

Research Article

Estimate Stress-Strength Reliability Model Using Rayleigh and Half-Normal Distribution

Osama Abdulaziz Alamri,¹ M. M. Abd El-Raouf ^(b),² Eman Ahmed Ismail,³ Zahra Almaspoor ^(b),⁴ Basim S. O. Alsaedi ^(b),¹ Saima Khan Khosa,⁵ and M. Yusuf ^(b)

¹Department of Statistics, Faculty of Science, University of Tabuk, Tabuk 71491, Saudi Arabia

²Basic and Applied Science Institute, Arab Academy for Science, Technology and Maritime Transport (AASTMT), Alexandia, Egypt

³College of International Transport and Logistics,

Arab Academy for Science, Technology and Maritime Transport, Smart Village Campus, Cairo, Egypt

⁴Department of Statistics, Yazd University, P.O. Box 89175-741, Yazd, Iran

⁵Department of Statistics, Bahauddin Zakariya University, Multan, Pakistan

⁶Department of Mathematics, Faculty of science, Helwan University, Cairo, Egypt

Correspondence should be addressed to Zahra Almaspoor; z.almaspoor@stu.yazd.ac.ir

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In the field of life testing, it is very important to study the reliability of any component under testing. One of the most important subjects is the "stress-strength reliability" term which always refers to the quantity P(X > Y) in any statistical literature. It resamples a system with random strength (X) that is subjected to a random strength (Y) such that a system fails in case the stress exceeds the strength. In this study, we consider stress-strength reliability where the strength (X) follows Rayleigh-half-normal distribution and stress (Y_1, Y_2, Y_3 , and Y_4) follows Rayleigh-half-normal distribution, exponential distribution, Rayleigh distribution, and half-normal distribution, respectively. This effort comprises determining the general formulations of the reliabilities of a system. Also, the maximum likelihood estimation approach and method of moment (MOM) will be utilized to estimate the parameters. Finally, reliability has been attained utilizing various values of stress and strength parameters.

1. Introduction

The life of a component is described using the stress-strength models, in reliability theory, that is including a random strength (*X*) which is subjected to a random stress (*Y*). The failure of a component is occurred instantaneously when the stress level applied to it exceeds the level of the strength. Thus, the component reliability is measured by R = P(Y < X). This measurement has a variety of applications, most notably in the engineering industry, such as the degradation of rocket motors and structures, the fatigue failure of aircraft structures, the ageing of concrete pressure vessels, and static fatigue of ceramic components. Therefore, the estimation of R = P(Y < X) has a great importance in the

practical applications. The literature demonstrates that reliability estimation (R) has already been performed when the distributions of (X) and (Y) are Weibull, exponential, or log normal.

Church and Harris [1] firstly introduced the term stressstrength. Many authors have adopted various distributions types for stress and strength. The works of Church and Harris, Surles and Padgett [2], Raqab and Kundu [3], Mokhlis [4], and Saraçoğlu et al. [5] contain the discussion of the estimation problems of the stress-strength reliability model for different distributions. Recently, a review of all methods and results on the stress-strength reliability have presented by Kotz et al. [6]. Bayes estimators and reliability function and the parameters of the Consul, Geeta, and sizebiased Geeta distributions are obtained by Khan Adil and Jan [7]. Akman et al. [8] studied the estimation of reliability using a finite mixture of inverse Gaussian distributions. The estimation of R = P(Y < X) is studied by AI-Hussaini [9] based on a finite mixture of lognormal components. For more reading, see [10–14].

2. Finite Mixture of Rayleigh and Half-Normal Distribution

The Rayleigh-half-normal distribution is denoted as **RHN**(θ) by Abd El-Monsef and Abd El-Raouf [15]. A mixture of Rayleigh and half-normal distribution with a parameter $(1/\sqrt{2\theta})$ is used to represent this model:

$$f(x, \theta) = K f_R \left(\frac{x \cdot 1}{\sqrt{2\theta}}\right) + (1 - K) f_{HN} \left(\frac{x \cdot 1}{\sqrt{2\theta}}\right)$$
$$= K \left(2\theta x e^{-\theta x^2}\right) + (1 - K) \left(2\sqrt{\frac{\theta}{\pi}} e^{-\theta x^2}\right),$$
(1)

where $K = (1/(1 + \sqrt{\pi\theta}))$.

