

Research Article

Adaptive Wavelet Estimation of a Biased Density for Strongly Mixing Sequences

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The estimation of a biased density for exponentially strongly mixing sequences is investigated. We construct a new adaptive wavelet estimator based on a hard thresholding rule. We determine a sharp upper bound of the associated mean integrated square error for a wide class of functions.

1. Introduction

In the standard density estimation problem, we observe n random variables X_1, \dots, X_n with common density function f . The goal is to estimate f from X_1, \dots, X_n . However, in some applications, X_1, \dots, X_n are not accessible; we only have n random variables Z_1, \dots, Z_n with the common density

$$g(x) = \mu^{-1}w(x)f(x), \quad (1.1)$$

where w denotes a known positive function and μ is the unknown normalization parameter: $\mu = \int w(y)f(y)dy$. Our goal is to estimate the "biased density" f from Z_1, \dots, Z_n . Practical examples can be found in, for example, [1–3] and the survey by the author of [4].

The standard i.i.d. case has been investigated in several papers. See, for example, [5–9]. To the best of our knowledge, the dependent case has only been examined in [10] for associated (positively or negatively) Z_1, \dots, Z_n . In this paper, we study another dependent (and realistic) structure which has not been addressed earlier: we suppose that Z_1, \dots, Z_n is a sample of a strictly stationary and exponentially strongly mixing process $(Z_i)_{i \in \mathbb{Z}}$ (to be defined in Section 2). Such a dependence condition arises for a wide class of GARCH-type time series models classically encountered in finance. See, for example, [11, 12] for an overview.

We focus our attention on the wavelet methods because they provide a coherent set of procedures that are spatially adaptive and near optimal over a wide range of function spaces. See, for example, [13, 14] for a detailed coverage of wavelet theory in statistics. We develop two new wavelet estimators: a linear nonadaptive based on projections and a nonlinear adaptive using the hard thresholding rule introduced by [15]. We measure their performances by determining upper bounds of the mean integrated squared error (MISE) over Besov balls (to be defined in Section 3). We prove that our adaptive estimator attains a sharp rate of convergence, close to the one attained by the linear wavelet estimator (constructed in a nonadaptive fashion to minimize the MISE).

The rest of the paper is organized as follows. Section 2 is devoted to the assumptions on the model. In Section 3, we present wavelets and Besov balls. The considered wavelet estimators are defined in Section 4. Section 5 is devoted to the results. The proofs are postponed in Section 6.

2. Assumptions on the Model

We assume that Z_1, \dots, Z_n coming from a strictly stationary process $(Z_i)_{i \in \mathbb{Z}}$. For any $m \in \mathbb{Z}$, we define the m th strongly mixing coefficient of $(Z_i)_{i \in \mathbb{Z}}$ by

$$a_m = \sup_{(A,B) \in \mathcal{F}_{-\infty,0}^Z \times \mathcal{F}_{m,\infty}^Z} |\mathbb{P}(A \cap B) - \mathbb{P}(A)\mathbb{P}(B)|, \quad (2.1)$$

where, for any $u \in \mathbb{Z}$, $\mathcal{F}_{-\infty,u}^Z$ is the σ -algebra generated by the random variables \dots, Z_{u-1}, Z_u and $\mathcal{F}_{u,\infty}^Z$ is the σ -algebra generated by the random variables Z_u, Z_{u+1}, \dots

We consider the exponentially strongly mixing case, that is, there exist three known constants, $\gamma > 0$, $c > 0$, and $\theta > 0$, such that, for any $m \in \mathbb{Z}$,

$$a_m \leq \gamma \exp(-c|m|^\theta). \quad (2.2)$$

This assumption is satisfied by a large class of GARCH processes. See, for example, [11, 12, 16, 17].

Note that, when $\theta \rightarrow \infty$, we are in the standard i.i.d. case.

W.o.l.g., the support of the functions f , and w are $[0, 1]$.

There exist two constants, $c > 0$ and $C > 0$, such that

$$c \leq \inf_{x \in [0,1]} w(x), \quad \sup_{x \in [0,1]} w(x) \leq C. \quad (2.3)$$

There exists a (known) constant $C > 0$ such that

$$\sup_{x \in [0,1]} f(x) \leq C. \quad (2.4)$$

For any $m \in \{1, \dots, n\}$, let $g_{(Z_0, Z_m)}$ be the density of (Z_0, Z_m) . There exists a constant $C > 0$ such that

$$\sup_{m \in \{1, \dots, n\}} \sup_{(x, y) \in [0, 1]^2} |g_{(Z_0, Z_m)}(x, y) - g(x)g(y)| \leq C. \tag{2.5}$$

The two first boundedness assumptions are standard in the estimation of biased densities. See, for example, [6–8].

3. Wavelets and Besov Balls

Let N be an integer ϕ and ψ be the initial wavelets of dbN (so $\text{supp}(\phi) = \text{supp}(\psi) = [1 - N, N]$). Set

$$\phi_{j,k}(x) = 2^{j/2} \phi(2^j x - k), \quad \psi_{j,k}(x) = 2^{j/2} \psi(2^j x - k). \tag{3.1}$$

With an appropriate treatments at the boundaries, there exists an integer τ satisfying $2^\tau \geq 2N$ such that the collection $\mathcal{B} = \{\phi_{\tau,k}(\cdot), k \in \{0, \dots, 2^\tau - 1\}; \psi_{j,k}(\cdot); j \in \mathbb{N} - \{0, \dots, \tau - 1\}, k \in \{0, \dots, 2^j - 1\}\}$, is an orthonormal basis of $\mathbb{L}^2([0, 1])$ (the space of square-integrable functions on $[0, 1]$). See [18].

