_n^{SRW}(\alpha) \right)} > 0, \tag{5}$$

then $\mathbb{E}|X|^2 < \infty$ and $\mathbb{E}X = 0$.

As it follows from Theorem 3 given below for d = 2, 3and from Theorem 5.2.3 in Chen [12] for d = 1, if $\mathbb{E}X = 0$ and $0 < \mathbb{E}|X|^2 < \infty$, then $\liminf_n \operatorname{var}(L_n(\alpha))/v_{d,\alpha}(n) > 0$.

For any genuinely *d*-dimensional random walk with finite second moments and zero mean, the asymptotic behaviour of $var(L_n(\alpha))$ is similar to that of the *d*-dimensional simple symmetric random walk. Also, as it follows from our general bounds (see Proposition 4 and Corollary 7) that the asymptotic results for the genuinely *d*-dimensional random walk can be reproduced by those of the symmetric *one-dimensional* random walk with appropriately chosen heavy tails, as was indicated by Kesten and Spitzer [2]. The proofs are based on adapting the Tauberian approach developed in [13].

Theorem 3. Let d = 1, 2, 3, and suppose that for $t \in \Gamma := [-\pi, \pi]^d$ one has

$$f(t) = 1 - \gamma |t| + R(t), \text{ for } d = 1,$$

or $f(t) = 1 - \langle \Sigma t, t \rangle + R(t), \text{ for } d = 2, 3,$ (6)

where Σ is a nonsingular covariance matrix and R(t) = o(|t|)for d = 1 and $o(|t|^2)$ for d = 2, 3 as $t \to 0$. Then

$$\operatorname{var}\left(L_{n}\left(\alpha\right)\right) \\ \sim \begin{cases} \frac{\left(\pi^{2}+6\right)}{12} \frac{\left(\alpha!\right)^{2} \left(\alpha-1\right)^{2}}{\left(\gamma\pi\right)^{2\alpha-2}} n^{2} \log\left(n\right)^{2\alpha-4}, & \text{for } d=1, \\ \\ \frac{\left(\alpha!\right)^{2} \left(\alpha-1\right)^{2}}{2 \left(2\pi\sqrt{|\Sigma|}\right)^{2\alpha-2}} n^{2} \log\left(n\right)^{2\alpha-4} \left(\kappa+1\right), & \text{for } d=2, \\ \\ \left(\kappa_{1}+\kappa_{2}\right) n \log n, & \text{for } d=3, \ \alpha=2, \end{cases}$$

$$\tag{7}$$

where

$$\kappa \coloneqq \iint_{0}^{\infty} dr \, ds \left[(1+r) \, (1+s) \, \sqrt{(1+r+s)^2 - 4rs} \right]^{-1} \\ - \frac{\pi^2}{6}, \tag{8}$$

and κ_1 and κ_2 are defined in (58) and (63), respectively.

Moreover, if $L'(n, \alpha)$ is the self-intersection local time of another random walk, independent of $(S_n)_n$, whose characteristic function also satisfies (6), then $\operatorname{var}(L'_n(\alpha)) = \operatorname{var}(L_n(\alpha))(1+$ o(1)).

2. Proofs

2.1. General Bounds. We first develop a technique to treat random walks with independent but not necessarily identically distributed increments.

Proposition 4 (general upper bound). Assume that X_1 , X_2 ,... are independent \mathbb{Z}^d -valued random variables and let $S_{u,v} := X_u + \cdots + X_{u+v}$. Suppose further that for all $n \in \mathbb{N}$ and integers $a, u, b, v \ge 0$, with $a + u \le b$ and any $x \in \mathbb{Z}^d$, one has

$$\mathbb{P}\left(S_{a,u} \pm S_{b,v} = x\right) \le \phi\left(u + v\right), \qquad (A)$$

$$\mathbb{P}\left(S_{a,u}=0\right) - \mathbb{P}\left(S_{a,u}+S_{b,v}=0\right) \le \psi\left(u,v\right),\tag{B}$$

where $\phi(u)$ is nonincreasing and $\psi(u, v)$ is nonincreasing in uand is nondecreasing and subadditive in v in the sense that $\psi(u, v + w) \leq A_{\psi}[\psi(u, v) + \psi(u, w)]$, for some constant A_{ψ} independent of u, v, and w. Then, for some constant $K = cA_{\psi}(1 + A_{\psi})^{\alpha-2}$ depending only on α

$$\operatorname{var}\left(L_{n}\left(\alpha\right)\right) \leq Kn\left(\sum_{i=0}^{n-1}\phi\left(i\right)\right)^{2\alpha-4}$$

$$\cdot \sum_{i,j,k=0}^{n-1}\left[\phi\left(j\vee i\right)\phi\left(k\vee i\right) + \phi\left(j\right)\psi\left(i+k,j\right)\right].$$
(9)

Proof of Proposition 4. We first write out the variance as a sum

 $\operatorname{var}(L_n(\alpha)) = (\alpha!)^2$

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$$\cdot \sum_{k_1 \leq \dots \leq k_{\alpha}} \sum_{l_1 \leq \dots \leq l_{\alpha}} \left(\mathbb{P} \left[S_{k_1} = \dots = S_{k_{\alpha}}, S_{l_1} = \dots = S_{l_{\alpha}} \right] \right.$$

$$- \mathbb{P} \left[S_{k_1} = \dots = S_{k_{\alpha}} \right] \mathbb{P} \left[S_{l_1} = \dots = S_{l_{\alpha}} \right] \right).$$

$$(10)$$

An important role is played by the manner in which the two sequences are interlaced, since, for example, if $k_{\alpha} \leq l_1$ or $l_{\alpha} \leq k_1$, the term vanishes by the Markov property.

We will treat the sum over indices with $k_1 \leq l_1$. The sum over the remaining index set with $k_1 > l_1$ can be treated in a similar fashion and will contribute a constant factor. Therefore, we assume that $k_1 \leq l_1$ and we arrange the two sequences in an ordered sequence of combined length 2α which we denote as $(p_1, \ldots, p_{2\alpha})$; we also define $(\epsilon_1, \ldots, \epsilon_{2\alpha})$ where $\epsilon_i = 0$ if p_i came from $\mathbf{k} \coloneqq \{k_1, \ldots, k_{\alpha}\}$ and $\epsilon_i = 1$ if p_i came from $\mathbf{l} \coloneqq \{l_1, \ldots, l_{\alpha}\}$. Finally we define two new sequences $m_0, m_1, \ldots, m_{2\alpha-1}$, and $\delta_1, \ldots, \delta_{2\alpha-1}$, where $m_0 \coloneqq$ $p_1, m_i = p_{i+1} - p_i$, and $\delta_i = \epsilon_{i+1} - \epsilon_i$, for $i = 1, \ldots, 2\alpha - 1$. Notice that since we assume that $k_1 \leq l_1$, we have $p_1 = k_1$ and $\epsilon_1 = 0$. Let $v(\delta) \coloneqq \sum_{i=1}^{2\alpha-1} |\delta_i|$ denote the *interlacement index*. The terms with v = 1 vanish, while the terms with v = 2 will be considered separately.