Thus, the Rayleigh-half normal distribution probability density function (pdf) is given by

$$f(x, \theta) = \frac{2\theta(x+1)e^{-\theta x^2}}{1+\sqrt{\pi\theta}}, \quad x, \theta > 0.$$
⁽²⁾

The corresponding cumulative distribution function is given by

$$F(x, \theta) = \frac{1 - e^{-\theta x^2} + \sqrt{\pi \theta} \operatorname{erf}(\sqrt{\theta} x)}{1 + \sqrt{\pi \theta}}, \quad x, \theta > 0, \qquad (3)$$

where erf(u) is the Gauss error function defined as

$$\operatorname{erf}(u) = \frac{2}{\sqrt{\pi}} \int_{0}^{u} e^{-t^{2}} \mathrm{d}t.$$
 (4)

2.1. The Survival Function and the Hazard Function. The reliability function or the survival function S(x) tests the chance of occurring of a breakdown of units beyond certain given point in time. For monitoring, a unit lifetime across the support of its lifetime distribution; generally, the probability that an item will work properly for a specified time period with no failure is the survival function. The definition of the survival function is represented as follows:

$$S(x) = 1 - F(x) = \frac{e^{-\theta x^2} + \sqrt{\pi \theta} \operatorname{erfc}(\sqrt{\theta} x)}{1 + \sqrt{\pi \theta}},$$
 (5)

where $\operatorname{erfc}(u)$ is the complementary error function, and its definition is

$$\operatorname{erfc}(u) = 1 - \operatorname{erf}(u) = \frac{2}{\sqrt{\pi}} \int_{u}^{\infty} e^{-t^{2}} dt.$$
 (6)

The definition of the hazard rate function is the ratio between the density function and its survival function, which measures the tendency to die or to fail depending on the reached age, and therefore, it has a critical role in the classification of the distributions of lifetime, so the hazard rate function of the RHN distribution is given by

$$h(x) = \frac{f(x)}{S(x)} = \frac{2\theta(1+x)}{1+e^{\theta x^2}\sqrt{\pi\theta}\operatorname{erf}(\sqrt{\theta}x)}.$$
(7)

3. Stress-Strength Reliability Computations

In this section, the reliability R = P(Y < X) was derived, where the random variables (X) and (Y) are the independent random variables, where the strength X follows Rayleighhalf normal distribution and the stress (Y) takes different cases (Rayleigh-half normal distribution, exponential distribution, Rayleigh distribution, and half-normal distribution).

Let (X) and (Y) be two independent random variables, where (X) represents "strength" and (Y) represents "stress" and (X), and (Y) follows a joint pdf $f(x, \theta)$; thus, the component reliability is

$$R = P(Y < X) = \int_{-\infty}^{\infty} \int_{-\infty}^{x} f(x, y) dy dx.$$
 (8)

In case that the random variables are statistically independent, then f(x, y) = f(x)g(y) so that

$$R = \int_{-\infty}^{\infty} \int_{-\infty}^{x} f(x)g(x)dydx,$$
 (9)

where f(x) and g(y) are pdf's of X and Y, respectively.

3.1. The Stress and the Strength Follows Rayleigh-Half-Normal Distribution. As the strength $X \sim \text{RHN}(\theta)$ and $Y_1 \sim \text{RHN}(\theta_1)$, they are independent random variables with pdf f(x) and $g(y_1)$, respectively:

$$f(x) = \frac{2\theta(x+1)e^{-\theta x^2}}{1+\sqrt{\pi\theta}}, \quad 0 < \theta \cdot x,$$

$$g(y_1) = \frac{2\theta_1(y_1+1)e^{-\theta_1 y_1^2}}{1+\sqrt{\pi\theta_1}}, \quad 0 < \theta_1 \cdot y_1.$$
(10)

We derive the reliability R = P(Y < X) as follows:

$$R_{1} = P(Y < X) = \int_{0}^{\infty} \int_{0}^{x} f(x)g(y_{1})dydx$$
$$= \int_{0}^{\infty} \int_{0}^{x} \left(\frac{2\theta_{1}(y_{1}+1)e^{-\theta_{1}y_{1}^{2}}}{1+\sqrt{\pi\theta_{1}}}\right) \left(\frac{2\theta(x+1)e^{-\theta_{x}^{2}}}{1+\sqrt{\pi\theta}}\right)dydx.$$
(11)

And, we get after the simplification:

$$R_{1} = \frac{\left(1 + \sqrt{\pi\theta_{1}} + 2\sqrt{\theta\theta_{1}}\operatorname{Tan}^{-1}\left(\sqrt{\theta/\theta_{1}}\right) + \left(\sqrt{\pi}\left(\theta - \theta_{1}\right)/\sqrt{\theta + \theta_{1}}\right) - \left(\theta_{1}/\left(\theta + \theta_{1}\right)\right)\right)}{1 + \sqrt{\pi\theta} + \sqrt{\pi\theta_{1}} + \pi\sqrt{\theta\theta_{1}}}.$$
(12)

3.2. The Strength Follows RHN Distribution and the Stress Follows Exponential Distribution. In this case, the probability density function (pdf) for the stress Y_2 that follows the exponential distribution is given by

$$g(y_2) = \theta_2 e^{-y_1 \theta_2}, \quad y_2, \theta_2 > 0.$$
 (13)