For any integer $\ell \geq \tau$, any $h \in \mathbb{L}^2([0, 1])$ can be expanded on \mathcal{B} as

$$h(x) = \sum_{k=0}^{2^\ell - 1} \alpha_{\ell,k} \phi_{\ell,k}(x) + \sum_{j=\ell}^{\infty} \sum_{k=0}^{2^j - 1} \beta_{j,k} \psi_{j,k}(x), \quad x \in [0, 1], \tag{3.2}$$

where $\alpha_{j,k}$ and $\beta_{j,k}$ are the wavelet coefficients of h defined by

$$\alpha_{j,k} = \int_0^1 h(x) \phi_{j,k}(x) dx, \quad \beta_{j,k} = \int_0^1 h(x) \psi_{j,k}(x) dx. \tag{3.3}$$

Let $M > 0, s > 0, p \geq 1$, and $r \geq 1$. A function h belongs to $B_{p,r}^s(M)$ if and only if there exists a constant $M^* > 0$ (depending on M) such that the associated wavelet coefficients (3.3) satisfy

$$2^{\tau(1/2-1/p)} \left(\sum_{k=0}^{2^\tau - 1} |\alpha_{\tau,k}|^p \right)^{1/p} + \left(\sum_{j=\tau}^{\infty} \left(2^{j(s+1/2-1/p)} \left(\sum_{k=0}^{2^j - 1} |\beta_{j,k}|^p \right)^{1/p} \right)^r \right)^{1/r} \leq M^*. \tag{3.4}$$

In this expression, s is a smoothness parameter and p and r are norm parameters. For a particular choice of s, p , and r , $B_{p,r}^s(M)$ contains some classical sets of functions as the Hölder and Sobolev balls. See [19].

4. Estimators

Firstly, we consider the following estimator for μ :

$$\hat{\mu} = \left(\frac{1}{n} \sum_{i=1}^n \frac{1}{w(Z_i)} \right)^{-1}. \quad (4.1)$$

It is obtained by the method of moments (see Proposition 6.2 below).

Then, for any integer $j \geq \tau$ and any $k \in \{0, \dots, 2^j - 1\}$, we estimate the unknown wavelet coefficient

$$(i) \alpha_{j,k} = \int_0^1 f(x) \phi_{j,k}(x) dx \text{ by}$$

$$\hat{\alpha}_{j,k} = \frac{\hat{\mu}}{n} \sum_{i=1}^n \frac{\phi_{j,k}(Z_i)}{w(Z_i)}, \quad (4.2)$$

$$(ii) \beta_{j,k} = \int_0^1 f(x) \psi_{j,k}(x) dx \text{ by}$$

$$\hat{\beta}_{j,k} = \frac{\hat{\mu}}{n} \sum_{i=1}^n \frac{\psi_{j,k}(Z_i)}{w(Z_i)}. \quad (4.3)$$

Note that they are those considered in the i.i.d. case (see, e.g., [8, 9]). Their statistical properties, with our dependent structure, are investigated in Propositions 6.2, 6.3, and 6.4 below.

Assuming that $f \in B_{p,r}^s(M)$ with $p \geq 2$, we define the linear estimator \hat{f}^L by

$$\hat{f}^L(x) = \sum_{k=0}^{2^{j_0}-1} \hat{\alpha}_{j_0,k} \phi_{j_0,k}(x), \quad x \in [0, 1], \quad (4.4)$$

where $\hat{\alpha}_{j,k}$ is defined by (4.2) and j_0 is the integer satisfying

$$\frac{1}{2} n^{1/(2s+1)} < 2^{j_0} \leq n^{1/(2s+1)}. \quad (4.5)$$

For a survey on wavelet linear estimators for various density models, we refer the reader to [20]. For the consideration of strongly mixing sequences, see, for example, [21, 22].

We define the hard thresholding estimator \hat{f}^H by

$$\hat{f}^H(x) = \sum_{k=0}^{2^\tau-1} \hat{\alpha}_{\tau,k} \phi_{\tau,k}(x) + \sum_{j=\tau}^{j_1} \sum_{k=0}^{2^j-1} \hat{\beta}_{j,k} \mathbb{I}_{\{|\hat{\beta}_{j,k}| \geq \kappa \lambda_n\}} \psi_{j,k}(x), \quad (4.6)$$

$x \in [0, 1]$, where $\hat{\alpha}_{\tau,k}$ is defined by (4.2) and $\hat{\beta}_{j,k}$ by (4.3), for any random event \mathcal{A} , $\mathbb{I}_{\mathcal{A}}$ is the indicator function on \mathcal{A} , j_1 is the integer satisfying

$$\frac{1}{2} \frac{n}{(\ln n)^{1+1/\theta}} < 2^{j_1} \leq \frac{n}{(\ln n)^{1+1/\theta}}, \quad (4.7)$$

θ is the one in (2.2), κ is a large enough constant (the one in Proposition 6.4 below) and λ_n is the threshold

$$\lambda_n = \sqrt{\frac{(\ln n)^{1+1/\theta}}{n}}. \quad (4.8)$$

The feature of the hard thresholding estimator is to only estimate the “large” unknown wavelet coefficients of f which contain his main characteristics.