Terms with $v \ge 3$. We first consider the sum I_n over the terms with $v \ge 3$ for which we drop the negative part and obtain the bound

$$I_{n} \coloneqq \sum_{\substack{k_{1} \leq \dots \leq k_{\alpha} \\ l_{1} \leq \dots \leq l_{\alpha} \\ k_{1} \leq l_{1}, v(\delta) \geq 3}} \mathbb{P} \left[S_{k_{1}} = \dots = S_{k_{\alpha}}, S_{l_{1}} = \dots = S_{l_{\alpha}} \right]$$

$$= \sum_{x, y \in \mathbb{Z}^{d}} \sum_{p_{1} \leq \dots \leq p_{2\alpha} \leq n} \sum_{\epsilon: v(\delta) \geq 3} \mathbb{P} \left[S_{p_{1}} = x, S_{p_{2}} = x + \epsilon_{2}y, \dots, S_{p_{2\alpha}} = x + \epsilon_{2\alpha}y \right]$$

$$\leq \sum_{x, y \in \mathbb{Z}^{d}} \sum_{m_{0}, \dots, m_{2\alpha-1} \leq n} \sum_{\delta: v(\delta) \geq 3} \mathbb{P} \left(S_{m_{0}} = x \right) \mathbb{P} \left(S_{m_{0}, m_{1}} = \delta_{1}y \right) \cdots \mathbb{P} \left(S_{m_{2\alpha-2}, m_{2\alpha-1}} = \delta_{2\alpha-1}y \right)$$

$$= \sum_{y \in \mathbb{Z}^{d}} \sum_{m_{0}, \dots, m_{2\alpha-1} \leq n} \sum_{\delta: v(\delta) \geq 3} \mathbb{P} \left(S_{m_{0}, m_{1}} = \delta_{1}y \right) \cdots \mathbb{P} \left(S_{m_{2\alpha-2}, m_{2\alpha-1}} = \delta_{2\alpha-1}y \right).$$
(11)

Summing over the free index m_0 , it is clear that

$$I_n \le (n+1)$$

$$\cdot \sum_{m_1,\dots,m_{2\alpha-1}} \sum_{y \in \mathbb{Z}^d} \sum_{\delta: v(\delta) \ge 3} \prod_{t=1}^{2\alpha-1} \sup_{w} \mathbb{P}\left(S_{w,m_t} = \delta_t y\right).$$
⁽¹²⁾

For any $\delta = (\delta_1, \dots, \delta_{2\alpha-1})$ with $v(\delta) = v$, exactly $u \coloneqq 2\alpha - 1 - v$ elements are equal to 0, and therefore by Assumption (A) with x = 0 we have

$$I_{n} \leq C(n+1) \sum_{\nu=3}^{\alpha} \left[\sum_{i=0}^{n} \phi(i) \right]^{2\alpha-1-\nu}$$

$$\cdot \sum_{j_{1},\dots,j_{\nu}=0}^{n} \sum_{y \in \mathbb{Z}^{d}} \sum_{\delta' \in \{-1,+1\}^{\nu}} \prod_{t=1}^{\nu} \sup_{w_{t}} \mathbb{P}\left(S_{w_{t},j_{t}} = \delta_{t} y \right).$$

$$(13)$$

Letting $(\tilde{S}_n)_{n \in \mathbb{N}_0}$ denote an independent copy of the random walk $(S_n)_{n \in \mathbb{N}_0}$ and assuming without loss of generality that $j_1 \leq \cdots \leq j_{\nu}$, we have that for any $\delta \in \{-1, +1\}^{\nu}$

$$\sum_{y \in \mathbb{Z}^{d}} \prod_{t=1}^{\nu} \sup_{w_{t}} \mathbb{P} \left(S_{w_{t}, j_{t}} = \delta_{t} y \right)$$

$$\leq \left(\prod_{t=2}^{\nu-1} \sup_{y} \sup_{w_{t}} \mathbb{P} \left(S_{w_{t}, j_{t}} = y \right) \right)$$

$$\cdot \sup_{w_{1}, w_{\nu}} \mathbb{P} \left(S_{w_{1}, j_{1}} - \delta_{\nu} \widetilde{S}_{w_{\nu}, j_{\nu}} = 0 \right) \leq \left[\prod_{t=2}^{\nu-1} \phi \left(j_{t} \right) \right]$$

$$\cdot \phi \left(j_{1} + j_{\nu} \right) \leq \prod_{t=2}^{\nu} \phi \left(j_{t} \lor j_{1} \right).$$
(14)

Let $G_n := \sum_{i=0}^n \phi(i)$. Since ϕ is nonincreasing we have that

$$\Delta_{n,v} := \sum_{0 \le j_1 \le \dots \le j_v \le n} \prod_{t=2}^{v} \phi(j_t \lor j_1)$$

$$\leq \sum_{j_v=0}^{n} \phi(j_v) \sum_{0 \le j_1 \le \dots \le j_{v-1} \le n} \prod_{t=2}^{v-1} \phi(j_t \lor j_1)$$

$$= G_n \Delta_{n,v-1},$$
 (15)

and iterating this procedure, for $v \ge 3$, we have that $\Delta_{n,v} \le \Delta_{n,3}G_n^{\nu-3}$. Combining the two bounds and summing over $v = 3, \ldots, 2\alpha - 1$, we have that

$$I_{n} \leq \sum_{\nu=3}^{2\alpha-1} c(\alpha) n G_{n}^{2\alpha-1-\nu} \Delta_{n,\nu} \leq c(\alpha) n G_{n}^{2\alpha-1-\nu+\nu-3} \Delta_{n,3}$$

= $c(\alpha) n G_{n}^{2\alpha-4} \Delta_{n,3},$ (16)

where $c(\alpha)$ is a constant depending only on α .