Then, reliability function R_2 for the independent random variables X and Y_2 :

$$R_{2} = \int_{0}^{\infty} \int_{0}^{x} \theta_{2} e^{-y_{2}\theta_{2}} \left(\frac{2\theta(x+1)e^{-\theta x^{2}}}{1+\sqrt{\pi\theta}} \right) dy dx,$$

$$R_{2} = \frac{\sqrt{\pi}}{2(\sqrt{\theta}+\sqrt{\pi}\theta)} \left(2\theta - (2\theta - \theta_{2})e^{(\theta_{2}^{2}/4\theta)} \operatorname{erfc}\left(\frac{\theta_{2}}{2\sqrt{\theta}}\right) \right),$$
(14)

where the strength follows RHN distribution.

3.3. The Strength Follows RHN Distribution and the Stress Follows Rayleigh Distribution. In this case, the probability density function (pdf) for the stress Y_3 that follows the Rayleigh distribution is given by

$$g(y_3) = \frac{y_3}{\theta_2^2} e^{-(y_3^2/2\theta_3^2)}, \quad y, \theta_3 > 0.$$
(15)

Then, reliability function R_3 for the independent random variables X and Y_3 is

$$\begin{split} R_{3} &= \int_{0}^{\infty} \int_{0}^{x} \left(\frac{y_{3}}{\theta_{3}^{2}} e^{-\left(y_{3}^{2}/2\theta_{3}^{2}\right)} \right) \left(\frac{2\theta\left(x+1\right)e^{-\theta x^{2}}}{1+\sqrt{\pi\theta}} \right) \mathrm{d}y \mathrm{d}x \\ &= \frac{2\theta}{\left(1+\sqrt{\pi\theta}\right)} \int_{0}^{\infty} (x+1)e^{-\theta x^{2}} \left(1-e^{-\left(x^{2}/2\theta_{3}^{2}\right)}\right) \mathrm{d}x, \\ R_{3} &= \frac{\theta}{1+\sqrt{\pi\theta}} \left(\frac{1}{\theta} + \frac{\sqrt{\pi}}{\sqrt{\theta}} - \frac{2}{2\theta + \left(1/\theta_{3}^{2}\right)} - \frac{\sqrt{2\pi}}{\sqrt{2\theta + \left(1/\theta_{3}^{2}\right)}} \right), \end{split}$$
(16)

where the strength follows RHN distribution.

3.4. The Strength Follows RHN Distribution and the Stress Follows Half-Normal Distribution. In this case, the probability density function (pdf) for the stress Y_4 that follows half-normal distribution is given by

$$g(y_4) = \left(\frac{\sqrt{2}}{\theta_4 \sqrt{\pi}}\right) e^{-(y_4^2/2\theta_4^2)}, \quad y, \theta_4 > 0.$$
(17)

Then, reliability function R_4 for the independent random variables X and Y_4

$$\begin{split} R_4 &= \int_0^\infty \int_0^x \left(\left(\frac{\sqrt{2}}{\theta_4 \sqrt{\pi}} \right) e^{-\frac{y_4^2}{2\theta_4^2}} \right) \left(\frac{2\theta \left(x + 1 \right) e^{-\theta x^2}}{1 + \sqrt{\pi \theta}} \right) \mathrm{d}y \mathrm{d}x \\ &= \frac{2\theta}{\left(1 + \sqrt{\pi \theta} \right)} \int_0^\infty \mathrm{Erf} \left(\frac{x}{\sqrt{2} \theta_4} \right) (x+1) e^{-\theta x^2} \mathrm{d}x, \\ R_4 &= \frac{1}{\sqrt{\theta} \left(1 + \sqrt{\pi \theta} \right)} \left(\frac{2\theta \mathrm{Cot}^{-1} \left(\theta_4 \sqrt{2\theta} \right)}{\sqrt{\pi}} + \frac{1}{\theta_4 \sqrt{2} + \left(1/\theta \theta_4^2 \right)} \right), \end{split}$$
(18)

where the strength follows RHN distribution.

4. Estimation of Stress-Strength Reliability

In the literature, a discussion of the estimation R = P(Y < X) when random variables (*X*) and (*Y*) are following the specified distributions have been presented including engineering statistics, quality control, medicine, reliability, biostatistics, and psychology. This quantity for a limited number of cases could be calculated in a closed form (Nadarajah [16] and Barreto-Souza et al. [17]). Several authors including Milan and Vesna [18] have considered the estimation of (*R*) for independent variables and normally distributed (*X*) and (*Y*). Later, a list of papers related to the estimation problem of (*R*) were reported by Greco and Venture [19] when (*X*) and (*Y*) are independent and follow a class of lifetime distributions containing Gamma distributions, exponential, generalized exponential, bivariate exponential, Weibull distribution, Burr type t model, and others.