For the construction of hard thresholding wavelet estimators in the standard density model, see, for example, [15, 23].

5. Results

Theorem 5.1 (upper bound for \hat{f}^L). *Consider (1.1) under the assumptions of Section 2. Suppose that $f \in B_{p,r}^s(M)$ with $s > 0$, $p \geq 2$, and $r \geq 1$. Let \hat{f}^L be (4.4). Then there exists a constant $C > 0$ such that*

$$\mathbb{E} \left(\int_0^1 (\hat{f}^L(x) - f(x))^2 dx \right) \leq C n^{-2s/(2s+1)}. \quad (5.1)$$

The proof of Theorem 5.1 uses a suitable decomposition of the MISE and a moment inequality on (4.2) (see Proposition 6.3 below).

Note that $n^{-2s/(2s+1)}$ is the optimal rate of convergence (in the minimax sense) for the standard density model in the independent case (see, e.g., [14, 23]).

Theorem 5.2 (upper bound for \hat{f}^H). *Consider (1.1) under the assumptions of Section 2. Let \hat{f}^H be (4.6). Suppose that $f \in B_{p,r}^s(M)$ with $r \geq 1$, $\{p \geq 2 \text{ and } s > 0\}$ or $\{p \in [1, 2) \text{ and } s > 1/p\}$. Then there exists a constant $C > 0$ such that*

$$\mathbb{E} \left(\int_0^1 (\hat{f}^H(x) - f(x))^2 dx \right) \leq C \left(\frac{(\ln n)^{1+1/\theta}}{n} \right)^{2s/(2s+1)}. \quad (5.2)$$

The proof of Theorem 5.2 uses a suitable decomposition of the MISE, some moment inequalities on (4.2) and (4.3) (see Proposition 6.3 below), and a concentration inequality on (4.3) (see Proposition 6.4 below).

Theorem 5.2 shows that, besides being adaptive, \hat{f}^H attains a rate of convergence close to the one of \hat{f}^L . The only difference is the logarithmic term $(\ln n)^{(1+1/\theta)(2s/(2s+1))}$.

Note that, if we restrict our study to the independent case, that is, $\theta \rightarrow \infty$, the rate of convergence attained by \hat{f}^H becomes the standard one: $(\log n/n)^{2s/(2s+1)}$. See, for example, [14, 15, 23].

6. Proofs

In this section, we consider (1.1) under the assumptions of Section 2. Moreover, C denotes any constant that does not depend on j , k and n . Its value may change from one term to another and may depend on ϕ or ψ .

6.1. Auxiliary Results

Lemma 6.1. For any integer $j \geq \tau$ and any $k \in \{0, \dots, 2^j - 1\}$, let $\hat{\alpha}_{j,k}$ be (4.2) and $\alpha_{j,k} = \int_0^1 f(x)\phi_{j,k}(x)dx$. Then, under the assumptions of Section 2, there exists a constant $C > 0$ such that

$$|\hat{\alpha}_{j,k} - \alpha_{j,k}| \leq C \left(\left| \frac{\mu}{n} \sum_{i=1}^n \frac{\phi_{j,k}(Z_i)}{w(Z_i)} - \alpha_{j,k} \right| + \left| \frac{1}{\hat{\mu}} - \frac{1}{\mu} \right| \right). \quad (6.1)$$

This inequality holds for ψ instead of ϕ (and, a fortiori, $\hat{\beta}_{j,k}$ defined by (4.3) instead of $\hat{\alpha}_{j,k}$ and $\beta_{j,k} = \int_0^1 f(x)\psi_{j,k}(x)dx$ instead of $\alpha_{j,k}$).

Proof of Lemma 6.1. We have

$$\hat{\alpha}_{j,k} - \alpha_{j,k} = \frac{\hat{\mu}}{\mu} \left(\frac{\mu}{n} \sum_{i=1}^n \frac{\phi_{j,k}(Z_i)}{w(Z_i)} - \alpha_{j,k} \right) + \alpha_{j,k} \hat{\mu} \left(\frac{1}{\hat{\mu}} - \frac{1}{\mu} \right). \quad (6.2)$$

Due to (2.3), we have $|\hat{\mu}| \leq C$ and $|\hat{\mu}/\mu| \leq C$. Therefore

$$|\hat{\alpha}_{j,k} - \alpha_{j,k}| \leq C \left(\left| \frac{\mu}{n} \sum_{i=1}^n \frac{\phi_{j,k}(Z_i)}{w(Z_i)} - \alpha_{j,k} \right| + |\alpha_{j,k}| \left| \frac{1}{\hat{\mu}} - \frac{1}{\mu} \right| \right). \quad (6.3)$$

Using (2.4) and the Cauchy-Schwarz inequality, we obtain

$$\begin{aligned} |\alpha_{j,k}| &\leq \int_0^1 f(x) |\phi_{j,k}(x)| dx \leq C \int_0^1 |\phi_{j,k}(x)| dx \\ &\leq C \left(\int_0^1 (\phi_{j,k}(x))^2 dx \right)^{1/2} = C. \end{aligned} \quad (6.4)$$