Terms with v = 2. Next we consider the sum J_n over the terms with v = 2, which occurs when, for some j, the indices l_1, \ldots, l_{α} all lie in $[k_j, k_{j+1}]$. Then it is easy to see that this sum J_n is bounded above by

$$J_{n} \leq Cn \sup_{w_{0},...,w_{2\alpha-1}} \sum_{m_{\alpha+1},...,m_{2\alpha-2}=0}^{n} \prod_{r=\alpha+1}^{2\alpha-2} \mathbb{P}\left(S_{w_{r},m_{r}}=0\right)$$

$$\cdot \sum_{m_{0},...,m_{\alpha}=0}^{n} \left[\prod_{t=1}^{\alpha-1} \mathbb{P}\left(S_{w_{t},m_{t}}=0\right)\right] \left[\mathbb{P}\left(S_{w_{0},m_{0}}+S_{w_{\alpha},m_{\alpha}}\right) - \mathbb{P}\left(S_{w_{0},m_{0}}+\cdots+S_{w_{\alpha},m_{\alpha}}=0\right)\right] \leq CnG_{n}^{\alpha-2}$$

$$\cdot \sup_{w_{0},...,w_{\alpha}} \sum_{m_{0},...,m_{\alpha}=0}^{n} \left[\prod_{t=1}^{\alpha-1} \mathbb{P}\left(S_{w_{t},m_{t}}=0\right)\right]$$

$$\cdot \left[\mathbb{P}\left(S_{w_{0},m_{0}}+S_{w_{\alpha},m_{\alpha}}=0\right) - \mathbb{P}\left(S_{w_{0},m_{0}}+\cdots+S_{w_{\alpha},m_{\alpha}}=0\right)\right] \qquad (17)$$

$$\leq CnG_{n}^{\alpha-2} \sum_{m_{0},...,m_{\alpha}=0}^{n} \left[\prod_{t=1}^{\alpha-1} \phi\left(m_{t}\right)\right] \psi\left(m_{0}+m_{\alpha},m_{1}+\cdots+m_{\alpha-1}\right) \leq C\alpha nG_{n}^{\alpha-2}A_{\psi}\left(1+A_{\psi}\right)^{\alpha-2}$$

$$\cdot \left(\sum_{m_{2},...,m_{\alpha-1}}\prod_{t=2}^{\alpha-1} \phi\left(m_{t}\right)\right) \sum_{m_{0},m_{1},m_{\alpha}} \phi\left(m_{1}\right) \psi\left(m_{0}+m_{\alpha},m_{1}+k,j\right).$$

The following corollary provides explicit bounds in the cases that are usually considered in the literature.

Corollary 5. Assume that the conditions of Proposition 4 are satisfied with $\phi(m) = Tm^{-r}$ and $\psi(m, k) = Tm^{-r-1}(k \wedge m)$. Then,

$$\leq c_{\alpha} T^{2\alpha-2} \begin{cases} n^{2} \log (n)^{2\alpha-4}, & if \ r = 1, \\ n^{4-2r}, & if \ 1 < r < \frac{3}{2}, \\ n \log (n), & if \ r = \frac{3}{2}, \\ n, & if \ r > \frac{3}{2}. \end{cases}$$
(18)

It is straightforward to see that Corollary 5 includes random walks with mean zero and finite second moment; for example, d = 2 corresponds to r = 1 and d = 3to r = 3/2. Therefore several relevant results in [3, 7–13] are obtained as a special case of Corollary 5 and extended to the case of independent but not necessarily identically distributed variables, for example, by applying the local limit theorem, as conducted in [8].

Also when the random walk increment *X* is in the domain of attraction of the one-dimensional symmetric Cauchy law [13, 14] or in the case of planar random walk with second moments [3, 7–9, 11], it is well known that the conditions of Proposition 4 are satisfied with $\phi(m) = T/m$ and $\psi(m, k) = Tm^{-2}(k \wedge m)$.

However, we can do better for symmetric variables and show that condition (A) implies (B), which together with the comparison technique motivates the following results. For a real number x, we write [x] for the integer part of x.

Proposition 6 (bounds via comparison with characteristic function of symmetric random variables). Let X_1, X_2, \ldots , be independent \mathbb{Z}^d -valued random variables and let $f_i(t) := \mathbb{E} \exp(itX_i)$. Assume that there exist a measurable function $f: \Gamma \to [0, 1]$ and a positive nonincreasing sequence $(\phi(m))_{m \in \mathbb{N}_0}$, such that

$$|1 - f_i(t)| \le Tf(t),$$

$$|f_i(\pm t)| \le f(t),$$

$$\int_{\Gamma} f(t)^m dt \le \phi(m),$$
(19)

for all integers $i, m \ge 0$, all $t \in \Gamma$, and some positive constant T. Then there exists another positive constant $K = c(\alpha, d, T)$ such that

$$\operatorname{var}\left(L_{n}\left(\alpha\right)\right) \leq Kn\left(\sum_{i=0}^{n-1}\phi\left(\left[\frac{i}{2}\right]\right)\right)^{2\alpha-4}\sum_{j=0}^{n}j\phi\left(\left[\frac{j}{2}\right]\right)\sum_{k=j}^{2n}\phi\left(\left[\frac{k}{2}\right]\right).$$
⁽²⁰⁾

Proof of Proposition 6. Using the notation of Proposition 4, for positive integers *a*, *u*, *b*, and *v*, with $a + u \le b$, $\epsilon_j = \pm 1$, and any $x \in \mathbb{Z}^d$

$$\mathbb{P}\left(S_{a,u} + \epsilon \cdot S_{b,v} = x\right)$$

$$\leq \frac{1}{(2\pi)^d} \int_{\Gamma} \prod_{j \in [a,a+u] \cup [b,b+v]} \left| f_j\left(\epsilon_j t\right) \right| dt \qquad (21)$$

$$\leq \frac{1}{(2\pi)^d} \int_{\Gamma} f\left(t\right)^{u+v} dt \leq \frac{1}{(2\pi)^d} \phi\left(u+v\right).$$

To find $\psi(u, v)$, notice that since $f(t) \ge 0$,

$$\phi(m) \ge \int_{\Gamma} f(t)^{m} [1 - f(t)^{m}] dt$$

= $\sum_{j=0}^{m-1} \int_{\Gamma} f(t)^{m+j} (1 - f(t)) dt$ (22)
 $\ge m \int_{\Gamma} f(t)^{2m} (1 - f(t)) dt =: mQ(2m)$

whence $Q(m) \leq 2\phi([m/2])/m$. Therefore,

$$\left| \mathbb{P} \left(S_{a,u} = 0 \right) - \mathbb{P} \left(S_{a,u} + S_{b,1} = 0 \right) \right|$$

$$= \left| \frac{1}{(2\pi)^d} \int_{\Gamma} \left[\prod_{j=a}^{a+u} f_j(t) \right] \left(1 - f_{b+1}(t) \right) dt \right| \qquad (23)$$

$$\leq CT \int_{\Gamma} \left| f(t) \right|^u \left| 1 - f(t) \right| dt \leq \frac{CT\phi\left([u/2] \right)}{u}.$$

A telescoping argument implies that

$$\left|\mathbb{P}\left(S_{a,u}=0\right)-\mathbb{P}\left(S_{a,u}+S_{b,v}=0\right)\right| \le CT\phi\left(\left[\frac{u}{2}\right]\right)\frac{v}{u}.$$
 (24)

On the other hand for $u \leq v$ we can obtain a tighter bound through

$$\mathbb{P}\left(S_{a,u}=0\right) - \mathbb{P}\left(S_{a,u}+S_{b,v}=0\right) \le \mathbb{P}\left(S_{a,u}=0\right)$$

$$\le \phi\left(u\right).$$
(25)

Combining the two bounds above it follows that (B) is satisfied with $\psi(u, v) \coloneqq \phi([u/2]) \min(u, v)/u$. Thus all conditions of Proposition 4 are satisfied and the result follows.