4.1. Method of Moment (MOM) Estimation of R. The estimation of reliability is very common in the statistical literature. Now, to compute \hat{R} , we need to estimate the parameters θ and θ_i , i = 1, 2, 3, 4, in four cases of stress.

Since the strengths X follow RHN (θ), the stress have four cases:

(i) Y_1 follows Rayleigh-half normal distribution with parameter θ_1

- (ii) Y_2 follows exponential distribution with parameter θ_2
- (iii) Y_3 follows Rayleigh distribution with parameter θ_3
- (iv) Y_4 follows half-normal distribution with parameter θ_4 ; then, their population means are given by

$$\overline{x} = \frac{2\sqrt{\theta} + \sqrt{\pi}}{2\sqrt{\theta}(1 + \sqrt{\theta\pi})},$$

$$\overline{y}_1 = \frac{2\sqrt{\theta_1} + \sqrt{\pi}}{2\sqrt{\theta_1}(1 + \sqrt{\theta_1\pi})},$$

$$\overline{y}_2 = \frac{1}{\theta_2},$$

$$\overline{y}_3 = \theta_3\sqrt{\frac{\pi}{2}},$$

$$\overline{y}_4 = \theta_4\sqrt{\frac{2}{\pi}}.$$
(19)

The ME's of θ , θ_1 , θ_2 , θ_3 , and θ_4 , denoted by $\hat{\theta}$, $\hat{\theta}_1$, $\hat{\theta}_2$, $\hat{\theta}_3$, and $\hat{\theta}_4$, respectively, can be obtained by solving $(\overline{x}, \overline{y}_1, \overline{y}_2, \overline{y}_3, \overline{y}_3, \overline{y}_4)$ numerically:

$$\begin{aligned} \widehat{\theta} &= \frac{\left(\sum_{i=1}^{m} x_{i} - n\right)^{2} + n\pi \sum_{i=1}^{m} x_{i}}{2\pi \left(\sum_{i=1}^{m} x_{i}\right)^{2}} + \frac{\sqrt{\left(\sum_{i=1}^{m} x_{i} - n\right)^{4} + 2n\pi \sum_{i=1}^{m} x_{i} \left(\sum_{i=1}^{m} x_{i} - n\right)^{2}}}{2\pi \left(\sum_{i=1}^{m} x_{i}\right)^{2}}, \\ \widehat{\theta}_{1} &= \frac{\left(\sum_{j=1}^{m} y_{1_{j}} - m\right)^{2} + m\pi \sum_{j=1}^{m} y_{1_{j}}}{2\pi \left(\sum_{j=1}^{m} y_{1_{j}}\right)^{2}} + \frac{\sqrt{\left(\sum_{j=1}^{m} y_{1_{j}} - n\right)^{4} + 2m\pi \sum_{j=1}^{m} y_{1_{j}} \left(\sum_{j=1}^{m} y_{1_{j}} - m\right)^{2}}}{2\pi \left(\sum_{j=1}^{m} y_{1_{j}}\right)^{2}}, \\ \widehat{\theta}_{2} &= \frac{m}{\sum_{j=1}^{m} y_{2_{j}}}, \\ \widehat{\theta}_{3} &= \sqrt{\frac{2}{\pi}} \left(\frac{\sum_{j=1}^{m} y_{3_{j}}}{m}\right), \\ \widehat{\theta}_{4} &= \sqrt{\frac{\pi}{2}} \left(\frac{\sum_{j=1}^{m} y_{4_{j}}}{m}\right). \end{aligned}$$

$$(20)$$

The ME of *R*, denoted by $\hat{R}_1 \cdot \hat{R}_2 \cdot \hat{R}_3$ and \hat{R}_4 is obtained by substitute $\hat{\theta}$ with $\hat{\theta}_1 \cdot \hat{\theta}_2 \cdot \hat{\theta}_3$ and $\hat{\theta}_4$ in $R_1 \cdot R_2 \cdot R_3$ and R_4 .

4.2. The Maximum Likelihood Estimators of R. The maximum likelihood estimator (MLE) is the most popular method for reliability estimation R = p(Y < X) because of its generality and flexibility. This method can be used

if the joint distribution of the strength (X) and the stress (Y) is a known function with some unknown parameters.