Hence

$$|\hat{\alpha}_{j,k} - \alpha_{j,k}| \leq C \left(\left| \frac{\mu}{n} \sum_{i=1}^n \frac{\phi_{j,k}(Z_i)}{w(Z_i)} - \alpha_{j,k} \right| + \left| \frac{1}{\hat{\mu}} - \frac{1}{\mu} \right| \right). \quad (6.5)$$

Lemma 6.1 is proved. \square

Proposition 6.2. For any integer $j \geq \tau$ such that $2^j \leq n$ and any $k \in \{0, \dots, 2^j - 1\}$, let $\alpha_{j,k} = \int_0^1 f(x)\phi_{j,k}(x)dx$ and $\hat{\mu}$ be (4.1). Then,

(1) one has

$$\mathbb{E}\left(\frac{\mu}{n} \sum_{i=1}^n \frac{\phi_{j,k}(Z_i)}{w(Z_i)}\right) = \alpha_{j,k}, \quad \mathbb{E}\left(\frac{1}{\hat{\mu}}\right) = \frac{1}{\mu}, \quad (6.6)$$

(2) there exists a constant $C > 0$ such that

$$\mathbb{V}\left(\frac{\mu}{n} \sum_{i=1}^n \frac{\phi_{j,k}(Z_i)}{w(Z_i)}\right) \leq C \frac{1}{n}, \quad (6.7)$$

(3) there exists a constant $C > 0$ such that

$$\mathbb{V}\left(\frac{1}{\hat{\mu}}\right) \leq C \frac{1}{n}. \quad (6.8)$$

These results hold for ψ instead of ϕ (and, a fortiori, $\beta_{j,k} = \int_0^1 f(x)\psi_{j,k}(x)dx$ instead of $\alpha_{j,k}$).

Proof of Proposition 6.2. (1) We have

$$\begin{aligned} \mathbb{E}\left(\frac{\mu}{n} \sum_{i=1}^n \frac{\phi_{j,k}(Z_i)}{w(Z_i)}\right) &= \mu \mathbb{E}\left(\frac{\phi_{j,k}(Z_1)}{w(Z_1)}\right) = \mu \int_0^1 \frac{\phi_{j,k}(x)}{w(x)} g(x) dx \\ &= \mu \int_0^1 \frac{\phi_{j,k}(x)}{w(x)} \mu^{-1} w(x) f(x) dx = \int_0^1 f(x) \phi_{j,k}(x) dx = \alpha_{j,k}. \end{aligned} \quad (6.9)$$

Since f is a density, we obtain

$$\begin{aligned} \mathbb{E}\left(\frac{1}{\hat{\mu}}\right) &= \mathbb{E}\left(\frac{1}{n} \sum_{i=1}^n \frac{1}{w(Z_i)}\right) = \mathbb{E}\left(\frac{1}{w(Z_1)}\right) = \int_0^1 \frac{1}{w(x)} g(x) dx \\ &= \int_0^1 \frac{1}{w(x)} \mu^{-1} w(x) f(x) dx = \frac{1}{\mu} \int_0^1 f(x) dx = \frac{1}{\mu}. \end{aligned} \quad (6.10)$$

(2) We have

$$\begin{aligned}
 \mathbb{V}\left(\frac{\mu}{n}\sum_{i=1}^n\frac{\phi_{j,k}(Z_i)}{w(Z_i)}\right) &= \frac{\mu^2}{n^2}\sum_{v=1}^n\sum_{\ell=1}^n\mathbb{C}\left(\frac{\phi_{j,k}(Z_v)}{w(Z_v)},\frac{\phi_{j,k}(Z_\ell)}{w(Z_\ell)}\right) \\
 &= \frac{\mu^2}{n}\mathbb{V}\left(\frac{\phi_{j,k}(Z_1)}{w(Z_1)}\right) + 2\frac{\mu^2}{n^2}\sum_{v=2}^n\sum_{\ell=1}^{v-1}\mathbb{C}\left(\frac{\phi_{j,k}(Z_v)}{w(Z_v)},\frac{\phi_{j,k}(Z_\ell)}{w(Z_\ell)}\right) \\
 &\leq \frac{\mu^2}{n}\mathbb{V}\left(\frac{\phi_{j,k}(Z_1)}{w(Z_1)}\right) + 2\frac{\mu^2}{n^2}\left|\sum_{v=2}^n\sum_{\ell=1}^{v-1}\mathbb{C}\left(\frac{\phi_{j,k}(Z_v)}{w(Z_v)},\frac{\phi_{j,k}(Z_\ell)}{w(Z_\ell)}\right)\right|.
 \end{aligned} \tag{6.11}$$

Using (2.3) and (2.4), we have $\sup_{x\in[0,1]}g(x) \leq C$. Hence,

$$\begin{aligned}
 \mathbb{V}\left(\frac{\phi_{j,k}(Z_1)}{w(Z_1)}\right) &\leq \mathbb{E}\left(\left(\frac{\phi_{j,k}(Z_1)}{w(Z_1)}\right)^2\right) \leq C\mathbb{E}\left((\phi_{j,k}(Z_1))^2\right) \\
 &= C\int_0^1(\phi_{j,k}(x))^2g(x)dx \leq C\int_0^1(\phi_{j,k}(x))^2dx = C.
 \end{aligned} \tag{6.12}$$