The following corollary allows for the case where $\phi(m)$ is regularly varying.

Corollary 7. Assume that the conditions of Proposition 6 are satisfied with $\phi(m) = h(m)m^{-r}$, $r \ge 1$, where $h(\cdot)$ is slowly varying at ∞ . Then,

$$\operatorname{var}(L_{n}(\alpha)) \leq K\Delta_{n}(\alpha,\phi) \\ \leq c_{\alpha}T^{2\alpha-2} \begin{cases} n^{2} \left[\sum_{k=1}^{n} \frac{h(k)}{k}\right]^{2\alpha-4}, & \text{for } r = 1, \\ n^{4-2r}h^{2}(n), & \text{for } 1 < r < \frac{3}{2}, \\ n\sum_{k=1}^{n} \frac{h(k)^{2}}{k}, & \text{for } r = \frac{3}{2}, \\ n, & \text{for } r > \frac{3}{2}. \end{cases}$$
(26)

Several results in [3, 7–13] are obtained as a special case of Corollary 7 and can be extended to dependent variables, for example, a random walk driven by a hidden Markov chain. In addition, following [2], we can construct a one-dimensional symmetric random walk with characteristic function $f(t) = 1 - c|t|^{1/r} + o(|t|^{1/r})$, where r = 2/d for d = 2, 3 and r = 1/2 for $d \ge 4$, whose asymptotic behaviour is similar to that of genuinely *d*-dimensional random walk.

The following example of genuinely 2-dimensional recurrent walk with infinite variance was motivated by Spitzer [1, pp. 87].

Example 8. Let X_1, X_2, \ldots be independent, identically distributed, \mathbb{Z}^2 -valued random variables, such that $\mathbb{P}(|X_1| = k) = c/(k^3 \log(k)^g)$, for $k \ge 4$ and $g \in [0, 1)$. Let $(S_n)_{n \in \mathbb{N}_0}$ be the corresponding random walk in \mathbb{Z}^2 . Then we have

$$\operatorname{var}\left(L_{n}\left(\alpha\right)\right) \leq cn^{2} \max\left\{\left[\log n\right]^{g}, \log\log n\right\}^{2\alpha-4} \log n^{-2(1-g)},$$

$$(27)$$

for $n \ge 10$. Under these assumptions we have that $\mathbb{P}(S_n = 0) \le c/n \log(n)^{1-g}$, which is in the *critical range*, where the random walk is recurrent, without second moment. To see why, we note that by a lengthy but straightforward calculation it can be shown that the characteristic function of X satisfies (19) with

$$\phi(n) = \frac{c}{n\log(e \vee n)^{1-g}},$$

$$f(t) = \exp\left[-A|t|^2 h\left(|t|^2\right)\right],$$
(28)
where $h(r) \coloneqq \left[1 + \log\left(\frac{1}{r}\right)_+\right]^{1-g}.$

The sequence $\phi(m)$ is identified via Fourier inversion, polar coordinates, and a Laplace argument,

$$\int_{\Gamma} f(t)^{n} dt \leq c \int_{0}^{1} \exp\left(-nr\left(1+\log\left(\frac{1}{r}\right)\right)^{1-g}\right) + O\left(e^{-n}\right) \leq \frac{c}{n\log\left(e \vee n\right)^{1-g}} =: \phi(n).$$
(29)

2.2. Bounds for Identically Distributed Variables

Proposition 9 (general upper bound for i.i.d.). Let X, X_1, X_2, \ldots , be independent, identically distributed,

 \mathbb{Z}^d -valued random variables. Suppose that for any $x \in \mathbb{Z}^d$ and all positive integers $a, u, b, and v, with <math>a + u \leq b$, it holds that

$$\mathbb{P}\left(S_{a,u} \pm S_{b,v} = x\right) \le \phi\left(u + v\right),\tag{30}$$

where $\{\phi(m)\}_{m \in \mathbb{N}_0}$ is a nonincreasing sequence. Then for some constant $K = c(\alpha)$ we have that

$$\operatorname{var}\left(L_{n}\left(\alpha\right)\right) \leq Kn\left(\sum_{i=0}^{n-1}\phi\left(i\right)\right)^{2\alpha-4}\sum_{j=0}^{n}j\phi\left(j\right)\sum_{k=j}^{\left[\alpha n\right]+1}\phi\left(\left[\frac{k}{\alpha}\right]\right).$$
(31)

Proof of Proposition 9. By inspecting the proof of Proposition 6, we notice that we only need to bound the term J_n . Consider typical ordering

$$0 \le i_1 \le \dots \le i_k \le j_1 \le \dots \le j_\alpha \le i_{k+1} \le \dots \le i_\alpha \le n, \quad (32)$$

and let us change variables to $(m_0, ..., m_{2\alpha})$ such that $m_0 + \cdots + m_{2\alpha} = n$. Then the contribution to J_n is given by

$$\sum_{\substack{m_0,\dots,m_{2\alpha}\\1\leq j\leq 2\alpha-1}} \prod_{\substack{j\neq k,k+\alpha\\1\leq j\leq 2\alpha-1}} \mathbb{P}\left(S_{m_j}=0\right)$$

$$\cdot \left[\mathbb{P}\left(S_{m_k+m_{k+\alpha}}=0\right) - \mathbb{P}\left(S_{m_k+\dots+m_{k+\alpha}}=0\right)\right].$$
(33)