Suppose $x_1 \cdot x_2 \cdots x_n$ is a random sample from RHN distribution with θ and $y_{11} \cdot y_{12} \cdots y_{1m}$ is a random sample from RHN distribution with θ_1 . Then, the likelihood function is given by

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$$L(\theta \cdot \theta_1; x \cdot y_1) = 2^{n+m} \theta^n \theta_1^m - (1 + \sqrt{\pi\theta})^n - \left(1 + \sqrt{\pi\theta_1}\right)^m \prod_{i=1}^n (x_i + 1) e^{-\theta x_i^2} \prod_{j=1}^m (y_{1j} + 1) e^{-\theta_1 y_{1j}^2}.$$
 (21)

And, the log-likelihood function of the observed samples is

$$\ln L(\theta \cdot \theta_{1}) = (m+n)\ln(2) + n \ln(\theta) + m \ln(\theta_{1}) - n \ln(1 + \sqrt{\pi\theta}) - m \ln\left(1 + \sqrt{\pi\theta_{1}}\right) - \theta \sum_{i=1}^{n} x_{i}^{2} - \theta_{1} \sum_{j=1}^{m} y_{1j}^{2} + \sum_{i=1}^{n} \ln(x_{i}+1) + \sum_{j=1}^{m} \ln(y_{1j}+1).$$
(22)

By solving the following equations, the MLE of θ and θ_2 can be obtained:

$$\frac{\left(\partial \ln L(\theta \cdot \theta_{1})\right)}{\partial \theta} = \frac{n}{\theta} - \frac{\left(n\sqrt{\pi}\right)}{\left(2\sqrt{\theta}\left(1 + \sqrt{\pi\theta}\right)\right)} - \sum_{(i=1)}^{n} x_{i}^{2} = 0,$$
$$\frac{\partial \ln L(\theta \cdot \theta_{1})}{\partial \theta_{1}} = \frac{m}{\theta_{1}} - \frac{m\sqrt{\pi}}{2\sqrt{\theta_{1}}\left(1 + \sqrt{\pi\theta_{1}}\right)} - \sum_{j=1}^{m} y_{1j}^{2} = 0.$$
(23)

The MLEs of θ and θ_1 can be obtained, respectively, as

$$\hat{\theta} = \frac{1}{6\pi A^2} \left(B + 2A(A + n\pi) + \frac{A^2 \left(n^2 \pi^2 + 4A(A - 4n\pi)\right)}{B} \right),$$
$$\hat{\theta}_1 = \frac{1}{6\pi C^2} \left(D + 2C(C + m\pi) + \frac{C^2 \left(m^2 \pi^2 + 4C(C - 4m\pi)\right)}{D} \right),$$
(24)

where $A = \sum_{i=0}^{n} x_i^2 \cdot C = \sum_{j=0}^{m} y_{1j}^2$,

$$B = \left(8A^{6} - 48\pi nA^{5} + 51\pi^{2}n^{2}A^{4} - \pi^{3}n^{3}A^{3} + 3\sqrt{3}\pi^{3/2}\sqrt{n^{3}A^{7}\left(-16A^{2} + 71\pi nA - 2\pi^{2}n^{2}\right)}\right)^{1/3},$$
(25)

$$\begin{split} D &= \left(8C^6 - 48\pi mC^5 + 51\pi^2 m^2 C^4 - \pi^3 m^3 C^3 + 3\sqrt{3} \\ \pi^{3/2} \sqrt{m^3 C^7 \left(-16C^2 + 71\pi mC - 2\pi^2 m^2\right)}\right)^{1/3}. \end{split}$$

Then, the maximum likelihood estimator of *R* when the strength *X* follows RHN(θ) distribution and stress *Y* follows RHN(θ_1) distribution is given as

$$\widehat{R}_{1} = \frac{\left(1 + \sqrt{\pi\widehat{\theta}_{1}} + 2\sqrt{\widehat{\theta}\widehat{\theta}_{1}}\operatorname{Tan}^{-1}\left(\sqrt{\left(\widehat{\theta}/\widehat{\theta}_{1}\right)}\right) + \left(\sqrt{\pi}\left(\widehat{\theta} - \widehat{\theta}_{1}\right)/\sqrt{\widehat{\theta} + \widehat{\theta}_{1}}\right) - \left(\widehat{\theta}_{1}/\left(\widehat{\theta} + \widehat{\theta}_{1}\right)\right)\right)}{1 + \sqrt{\pi\widehat{\theta}} + \sqrt{\pi\widehat{\theta}_{1}} + \pi\sqrt{\widehat{\theta}\widehat{\theta}_{1}}}.$$
(26)

Similarly, we perform the same steps to find (MLE) in other cases; we can obtain

(i) When the stress Y_2 that follows the exponential distribution with parameter θ_2 , the MLE of R_2 is given as

$$\widehat{R}_{2} = \frac{\sqrt{\pi}}{2\left(\sqrt{\widehat{\theta}} + \sqrt{\pi}\,\widehat{\theta}_{2}\right)} \left(2\widehat{\theta} - \left(2\widehat{\theta} - \widehat{\theta}_{2}\right)e^{\left(\widehat{\theta}_{2}^{2}/4\widehat{\theta}\right)}\operatorname{Erfc}\left(\frac{\widehat{\theta}_{2}}{2\sqrt{\widehat{\theta}}}\right)\right),\tag{27}$$

where the strength X follows Rayleigh-half-normal distribution with parameter θ .