It follows from the stationarity of $(Z_i)_{i\in\mathbb{Z}}$ and $2^j \leq n$ that

$$\begin{aligned}
 \left|\sum_{v=2}^n\sum_{\ell=1}^{v-1}\mathbb{C}\left(\frac{\phi_{j,k}(Z_v)}{w(Z_v)},\frac{\phi_{j,k}(Z_\ell)}{w(Z_\ell)}\right)\right| &= \left|\sum_{m=1}^n(n-m)\mathbb{C}\left(\frac{\phi_{j,k}(Z_0)}{w(Z_0)},\frac{\phi_{j,k}(Z_m)}{w(Z_m)}\right)\right| \\
 &\leq n\sum_{m=1}^n\left|\mathbb{C}\left(\frac{\phi_{j,k}(Z_0)}{w(Z_0)},\frac{\phi_{j,k}(Z_m)}{w(Z_m)}\right)\right| = T_1 + T_2,
 \end{aligned} \tag{6.13}$$

where

$$\begin{aligned}
 T_1 &= n\sum_{m=1}^{2^j-1}\left|\mathbb{C}\left(\frac{\phi_{j,k}(Z_0)}{w(Z_0)},\frac{\phi_{j,k}(Z_m)}{w(Z_m)}\right)\right|, \\
 T_2 &= n\sum_{m=2^j}^n\left|\mathbb{C}\left(\frac{\phi_{j,k}(Z_0)}{w(Z_0)},\frac{\phi_{j,k}(Z_m)}{w(Z_m)}\right)\right|.
 \end{aligned} \tag{6.14}$$

Let us now bound T_1 and T_2 .

Upper Bound for T_1

Using (2.5), (2.3), and doing the change a variables $y = 2^j x - k$, we obtain

$$\begin{aligned} \left| \mathbb{C} \left(\frac{\phi_{j,k}(Z_0)}{w(Z_0)}, \frac{\phi_{j,k}(Z_m)}{w(Z_m)} \right) \right| &= \left| \iint_0^1 (g_{(Z_0, Z_m)}(x, y) - g(x)g(y)) \frac{\phi_{j,k}(x)}{w(x)} \frac{\phi_{j,k}(y)}{w(y)} dx dy \right| \\ &\leq \iint_0^1 |g_{(Z_0, Z_m)}(x, y) - g(x)g(y)| \left| \frac{\phi_{j,k}(x)}{w(x)} \right| \left| \frac{\phi_{j,k}(y)}{w(y)} \right| dx dy \\ &\leq C \left(\int_0^1 |\phi_{j,k}(x)| dx \right)^2 = C \left(2^{-j/2} \int_0^1 |\phi(x)| dx \right)^2 = C 2^{-j}. \end{aligned} \quad (6.15)$$

Therefore,

$$T_1 \leq C n 2^{-j} = C n. \quad (6.16)$$

Upper Bound for T_2

By the Davydov inequality for strongly mixing processes (see [24]), for any $q \in (0, 1)$, it holds that

$$\begin{aligned} \left| \mathbb{C} \left(\frac{\phi_{j,k}(Z_0)}{w(Z_0)}, \frac{\phi_{j,k}(Z_m)}{w(Z_m)} \right) \right| &\leq 10 a_m^q \left(\mathbb{E} \left(\left| \frac{\phi_{j,k}(Z_0)}{w(Z_0)} \right|^{2/(1-q)} \right) \right)^{1-q} \\ &\leq 10 a_m^q \left(\sup_{x \in [0,1]} \left| \frac{\phi_{j,k}(x)}{w(x)} \right| \right)^{2q} \left(\mathbb{E} \left(\left(\frac{\phi_{j,k}(Z_0)}{w(Z_0)} \right)^2 \right) \right)^{1-q}. \end{aligned} \quad (6.17)$$

By (2.3), we have

$$\sup_{x \in [0,1]} \left| \frac{\phi_{j,k}(x)}{w(x)} \right| \leq C \sup_{x \in [0,1]} |\phi_{j,k}(x)| \leq C 2^{j/2} \quad (6.18)$$

and, by (6.12),

$$\mathbb{E} \left(\left(\frac{\phi_{j,k}(Z_0)}{w(Z_0)} \right)^2 \right) \leq C. \quad (6.19)$$

Therefore,

$$\left| \mathbb{C} \left(\frac{\phi_{j,k}(Z_0)}{w(Z_0)}, \frac{\phi_{j,k}(Z_m)}{w(Z_m)} \right) \right| \leq C 2^{qj} a_m^q. \quad (6.20)$$

Since $\sum_{m=2^j}^n m^q a_m^q \leq \sum_{m=1}^\infty m^q a_m^q = \gamma^q \sum_{m=1}^\infty m^q \exp(-cqm^\theta) < \infty$, we have

$$T_2 \leq Cn2^{qj} \sum_{m=2^j}^n a_m^q \leq Cn \sum_{m=2^j}^n m^q a_m^q \leq Cn. \tag{6.21}$$

It follows from (6.13), (6.16), and (6.21) that

$$\left| \sum_{v=2}^n \sum_{\ell=1}^{v-1} \mathbb{C} \left(\frac{\phi_{j,k}(Z_v)}{w(Z_v)}, \frac{\phi_{j,k}(Z_\ell)}{w(Z_\ell)} \right) \right| \leq Cn. \tag{6.22}$$

Combining (6.11), (6.12), and (6.22), we obtain

$$\mathbb{V} \left(\frac{\mu}{n} \sum_{i=1}^n \frac{\phi_{j,k}(Z_i)}{w(Z_i)} \right) \leq C \frac{1}{n}. \tag{6.23}$$