We keep m_j fixed for $j \neq \alpha, k + \alpha$ and we sum over $m = m_k + m_{k+\alpha}$ from 0 to some $M = M(n, \{m_j\}_{j \neq k, k+\alpha})$. Then for given $m_{k+1}, \ldots, m_{k+\alpha-1}$, the term in the sum is

$$\sum_{m=0}^{M} (m+1) \left[\mathbb{P} \left(S_m = 0 \right) - \mathbb{P} \left(S_{m+q} = 0 \right) \right], \qquad (34)$$

where $q \coloneqq m_{k+1} + \cdots + m_{k+\alpha-1}$. Then since $M \le n-q$, it is an easy exercise to show that this sum is bounded above by

$$\sum_{m=0}^{M} (m+1) \left[\mathbb{P} \left(S_{m} = 0 \right) - \mathbb{P} \left(S_{m+q} = 0 \right) \right]$$

$$\leq \sum_{m=0}^{q-1} (m+1) \mathbb{P} \left(S_{m} = 0 \right) + q \mathbb{1} (n-q \ge q)$$

$$\cdot \sum_{m=q}^{n-q} \mathbb{P} \left(S_{m} = 0 \right) \le \sum_{m=0}^{(\alpha m^{*}) \wedge n} (m+1) \mathbb{P} \left(S_{m} = 0 \right)$$

$$+ \alpha m^{*} \sum_{m=m^{*}}^{n} \mathbb{P} \left(S_{m} = 0 \right),$$
(35)

where $m^* = \max\{m_{k+1}, \ldots, m_{k+\alpha-1}\}$. The result follows by summing over all indices apart from m^* and changing the order of summation.

2.3. Proofs of Main Results

Proof of Theorem 1. We apply a comparison argument found to be useful in many areas (e.g., Montgomery-Smith and Pruss [15], and Lefèvre and Utev [16]). More specifically we

bound the quantity $var(L_n)$ by the corresponding quantity for the symmetrised random walk.

Following Spitzer's argument we notice that with $f(t) = \mathbb{E}[\exp(it \cdot X_1)]$

$$\mathbb{P}\left(S_{a,u} + \epsilon S_{b,v} = x\right) \le c \int_{\Gamma} \left|f\left(t\right)\right|^{u} \left|f\left(-t\right)\right|^{v} dt$$

$$= c \int_{\Gamma} \left[\left|f\left(t\right)\right|^{2}\right]^{u/2} \left[\left|f\left(-t\right)\right|^{2}\right]^{v/2} dt.$$
(36)

Since $|f(t)|^2$ is the characteristic function of a symmetric random variable in \mathbb{Z}^d , for some positive λ , we have $1 - |f(t)|^2 \ge \lambda |t|^2$, and, hence,

$$\mathbb{P}\left(S_{a,u} + \epsilon S_{b,v} = x\right) \le c \int_{\Gamma} \exp\left[-\frac{\lambda\left(u+v\right)}{2}\left|t\right|^{2}\right] dt$$

$$\le c \left(u+v\right)^{-d/2}.$$
(37)

The result follows from Proposition 9 applied with $\phi(m) = m^{-d/2}$.

The proof of Theorem 2 will be based on the following lemma.

Lemma 10. Assume $X, X_1, X_2, ...$ are independent, identically distributed, genuinely d-dimensional random variables such that $\mathbb{E}|X|^2 = \infty$. Then there exists a monotone, slowly varying sequence $(h_n)_{n \in \mathbb{N}_0}$, such that $h_n \to 0$ as $n \to \infty$ and

$$\sup_{x \in \mathbb{Z}^d} \mathbb{P}\left(S_n = x\right) \le c_d \int_{\Gamma} \left|\mathbb{E}e^{it \cdot X}\right|^n \mathrm{d}t \le h_n n^{-d/2}.$$
 (38)

Proof of Lemma 10. Without loss of generality we assume that *X* is symmetric. Let $\sigma_{e,L} := \mathbb{E}[(e \cdot X)^2 \mathbb{1}(|X| \leq L)]$. Following Spitzer, since *X* is genuinely *d*-dimensional, we may assume that there exist positive constants *c*, *W*, such that for any unit vector |e| = 1 we have that $\sigma_{e,W} \geq c$ and $1 - f(t) \geq c|t|^2$ for all $t \in \Gamma$. Let λ_d be the *d*-dimensional Lebesgue measure on \mathbb{R}^d and μ_d the Lebesgue-Haar measure on $S^{d-1} := \{e \in \Gamma : |e| = 1\}$. Notice that since $\mathbb{E}|X|^2 = \infty$, for any *K*, we have $\mu_d \{e : \sigma_{e,\infty} < K\} = 0$.

Fix a small positive *x* such that $\sqrt{c/x} \ge 2W$, and for any $\epsilon > 0$ let $K = K(\epsilon) = \epsilon^{-d/2}$. Then there exists $L = L(\epsilon) > 0$ small enough so that $\mu_d \{e : \sigma_{e,L} < K\} \le \epsilon^{d/2}$. We partition S^{d-1} in two sets

$$A_{L,K} = \{ e \in S_d : \sigma_{e,L} \ge K \},$$

$$\overline{A}_{L,K} = \{ e \in S_d : \sigma_{e,L} < K \},$$
(39)

so that, for any direction $e \in \overline{A}_{L,K}$,

$$\{z \in \mathbb{R} : 1 - f(ze) \le x\} \subseteq \{z : cz^2 \le x\}$$

$$\subseteq \{z : |z| \le \sqrt{\frac{x}{c}}\}.$$
(40)

Hence, using *d*-dimensional spherical coordinates,

$$\lambda_{d}\left\{(z,e) \in \mathbb{R} \times \overline{A}_{L,K} : 1 - f(ez) \le x\right\}$$

$$\le \mu_{d}\left\{\overline{A}_{L,K}\right\}\left(\frac{x}{c}\right)^{d/2}\left(\frac{1}{d}\right) \le \epsilon^{d/2}\left(\frac{x}{c}\right)^{d/2}\left(\frac{1}{d}\right).$$
(41)

On the other hand, for any *t*,

$$1 - f(t) = 2 \sum_{k \in \mathbb{Z}^d} \sin\left(\frac{[t \cdot k]}{2}\right)^2 P(X = k)$$

$$\geq \left(\frac{1}{4}\right) E\left[(t \cdot X)^2 I\left(|t \cdot X| \le \frac{1}{2}\right)\right] \qquad (42)$$

$$= \left(\frac{|t|^2}{4}\right) \sigma_{t/|t|, 1/2|t|}.$$

Now, assume that $\sqrt{c/x} \ge 2L$. Then for any direction $e \in A_{L,K}$, by choice of *x* and since $\sigma_{e,L}$ is increasing in *L*, for $cz^2 \le 1 - f(ez) \le x$ or $|z| \le \sqrt{x/c}$, it must be the case that

$$x \ge 1 - f(ez) \ge \left(\frac{z^2}{4}\right) \sigma_{e,1/2z} \ge \left(\frac{z^2}{4}\right) \sigma_{e,L}$$

$$\ge \left(\frac{z^2}{4}\right) K,$$
(43)

implying that, on the set $A_{L,K}$, it must be that $|z| \le 2\sqrt{x/K}$. Changing to *d*-dimensional polar coordinates, we find that

$$\lambda_d \left\{ (z, e) \in \mathbb{R} \times A_{L,K} : 1 - f(ez) \le x \right\}$$

$$\le \int_{A_{L,K}} \int_0^{\sqrt{4x/K}} r^{d-1} dr \, de \le C_d \epsilon^{d/2} x^{d/2}.$$
 (44)

Overall, for $x \le c/4L^2$, $\lambda_d \{t : 1 - f(t) \le x\} \le c_d(x\epsilon)^{d/2}$, and hence $\{t \in \Gamma : 1 - f(t) \le x\}$ has Lebesgue measure $o(x^{d/2})$.