(ii) When the stress Y_3 that follows Rayleigh distribution with parameter θ_3 and the strength X follows Rayleigh-half-normal distribution with parameter θ , the MLE of R_3 is given as

$$\widehat{R}_{3} = \frac{\widehat{\theta}}{1 + \sqrt{\pi\widehat{\theta}}} \left(\frac{1}{\widehat{\theta}} + \frac{\sqrt{\pi}}{\sqrt{\widehat{\theta}}} - \frac{2}{2\widehat{\theta} + \left(1/\widehat{\theta}_{3}^{2}\right)} - \frac{\sqrt{2\pi}}{\sqrt{2\widehat{\theta} + \left(1/\widehat{\theta}_{3}^{2}\right)}} \right).$$
(28)

(iii) When the stress Y_4 that follows half-normal distribution with parameter θ_4 , the M. L. E of R_2 is given as

				-	e						
θ	$ heta_1$										
	0.1	0.2	0.3	0.4	0.5	0.6	0.7	0.8	0.9	1	
0.1	0.208	0.359	0.472	0.558	0.624	0.677	0.718	0.752	0.780	0.803	
0.2	0.149	0.267	0.364	0.442	0.507	0.561	0.606	0.645	0.678	0.706	
0.3	0.121	0.222	0.307	0.379	0.440	0.493	0.538	0.577	0.611	0.641	
0.4	0.105	0.194	0.271	0.337	0.395	0.445	0.490	0.529	0.563	0.594	
0.5	0.093	0.174	0.245	0.307	0.362	0.410	0.453	0.491	0.526	0.556	
0.6	0.085	0.159	0.225	0.284	0.336	0.382	0.424	0.461	0.495	0.526	
0.7	0.078	0.148	0.210	0.265	0.315	0.360	0.400	0.437	0.470	0.500	
0.8	0.073	0.138	0.197	0.250	0.297	0.341	0.380	0.416	0.448	0.478	
0.9	0.068	0.130	0.186	0.237	0.283	0.325	0.363	0.398	0.430	0.459	
1	0.065	0.124	0.177	0.226	0.270	0.311	0.348	0.382	0.413	0.443	

TABLE 1: Variation in R_1 when strength and stress has RHN distribution.

$$\widehat{R}_{4} = \frac{1}{\sqrt{\widehat{\theta}}\left(1 + \sqrt{\pi\widehat{\theta}}\right)} \left(\frac{2\widehat{\theta}\operatorname{Cot}^{-1}\left(\widehat{\theta}_{2}\sqrt{2\widehat{\theta}}\right)}{\sqrt{\pi}} + \frac{1}{\widehat{\theta}_{4}\sqrt{2 + \left(1/\widehat{\theta}\widehat{\theta}_{4}^{2}\right)}}\right),$$
(29)

where the strength X follows Rayleigh-half-normal distribution with parameter θ .

5. Numerical Evaluation

In different cases, the system reliability R has evaluated for some specific values of the parameters involved in the expression of R.

5.1. Case 1: Strength and Stress Follows RHN Distribution. From Table 1 and Figures 1 and 2, it is noticed that, with the increase in the strength parameter values, the reliability value decreases. If the stress parameter increases, then the value of reliability increases.

5.2. Case 2: Strength Follows RHN Distribution and Stress Follows Exponential Distribution. From Table 2 and Figures 3 and 4, it is observed that if the strength parameter increases then the value of reliability increases. If the stress parameter increases, then the value of reliability increases.

5.3. Case 3: Strength Follows RHN Distribution and Stress Follows Rayleigh Distribution. From Table 3 and Figures 5 and 6, it is noticed that, with increasing the value of the strength and stress parameter, the reliability value decreases.

5.4. Case 4: Strength Follows RHN Distribution and Stress Follows Half-Normal Distribution. From Table 4 and Figures 7 and 8, it is noticed that, with increasing the value of the strength and stress parameter, the reliability value decreases.



FIGURE 1: Variation in R_1 for constant stress.



FIGURE 2: Variation in R_1 for constant strength.