(3) Proceeding in a similar fashion to 2-, we obtain

$$\begin{aligned} \mathbb{V} \left(\frac{1}{\hat{\mu}} \right) &= \mathbb{V} \left(\frac{1}{n} \sum_{i=1}^n \frac{1}{w(Z_i)} \right) \\ &= \frac{1}{n} \mathbb{V} \left(\frac{1}{w(Z_1)} \right) + 2 \frac{1}{n^2} \sum_{v=2}^n \sum_{\ell=1}^{v-1} \mathbb{C} \left(\frac{1}{w(Z_v)}, \frac{1}{w(Z_\ell)} \right) \\ &\leq \frac{1}{n} \mathbb{V} \left(\frac{1}{w(Z_1)} \right) + 2 \frac{1}{n} \sum_{m=1}^n \left| \mathbb{C} \left(\frac{1}{w(Z_0)}, \frac{1}{w(Z_m)} \right) \right|. \end{aligned} \tag{6.24}$$

Using (2.3) (which implies $\sup_{x \in [0,1]} (1/w(x)) \leq C$) and applying the Davydov inequality, we obtain

$$\mathbb{V} \left(\frac{1}{\hat{\mu}} \right) \leq C \frac{1}{n} \left(1 + \sum_{m=1}^n a_m^q \right) \leq C \frac{1}{n}. \tag{6.25}$$

The proof of Proposition 6.2 is complete. □

Proposition 6.3. For any integer $j \geq \tau$ such that $2^j \leq n$ and any $k \in \{0, \dots, 2^j - 1\}$, let $\alpha_{j,k} = \int_0^1 f(x)\phi_{j,k}(x)dx$ and $\hat{\alpha}_{j,k}$ be (4.2). Then,

(1) there exists a constant $C > 0$ such that

$$\mathbb{E} \left((\hat{\alpha}_{j,k} - \alpha_{j,k})^2 \right) \leq C \frac{1}{n}; \tag{6.26}$$

(2) there exists a constant $C > 0$ such that

$$\mathbb{E} \left((\hat{\alpha}_{j,k} - \alpha_{j,k})^4 \right) \leq C 2^j \frac{1}{n}. \tag{6.27}$$

These inequalities hold for $\widehat{\beta}_{j,k}$ defined by (4.3) instead of $\widehat{\alpha}_{j,k}$, and $\beta_{j,k} = \int_0^1 f(x)\psi_{j,k}(x)dx$ instead of $\alpha_{j,k}$.

Proof of Proposition 6.3. (1) Applying Lemma 6.1 and Proposition 6.2, we have

$$\begin{aligned} \mathbb{E}\left((\widehat{\alpha}_{j,k} - \alpha_{j,k})^2\right) &\leq C\left(\mathbb{E}\left(\left(\frac{\mu}{n}\sum_{i=1}^n\frac{\phi_{j,k}(Z_i)}{w(Z_i)} - \alpha_{j,k}\right)^2\right) + \mathbb{E}\left(\left(\frac{1}{\widehat{\mu}} - \frac{1}{\mu}\right)^2\right)\right) \\ &= C\left(\mathbb{V}\left(\frac{\mu}{n}\sum_{i=1}^n\frac{\phi_{j,k}(Z_i)}{w(Z_i)}\right) + \mathbb{V}\left(\frac{1}{\widehat{\mu}}\right)\right) \leq C\frac{1}{n}. \end{aligned} \tag{6.28}$$

(2) We have

$$|\widehat{\alpha}_{j,k} - \alpha_{j,k}| \leq |\widehat{\alpha}_{j,k}| + |\alpha_{j,k}|. \tag{6.29}$$

By (2.3), we have $|\widehat{\mu}| \leq C$ and $\sup_{x \in [0,1]} (1/w(x)) \leq C$. So,

$$\begin{aligned} \left|\frac{\widehat{\mu}}{n}\sum_{i=1}^n\frac{\phi_{j,k}(Z_i)}{w(Z_i)}\right| &\leq C\frac{1}{n}\sum_{i=1}^n\left|\frac{\phi_{j,k}(Z_i)}{w(Z_i)}\right| \leq C\sup_{x \in [0,1]}\left|\frac{\phi_{j,k}(x)}{w(x)}\right| \\ &\leq C\sup_{x \in [0,1]}|\phi_{j,k}(x)| \leq C2^{j/2}. \end{aligned} \tag{6.30}$$

By (6.4), we have $|\alpha_{j,k}| \leq C$. Therefore

$$|\widehat{\alpha}_{j,k} - \alpha_{j,k}| \leq C(2^{j/2} + 1) \leq C2^{j/2}. \tag{6.31}$$

It follows from (6.31) and (6.28) that

$$\mathbb{E}\left((\widehat{\alpha}_{j,k} - \alpha_{j,k})^4\right) \leq C2^j\mathbb{E}\left((\widehat{\alpha}_{j,k} - \alpha_{j,k})^2\right) \leq C2^j\frac{1}{n}. \tag{6.32}$$