Let F(x) be the cumulative distribution function of the random variable $\log(1/f(\cdot))$ defined on the probability space Γ with normalised Lebesgue measure. Then F is continuous at x = 0 and supported on \mathbb{R}^+ . Moreover, we have that $F(x) = o(x^{d/2})$ as $x \downarrow 0$. Therefore, for some positive sequence $(\epsilon_n)_{n \in \mathbb{N}_0}$ with $\epsilon_n \to 0$, we have that

$$\frac{1}{(2\pi)^2} \int_{\Gamma} f(t)^n dt = \int_0^\infty e^{-nx} dF(x)$$

$$= n \int_0^\infty e^{-nx} F(x) dx \le n^{-d/2} \epsilon_n.$$
(45)

It remains to show that there exists a positive, monotone, slowly varying sequence $(h_n)_{n \in \mathbb{N}_0}$, such that $\epsilon_n \leq h(n) \to 0$ as $n \to \infty$. Let $\delta_n = \sup_{j \geq n} \epsilon_j$ and $a_0 \coloneqq 0$ and for $n \geq 1$ define a_n recursively by $a_n = \min(2a_{2^{r-1}}, 1/\delta_n)$, for $2^{r-1} < n \leq 2^r$, so that $a_n \to \infty$ is monotone, $a_{2^r} \leq 2a_{2^{r-1}}$ implying that $a_{2n} \leq 4a_n$, and $1/a_n \geq \delta_n \geq \epsilon_n$. Finally, take $h_n \coloneqq 1/\max(a_0, \log a_n)$.

Proof of Theorem 2. Assume that $\mathbb{E}|X|^2 = \infty$ and d = 2 or d = 3. Then, by Lemma 10 there exists a slowly varying sequence $h_n \to 0$ as $n \to \infty$ such that $\int_{\Gamma} |\mathbb{E} \exp(it \cdot X)|^n dt \le h_n n^{-d/2}$. Applying Corollary 7 with r = 1 and r = 3/2 we, respectively, find that

$$\operatorname{var}\left(L_{n}\left(\alpha\right)\right) \leq \begin{cases} Kn^{2}\left(\sum_{k=1}^{n}\frac{h\left(k\right)}{k}\right)^{2\alpha-4} = o\left(n^{2}\left(\log n\right)^{2\alpha-4}\right), & \text{for } d = 2, \\ Kn\left(\sum_{k=1}^{n}\frac{h\left(k\right)^{2}}{k}\right) = o\left(n\ln n\right), & \text{for } d = 3. \end{cases}$$

$$(46)$$

Finally assume that $\mathbb{E}|X|^2 < \infty$ and $E[X] = \mu \neq 0$. Then $\mathbb{P}(S_n = 0) = \mathbb{P}(S'_n = -n\mu)$ whence it follows that $\mathbb{P}(S_n = 0) = o(n^{-d/2})$ (see, e.g., [17, Theorem 2.3.10]). Then inspecting the proof of Proposition 4, one can readily obtain the desired bound for the J_n term, while with slight modification the bound for the I_n term also follows.

Note that for d = 1 the situation is much simpler since then $\operatorname{var}(L_n^{\mathsf{SRW}}(\alpha)) \sim C[\mathbb{E}L_n^{\mathsf{SRW}}(\alpha, d)]^2$ and if $\mathbb{E}|X|^2 = \infty$ or $\mathbb{E}[X] \neq 0$, $\mathbb{E}L_n^{\mathsf{SRW}}(\alpha, d) = o(n^{(1+\alpha)/2})$.

Proof of Theorem 3. We first give the proof for the case d = 1. As in the proof of Proposition 4 we begin from expression (10) and define the sequences p_i and δ_i for $i = 1, ..., 2\alpha - 1$, and the quantity $v(\delta) = \sum_{i=1}^{2\alpha-1} |\delta_i|$. Recall that $v(\delta)$ measures the interlacement of the two sequences $k_1, ..., k_\alpha$ and $l_1, ..., l_\alpha$. For example, $v(\delta) = 1$ occurs when either $k_\alpha \leq l_1$ or $l_\alpha \leq k_1$, in which case the contribution vanishes by the Markov property. On the other hand $v(\delta) = 2$ when, for example, $l_1, ..., l_\alpha \in [k_i, k_{i+1}]$ for some *i*. Finally $v(\delta) = 3$ occurs when, for example,

$$k_{1} \leq \dots \leq k_{r} \leq l_{1} \leq \dots \leq l_{s} \leq k_{r+1} \leq \dots \leq k_{\alpha} \leq l_{s+1}$$

$$\leq \dots \leq l_{\alpha} \leq n.$$
(47)

From the proof of Proposition 4, and using the bound $\mathbb{P}(S_n = 0) \leq c/n$, the terms of the sum are bounded above by $n^2 \log(n)^{2\alpha-1-\nu(\delta)}$, and thus the leading term appears when either $\nu(\delta) = 2, 3$, with other terms giving strictly lower order. We will therefore analyze these two situations in detail in order to derive the exact asymptotic constants. When $\nu = 3$, the two terms in the difference individually give the correct order and will be treated by the classical Tauberian theory. However for $\nu = 2$, the two terms only give the correct order when considered together. This however forbids the use of Karamata's Tauberian theorem since the monotonicity restriction would require roughly that X_i is symmetric. Thus the complex Tauberian approach, as developed in [13], is required to justify the answer.