5.5. Simulation Study. In this section, some results are represented depending on Monte-Carlo simulation, for comparing the estimates of (*R*) performance using MLE and MOM estimators fundamentally for many sample sizes. The following sample sizes are considered; (n, m) = (5, 5), (10, 10), (20, 20), (30, 30), (40, 40), (50, 50), and (100, 100). From each sample, the estimates are computed for the parameters using MLE and method of moment estimation. Once the

			2	U			*					
θ	θ_2											
	0.1	0.2	0.3	0.4	0.5	0.6	0.7	0.8	0.9	1		
0.1	0.500	0.349	0.274	0.228	0.197	0.175	0.157	0.144	0.133	0.124		
0.2	0.651	0.500	0.412	0.354	0.312	0.280	0.256	0.236	0.219	0.205		
0.3	0.726	0.588	0.500	0.438	0.392	0.357	0.328	0.304	0.285	0.268		
0.4	0.772	0.646	0.562	0.500	0.453	0.415	0.384	0.359	0.337	0.318		
0.5	0.803	0.688	0.608	0.547	0.500	0.462	0.430	0.403	0.380	0.361		
0.6	0.825	0.720	0.643	0.585	0.538	0.500	0.468	0.441	0.417	0.396		
0.7	0.843	0.744	0.672	0.616	0.570	0.532	0.500	0.472	0.448	0.427		
0.8	0.856	0.764	0.696	0.641	0.597	0.559	0.528	0.500	0.476	0.455		
0.9	0.867	0.781	0.715	0.663	0.620	0.583	0.552	0.524	0.500	0.479		
1	0.876	0.795	0.732	0.682	0.639	0.604	0.573	0.545	0.521	0.500		

TABLE 2: Variation in R_2 when strength has RHN distribution and stress has Exponential distribution.



FIGURE 3: Variation in R_2 for constant stress.



FIGURE 4: Variation in R_2 for constant strength.

TABLE 3: Variation in R_3 when strength has RHN distribution and stress has Rayleigh distribution.

θ	θ_3											
	0.1	0.2	0.3	0.4	0.5	0.6	0.7	0.8	0.9	1		
0.1	0.983	0.963	0.941	0.917	0.891	0.864	0.835	0.806	0.777	0.747		
0.2	0.970	0.936	0.898	0.858	0.816	0.773	0.730	0.687	0.645	0.604		
0.3	0.959	0.913	0.863	0.810	0.756	0.702	0.650	0.600	0.552	0.508		
0.4	0.949	0.892	0.831	0.768	0.706	0.645	0.587	0.533	0.483	0.438		
0.5	0.940	0.874	0.804	0.732	0.662	0.596	0.535	0.479	0.429	0.385		
0.6	0.932	0.857	0.778	0.700	0.625	0.555	0.492	0.436	0.386	0.343		
0.7	0.924	0.841	0.755	0.671	0.591	0.519	0.455	0.399	0.351	0.309		
0.8	0.917	0.826	0.734	0.644	0.562	0.488	0.424	0.368	0.321	0.281		
0.9	0.910	0.812	0.714	0.620	0.535	0.460	0.396	0.342	0.296	0.257		
1	0.903	0.799	0.695	0.598	0.511	0.435	0.372	0.318	0.274	0.238		



FIGURE 5: Variation in R_3 for constant stress.



FIGURE 6: Variation in R_3 for constant strength.

TABLE 4: Variation in R_4 when strength has RHN distribution and stress has half-normal distribution.

θ	$ heta_4$											
	0.1	0.2	0.3	0.4	0.5	0.6	0.7	0.8	0.9	1		
0.1	0.989	0.977	0.964	0.950	0.934	0.918	0.901	0.884	0.866	0.848		
0.2	0.981	0.960	0.937	0.913	0.888	0.862	0.835	0.808	0.781	0.755		
0.3	0.974	0.946	0.915	0.883	0.850	0.816	0.783	0.750	0.718	0.687		
0.4	0.968	0.933	0.896	0.857	0.817	0.778	0.740	0.703	0.668	0.634		
0.5	0.963	0.921	0.878	0.834	0.789	0.745	0.704	0.664	0.627	0.592		
0.6	0.957	0.911	0.862	0.812	0.764	0.717	0.672	0.630	0.592	0.557		
0.7	0.952	0.901	0.847	0.793	0.741	0.691	0.644	0.602	0.562	0.527		
0.8	0.948	0.891	0.833	0.775	0.720	0.668	0.620	0.576	0.537	0.501		
0.9	0.943	0.883	0.820	0.759	0.701	0.647	0.598	0.554	0.514	0.479		
1	0.939	0.874	0.808	0.744	0.683	0.628	0.578	0.533	0.494	0.459		



FIGURE 7: Variation in R_4 for constant stress.



FIGURE 8: Variation in R_4 for constant strength.