The proof of Proposition 6.3 is complete. □

Proposition 6.4. For any $j \in \{\tau, \dots, j_1\}$ and any $k \in \{0, \dots, 2^j - 1\}$, let $\beta_{j,k} = \int_0^1 f(x)\psi_{j,k}(x)dx$, $\widehat{\beta}_{j,k}$ be (4.3) and λ_n be (4.8). Then there exist two constants, $\kappa > 0$ and $C > 0$, such that

$$\mathbb{P}\left(\left|\widehat{\beta}_{j,k} - \beta_{j,k}\right| \geq \frac{\kappa\lambda_n}{2}\right) \leq C\frac{1}{n^4}. \tag{6.33}$$

Proof of Proposition 6.4. It follows from Lemma 6.1 that

$$\mathbb{P}\left(\left|\widehat{\beta}_{j,k} - \beta_{j,k}\right| \geq \frac{\kappa\lambda_n}{2}\right) \leq P_1 + P_2, \tag{6.34}$$

where

$$P_1 = \mathbb{P} \left(\left| \frac{\mu}{n} \sum_{i=1}^n \frac{\psi_{j,k}(Z_i)}{w(Z_i)} - \beta_{j,k} \right| \geq \kappa C \lambda_n \right), \quad (6.35)$$

$$P_2 = \mathbb{P} \left(\left| \frac{1}{\hat{\mu}} - \frac{1}{\mu} \right| \geq \kappa C \lambda_n \right).$$

In order to bound P_1 and P_2 , let us present a Bernstein inequality for exponentially strongly mixing process. We refer to [25, 26].

Lemma 6.5 (see [25, 26]). *Let $\gamma > 0$, $c > 0$, $\theta > 1$ and $(Z_i)_{i \in \mathbb{Z}}$ be a stationary process such that, for any $m \in \mathbb{Z}$, the associated m th strongly mixing coefficient (2.2) satisfies $a_m \leq \gamma \exp(-c|m|^\theta)$. Let $n \in \mathbb{N}^*$, $h : \mathbb{R} \rightarrow \mathbb{R}$ be a measurable function and, for any $i \in \mathbb{Z}$, $U_i = h(Z_i)$. One assumes that $\mathbb{E}(U_1) = 0$ and there exists a constant $M > 0$ satisfying $|U_1| \leq M < \infty$. Then, for any $m \in \{1, \dots, n\}$ and any $\lambda > 4mM/n$, one has*

$$\mathbb{P} \left(\left| \frac{1}{n} \sum_{i=1}^n U_i \right| \geq \lambda \right) \leq 4 \exp \left(- \frac{\lambda^2 n}{m(64\mathbb{E}(U_1^2) + 8\lambda M/3)} \right) + 4\gamma \frac{n}{m} \exp(-cm^\theta). \quad (6.36)$$

Upper Bound for P_1

For any $i \in \{1, \dots, n\}$, set

$$U_i = \mu \frac{\psi_{j,k}(Z_i)}{w(Z_i)} - \beta_{j,k}. \quad (6.37)$$

Then U_1, \dots, U_n are identically distributed, depend on the stationary strongly mixing process $(Z_i)_{i \in \mathbb{Z}}$ which satisfies (2.2), Proposition 6.2 gives

$$\mathbb{E}(U_1) = 0, \quad \mathbb{E}(U_1^2) \leq \mathbb{E} \left(\left(\mu \frac{\psi_{j,k}(Z_1)}{w(Z_1)} \right)^2 \right) \leq C \quad (6.38)$$

and, by (2.3) and (6.4),

$$|U_1| \leq \mu \sup_{x \in [0,1]} \left| \frac{\psi_{j,k}(x)}{w(x)} \right| + |\beta_{j,k}| \leq C \left(\sup_{x \in [0,1]} |\psi_{j,k}(x)| + 1 \right) \quad (6.39)$$

$$\leq C(2^{j/2} + 1) \leq C2^{j/2}.$$

It follows from Lemma 6.5 applied with U_1, \dots, U_n , $\lambda = \kappa C \lambda_n$, $\lambda_n = ((\ln n)^{1+1/\theta} / n)^{1/2}$, $m = (u \ln n)^{1/\theta}$ with $u > 0$ (chosen later), $M = C2^{j/2}$ and $2^j \leq 2^i \leq n / (\ln n)^{1+1/\theta}$, that

$$\begin{aligned}
 P_1 &= \mathbb{P}\left(\left|\frac{1}{n} \sum_{i=1}^n U_i\right| \geq \kappa C \lambda_n\right) \\
 &\leq 4 \exp\left(-C \frac{\kappa^2 \lambda_n^2 n}{m(1 + \kappa \lambda_n M)}\right) + 4\gamma \frac{n}{m} \exp(-cm^\theta) \\
 &\leq 4 \exp\left(-C \frac{\kappa^2 (\ln n)^{1+1/\theta}}{(u \ln n)^{1/\theta} \left(1 + \kappa 2^{j/2} \left((\ln n)^{1+1/\theta} / n\right)^{1/2}\right)}\right) \\
 &\quad + 4\gamma \frac{n}{(u \ln n)^{1/\theta}} \exp(-cu \ln n) \\
 &\leq C \left(n^{-C\kappa^2 / (u^{1/\theta}(1+\kappa))} + n^{1-cu}\right).
 \end{aligned} \tag{6.40}$$