Case 1 ($v(\delta) = 3$). Assume that part of the sequence $\mathbf{l} = \{l_1, \ldots, l_{\alpha}\}$ lies between k_r and k_{r+1} and the rest between k_s and k_{s+1} . Then using the change of variables

$$i_{1} = m_{0},$$

$$i_{2} = m_{0} + m_{1},$$

$$\vdots$$

$$i_{r} = m_{0} + \dots + m_{r-1}$$

$$j_{1} = m_{0} + \dots + m_{r},$$

$$j_{2} = m_{0} + \dots + m_{r+1},$$

$$\vdots$$

$$j_{s} = m_{0} + \dots + m_{r+s-1},$$

$$i_{r+1} = m_{0} + \dots + m_{r+s+1},$$

$$\vdots$$

$$i_{\alpha} = m_{0} + \dots + m_{\alpha+s-1},$$

$$j_{s+1} = m_{0} + \dots + m_{\alpha+s+1},$$

$$\vdots$$

$$j_{s+2} = m_{0} + \dots + m_{\alpha+s+1},$$

$$\vdots$$

$$j_{\alpha} = m_{2\alpha-1},$$

$$n = m_{0} + \dots + m_{2\alpha},$$
(48)

we rewrite the positive term in (10) as

a(n)

$$= \sum \mathbb{P} \left[S(i_{1}) = \dots = S(i_{\alpha}); S(j_{1}) = \dots = S(j_{\alpha}) \right]$$

$$= \sum_{m_{0},\dots,m_{2\alpha-1}} \left[\prod_{\substack{j=1\\ j \neq r,r+s,\alpha+s}}^{2\alpha-1} \mathbb{P} \left(S_{m_{j}} = 0 \right) \right]$$

$$\cdot \mathbb{P} \left(S_{m_{r}} + S'_{m_{r+s}} = S'_{m_{r+s}} + S''_{m_{\alpha+s}} = 0 \right).$$
(49)

Notice that from [13] we have that $\sum_{n\geq 0} \lambda^n \mathbb{P}(S_n = 0) \sim \log(1/(1-\lambda))/\pi\gamma$. Let

$$a(\lambda) = (1 - \lambda)^{-3} \left[-\log(1 - \lambda) \right]^{2\alpha - 4},$$

$$c_{\gamma} = (\pi \gamma)^{-2\alpha + 4}.$$
(50)

Then, by direct calculations and Fourier inversion formula

$$\sum_{n\geq 0} \lambda^{n} a(n) = c_{\gamma} (1-\lambda) a(\lambda)$$

$$\cdot \sum_{x\in\mathbb{Z}} \sum_{k_{1},k_{2},k_{3}\geq 0} \lambda^{k_{1}+k_{2}+k_{3}} \mathbb{P} \left(S_{k_{1}} = x\right) \mathbb{P} \left(S_{k_{2}} = -x\right)$$

$$\cdot \mathbb{P} \left(S_{k_{3}} = x\right) = c_{\gamma} (1-\lambda) a(\lambda) \frac{1}{(2\pi)^{2}}$$

$$\cdot \iint_{\Gamma} \frac{dt \, ds}{(1-\lambda f(t)) (1-\lambda f(s)) (1-\lambda f(t+s))}$$

$$\cdot c_{\gamma} (1-\lambda) a(\lambda) \frac{1}{(2\pi)^{2} \gamma^{2}} \frac{1}{1-\lambda}$$

$$\cdot \iint_{\mathbb{R}^{2}} \frac{dx \, dy}{(1+|x|) (1+|y|) (1+|x+y|)} \sim \left(\frac{1}{4\gamma^{2}}\right)$$

$$\cdot c_{\gamma} a(\lambda).$$
(51)

Next we consider the negative term in (10)

$$b(n) := \sum_{m_0, \dots, m_{2\alpha-1}} \mathbb{P} \left[S_{m_1} = \dots = S_{m_{r-1}} = S_{m_r} + \dots + S_{m_{r+s}} = S_{m_{r+s+1}} = \dots = S_{m_{\alpha+s-1}} = 0 \right] \mathbb{P} \left[S_{m_{r+1}} = \dots + S_{m_{\alpha+s+1}} = \dots = S_{m_{2\alpha-1}} = 0 \right].$$
(52)

By direct calculations and (6),

$$\sum_{n} \lambda^{n} b(n) = \left(\frac{1}{\pi \gamma} \log\left(\frac{1}{1-\lambda}\right)\right)^{\alpha-s+r-2} (1-\lambda)^{-2}$$

$$\cdot \sum_{m_{r},\dots,m_{\alpha+s}=0}^{\infty} \lambda^{m_{r}+\dots+m_{\alpha+s}}$$

$$\cdot \prod_{\substack{t=r+1,\dots,\alpha+s-1\\t\neq r+s}} \mathbb{P}\left(S_{m_{t}}=0\right)$$

$$\cdot \mathbb{P}\left(S_{m_{r}}+\dots+S_{m_{r+s}}=0\right)$$

$$\cdot \mathbb{P}\left(S_{m_{r+s}}+\dots+S_{m_{\alpha+s}}=0\right),$$
(53)

and using Fourier inversion and (6) the internal sum behaves as

$$(2\pi)^{-\alpha-s+r} \int_{-\pi}^{\pi} \cdots \int_{-\pi}^{\pi} (1-\lambda\phi(x))^{-1} (1-\lambda\phi(x)\phi(y))^{-1} (1-\lambda\phi(y))^{-1} (1-\lambda\phi(y))^{-1} (1-\lambda\phi(y))^{-1} (1-\lambda\phi(y))^{-1} dt_j dt_k \right] dx dy \sim (\pi\gamma)^{-\alpha-s+r} (1-\lambda)^{-1}$$
(54)
 $\cdot \log\left(\frac{1}{1-\lambda}\right)^{\alpha-r+s-2} \frac{\pi^2}{6}.$

Then, we have $\sum_n \lambda^n b(n) \sim (\pi^2/6(\pi\gamma)^{2\alpha-2})a(\lambda)$, whence the Tauberian theorem implies that $a(n) - b(n) \sim n^2 \log(n)^{2\alpha-4}/24\pi^{2\alpha-4}\gamma^{2\alpha-2}$. Most importantly we see that the lengths and locations of the chains, *r* and *s*, do not affect the asymptotic behaviour. Noting that if $1 \le r$, $s \le \alpha - 1$, we can partition $2\alpha = r + s + (\alpha - r) + (\alpha - s)$ in $(\alpha - 1)^2$ ways, and thus overall the total contribution from terms with v = 3 is

$$\left[\frac{\left(\alpha!\left(\alpha-1\right)\right)^{2}}{12\pi^{2\alpha-4}\gamma^{2\alpha-2}}\right]n^{2}\log\left(n\right)^{2\alpha-4}.$$
(55)

Case 2 ($\nu(\delta) = 2$). The typical term c(n) was introduced in (33) in the proof of Proposition 9. Now we let $\lambda \in \mathbb{C}$, with $|\lambda| < 1$. By lengthy but direct calculations we can derive an expression of the form

$$\sum_{n} \lambda^{n} c(n) = \frac{\alpha - 1}{\left(\gamma \pi\right)^{2\alpha - 2}} a(\lambda) + o(a(\lambda)), \quad \lambda \longrightarrow 1.$$
 (56)

The approach developed in [13] can then be used to bound the error terms and show that $c(n) \sim [(\alpha - 1)/2(\gamma \pi)^{2\alpha-2}]n^2\log(n)^{2\alpha-4}$.