		1110112 01	interage mod of				
$(\theta \cdot \theta_1)$				(<i>n</i> , <i>m</i>)			
$(\theta \cdot \theta_1)$	(5, 5)	(10, 10)	(20, 20)	(30, 30)	(40, 40)	(50, 50)	(100, 100)
(1 1)	0.0196	0.0180	-0.0181	-0.0073	-0.0041	-0.0034	0.0005
(1, 1)	-0.0720	-0.0318	-0.0193	0.0138	-0.0084	0.0076	-0.0029
(1, 0.5)	-0.0594	-0.0476	-0.0132	-0.0120	-0.0101	-0.0043	0.0048
(1, 0.5)	-0.0790	-0.0747	-0.0643	-0.0662	-0.0617	-0.0658	-0.0555
$(1 \ 1 \ 5)$	-0.0419	0.0262	0.0172	0.0146	-0.0063	0.0024	0.0011
(1, 1.5)	0.0678	0.0643	0.0557	0.0478	0.0423	0.0372	0.0370
(1, 2)	-0.0125	0.0115	0.0111	0.0113	-0.0073	0.0044	-0.0015
(1, 2)	0.0800	-0.0777	0.0687	0.0663	0.0656	0.0649	0.0638
(0, 5, 1)	0.0252	0.0146	0.0122	-0.0097	-0.0030	-0.0018	0.0011
(0.5, 1)	0.0743	0.0737	0.0686	0.0656	0.0654	0.0639	0.0627
(1 5 1)	0.0594	-0.0090	0.0076	-0.0064	0.0062	0.0019	0.0019
(1.5, 1)	-0.0807	-0.0413	-0.0402	-0.0364	-0.0352	-0.0164	0.0125
(2, 1)	0.0153	0.0066	0.0066	0.0061	0.0048	0.0048	0.0045
(2, 1)	-0.0997	-0.0726	-0.0719	-0.0647	-0.0624	-0.0618	-0.0584

TABLE 5: Average MSE of the simulated estimates of R_{12}

TABLE 6: Average MSE of the simulated estimates of R_1 .

$(\boldsymbol{\rho} \boldsymbol{\rho})$	(n, m)										
$(\theta \cdot \theta_1)$	(5, 5)	(10, 10)	(20, 20)	(30, 30)	(40, 40)	(50, 50)	(100, 100)				
(1, 1)	0.0051	0.0020	0.0020	0.0013	0.0010	0.0005	0.0003				
())	0.0159	0.0104	0.0090	0.0035	0.0024	0.0020	0.0013				
(1 0 5)	0.0123	0.0095	0.0037	0.0033	0.0025	0.0018	0.0009				
(1, 0.3)	0.0148	0.0112	0.0051	0.0049	0.0046	0.0044	0.0044				
(1 1 5)	0.0077	0.0072	0.0044	0.0030	0.0022	0.0020	-0.0001				
(1, 1.5)	0.0135	0.0091	0.0054	0.0051	0.0034	0.0021	0.0020				
(1, 2)	0.0055	0.0050	0.0048	0.0033	0.0022	0.0019	-0.0003				
(1, 2)	0.0115	0.0086	0.0059	0.0059	0.0049	0.0049	0.0009				
(0, 5, 1)	0.0090	0.0038	0.0038	0.0036	0.0031	0.0024	0.0010				
(0.5, 1)	0.0092	0.0087	0.0070	0.0050	0.0050	0.0048	0.0046				
(1 5 1)	0.0147	0.0097	0.0024	0.0024	-0.0018	0.0014	-0.0008				
(1.5, 1)	0.0221	-0.0124	-0.0090	0.0062	0.0027	0.0022	0.0014				
(2, 1)	0.0121	0.0064	0.0059	0.0030	0.0024	0.0018	-0.0003				
(2, 1)	0.0478	-0.0272	0.0256	-0.0145	-0.0076	0.0050	0.0010				

parameters are estimated, the estimates of R_1 is obtained. The average biases of R_1 is reported in Table 5 and mean squared errors (MSEs) of R_1 are in Table 6.

The first row includes the average bias of R_1 using the MLE and second row includes the average bias of R_1 using the MOM, in each cell.

The first row includes the average MSE of R_1 using the MLE and second row includes the average MSE of R_1 using the MOM, in each cell.

6. Conclusions

The proposed model in this paper, the stress-strength reliability has been studied for Rayleigh-half normal when the strength (X) follows Rayleigh-half normal distribution, and the stress (Y) takes Rayleigh-half normal distribution, exponential distribution, Rayleigh distribution, and halfnormal distribution. Based on the computations and graphs, (i) it has been noticed that when the stress parameter is increased, the reliability value lowers, and when the strength parameter is increased, the reliability value increases. The numerical assessment demonstrates that increasing the stress parameter decreases the dependability value in case (ii), whereas increasing the strength parameter increases the reliability value. In cases (iii) and (vi), increasing the stress parameter decreases the reliability value, whereas increasing the strength parameter increases it. A comparison is carried out between two methods of reliability estimation R = P(X > Y) when (Y) and (X) both follow Rayleigh-half normal distributions for various parameters scale. We provide MLE and MOM procedure for estimating the unknown parameters that are used for reliability estimation (R). Based on the simulation findings, we can conclude that MLE outperforms MOM in terms of average bias and average MSE for a variety of parameter choices.

Data Availability

All data used to support the findings of the study are available within the article.

Conflicts of Interest

The authors declare that they have no conflicts of interest.

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