Finally taking into account the fact that l_1, \ldots, l_α can be in any of the $\alpha - 1$ intervals $[k_i, k_{i+1}]$, for $i = 1, \ldots, \alpha - 1$, the result follows the overall contribution of terms with $v(\delta) = 2$

$$\frac{(\alpha - 1)^2}{2(\gamma \pi)^{2\alpha - 2}} n^2 \log(n)^{2\alpha - 4}.$$
 (57)

The case for d = 2 is very similar, so we move on to the case d = 3.

Case 3 (d = 3 and $\alpha = 2$). Using the same notation as before, we have three terms to consider a(n), b(n), and c(n). We first consider c(n). Letting $K := \epsilon/\sqrt{1-\lambda}$ and using the usual power series construction and spherical coordinates

$$\sum_{n} \lambda^{n} c(n) = (1 - \lambda)^{-2} (2\pi)^{-6}$$

$$\cdot \iint_{J^{3} \times J^{3}} \frac{\lambda f(y) (1 - f(x)) dx dy}{(1 - \lambda f(x))^{2} (1 - \lambda f(y)) (1 - \lambda f(x) f(y))} \sim 2 (2\pi)^{-4} |\Sigma|^{-1} (1 - \lambda)^{-2} \cdot \iint_{0}^{K} \frac{r^{4} s^{2} dr ds}{(1 + r^{2})^{2} (1 + s)^{2} (1 + r^{2} + s^{2})} \sim 2 (2\pi)^{-4} |\Sigma|^{-1} \cdot \frac{\pi}{2} (1 - \lambda)^{-2} \log\left(\frac{1}{1 - \lambda}\right) =: \kappa_{1} (1 - \lambda)^{-2} \log\left(\frac{1}{1 - \lambda}\right),$$
(58)

and thus $c(n) \sim \kappa_1 n \log n$, where $\kappa_1 > 0$, where the answer can be justified following [13].

The term a(n) - b(n) is trickier to compute. As usual we consider the power series

$$\sum_{n\geq 0} \lambda^{n} (a (n) - b (n)) = (1 - \lambda)^{-2} (2\pi)^{-6}$$

$$\cdot \iint_{B(\epsilon)} \frac{dx \, dy}{(1 - \lambda f (x)) (1 - \lambda f (y)) (1 - \lambda f (x + y))}$$

$$- (1 - \lambda)^{-2} (2\pi)^{-6}$$

$$\cdot \iint_{B(\epsilon)} \frac{dx \, dy}{(1 - \lambda f (x)) (1 - \lambda f (y)) (1 - \lambda f (x) f (y))}$$

$$= (1 - \lambda)^{-2} (2\pi)^{-6} (I_{1} (\lambda) - I_{2} (\lambda)).$$
(59)

Let $A \in [-1, 1]$ be the cosine of the angle between x and y, which in spherical coordinates is

$$A = A (\theta_1, \theta_2, \phi_1, \phi_2)$$

= $\cos (\phi_1 - \phi_2) \sin (\theta_1) \sin (\theta_2)$ (60)
+ $\cos (\theta_1) \cos (\theta_2)$.

Then as $0 < \lambda \uparrow 1$, using the expansion (6)

$$I_{1}(\lambda) \sim |\Sigma|^{-1} \int_{r,s=0}^{\epsilon} \int_{\phi_{1,2}=0}^{2\pi} \int_{\theta_{1},\theta_{2}=0}^{\pi} \frac{r^{2}s^{2}\sin(\theta_{1})\sin(\theta_{2}) d\theta_{1} d\theta_{2} d\phi_{1} d\phi_{2} dr ds}{(1-\lambda+\lambda r^{2})(1-\lambda+\lambda s^{2})[1-\lambda+\lambda(r^{2}+s^{2}+2Ars)]}$$

$$= |\Sigma|^{-1} \int_{\theta_{1},\theta_{2}=0}^{\pi} \int_{\phi_{1},\phi_{2}=0}^{2\pi} \int_{r,s=0}^{K} \frac{\sin(\theta_{1})\sin(\theta_{2})r^{2}s^{2} ds dr d\phi_{1} d\phi_{2} d\theta_{1} d\theta_{2}}{(1+r^{2})(1+s^{2})[1+r^{2}+s^{2}+2Ars]}$$

$$\sim |\Sigma|^{-1} \log(K) \int_{\theta_{1},\theta_{2}=0}^{\pi} \int_{\phi_{1},\phi_{2}=0}^{2\pi} \sin(\theta_{1})\sin(\theta_{2}) \frac{\arccos(A(\theta_{1},\theta_{2},\phi_{1},\phi_{2}))}{\sqrt{1-A(\theta_{1},\theta_{2},\phi_{1},\phi_{2})^{2}}} d\phi_{1} d\phi_{2} d\theta_{1} d\theta_{2}.$$
(61)

The other integral is slightly easier

and thus overall we must have that

$$I_{2}(\lambda) \sim |\Sigma|^{-1} \frac{\pi}{2} \log K$$

$$\cdot \int_{\theta_{1},\theta_{2}=0}^{\pi} \int_{\phi_{1},\phi_{2}=0}^{2\pi} \sin(\theta_{1}) \sin(\theta_{2}) d\phi_{1} d\phi_{2} d\theta_{1} d\theta_{2},$$
(62)

$$(I_1 - I_2)(\lambda) \sim \frac{1}{2} (2\pi)^{-6} |\Sigma|^{-1} (1 - \lambda)^{-2} \log\left(\frac{1}{1 - \lambda}\right)$$
$$\cdot \int_{\theta_1, \theta_2 = 0}^{\pi} \int_{\phi_1, \phi_2 = 0}^{2\pi} \left[\frac{\arccos(A)}{\sqrt{1 - A^2}} - \frac{\pi}{2}\right] \sin(\theta_1)$$

whence it follows that $\operatorname{var}(L_n(2)) \sim (\kappa_1 + \kappa_2)n \log n$. To prove the last claim let $S'_n = X'_1 + \cdots + X'_n$ be another random walk, independent of S_n , such that its characteristic function $f'(t) = \mathbb{E}[\exp(itX'_i)]$ also satisfies the expansion (6). Then using [13, Lemma 3.1], one can adapt the proof of [13, Theorem 2.1] to show that $L'_n(\alpha) = L_n(\alpha) + o(L_n(\alpha))$.

Competing Interests

The authors declare that they have no competing interests.

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