Therefore, for large enough κ and u , we have

$$P_1 \leq C \frac{1}{n^4}. \tag{6.41}$$

Upper Bound for P_2

For any $i \in \{1, \dots, n\}$, set

$$U_i = \frac{1}{w(Z_i)} - \frac{1}{\mu}. \tag{6.42}$$

Then U_1, \dots, U_n are identically distributed, depend on the stationary strongly mixing process $(Z_i)_{i \in \mathbb{Z}}$ which satisfies (2.2), Proposition 6.2 gives

$$\mathbb{E}(U_1) = 0, \quad \mathbb{E}(U_1^2) \leq \mathbb{E}\left(\frac{1}{(w(Z_1))^2}\right) \leq C. \tag{6.43}$$

By (2.3), we have

$$|U_1| \leq \sup_{x \in [0,1]} \frac{1}{w(x)} + \left|\frac{1}{\mu}\right| \leq C. \tag{6.44}$$

It follows from Lemma 6.5 applied with U_1, \dots, U_n , $\lambda = \kappa C \lambda_n$, $\lambda_n = ((\ln n)^{1+1/\theta}/n)^{1/2}$, $m = (u \ln n)^{1/\theta}$ with $u > 0$ (chosen later) and $M = C$ that

$$\begin{aligned}
 P_2 &= \mathbb{P} \left(\left| \frac{1}{n} \sum_{i=1}^n U_i \right| \geq \kappa C \lambda_n \right) \\
 &\leq 4 \exp \left(-C \frac{\kappa^2 \lambda_n^2 n}{m(1 + \kappa \lambda_n M)} \right) + 4\gamma \frac{n}{m} \exp(-cm^\theta) \\
 &\leq 4 \exp \left(-C \frac{\kappa^2 (\ln n)^{1+1/\theta}}{(u \ln n)^{1/\theta} \left(1 + \kappa \left((\ln n)^{1+1/\theta}/n \right)^{1/2} \right)} \right) \\
 &\quad + 4\gamma \frac{n}{(u \ln n)^{1/\theta}} \exp(-cu \ln n) \\
 &\leq C \left(n^{-C\kappa^2/u^{1/\theta}} + n^{1-cu} \right).
 \end{aligned} \tag{6.45}$$

Therefore, for large enough κ and u , we have

$$P_2 \leq C \frac{1}{n^4}. \tag{6.46}$$

Putting (6.34), (6.41), and (6.46) together, this ends the proof of Proposition 6.4. \square

6.2. Proofs of the Main Results

Proof of Theorem 5.1. We expand the function f on \mathcal{B} as

$$f(x) = \sum_{k=0}^{2^{j_0}-1} \alpha_{j_0,k} \phi_{j_0,k}(x) + \sum_{j=j_0}^{\infty} \sum_{k=0}^{2^j-1} \beta_{j,k} \psi_{j,k}(x), \quad x \in [0, 1], \tag{6.47}$$

where $\alpha_{j_0,k} = \int_0^1 f(x) \phi_{j_0,k}(x) dx$ and $\beta_{j,k} = \int_0^1 f(x) \psi_{j,k}(x) dx$.

We have, for any $x \in [0, 1]$,

$$\widehat{f}^L(x) - f(x) = \sum_{k=0}^{2^{j_0}-1} (\widehat{\alpha}_{j_0,k} - \alpha_{j_0,k}) \phi_{j_0,k}(x) - \sum_{j=j_0}^{\infty} \sum_{k=0}^{2^j-1} \beta_{j,k} \psi_{j,k}(x). \tag{6.48}$$

Since \mathcal{B} is an orthonormal basis of $\mathbb{L}^2([0, 1])$, we have,

$$\mathbb{E} \left(\int_0^1 (\widehat{f}^L(x) - f(x))^2 dx \right) = \sum_{k=0}^{2^{j_0}-1} \mathbb{E} \left((\widehat{\alpha}_{j_0,k} - \alpha_{j_0,k})^2 \right) + \sum_{j=j_0}^{\infty} \sum_{k=0}^{2^j-1} \beta_{j,k}^2. \tag{6.49}$$

Using Proposition 6.3, we obtain

$$\sum_{k=0}^{2^{j_0}-1} \mathbb{E} \left((\hat{\alpha}_{j_0,k} - \alpha_{j_0,k})^2 \right) \leq C 2^{j_0} \frac{1}{n} \leq C n^{-2s/(2s+1)}. \tag{6.50}$$

Since $p \geq 2$, we have $B_{p,r}^s(M) \subseteq B_{2,\infty}^s(M)$. Hence

$$\sum_{j=j_0}^{\infty} \sum_{k=0}^{2^j-1} \beta_{j,k}^2 \leq C 2^{-2j_0s} \leq C n^{-2s/(2s+1)}. \tag{6.51}$$

Therefore,

$$\mathbb{E} \left(\int_0^1 (\hat{f}^L(x) - f(x))^2 dx \right) \leq C n^{-2s/(2s+1)}. \tag{6.52}$$

The proof of Theorem 5.1 is complete. □

Proof of Theorem 5.2. We expand the function f on \mathcal{B} as

$$f(x) = \sum_{k=0}^{2^\tau-1} \alpha_{\tau,k} \phi_{\tau,k}(x) + \sum_{j=\tau}^{\infty} \sum_{k=0}^{2^j-1} \beta_{j,k} \psi_{j,k}(x), \quad x \in [0, 1], \tag{6.53}$$

where $\alpha_{\tau,k} = \int_0^1 f(x) \phi_{\tau,k}(x) dx$ and $\beta_{j,k} = \int_0^1 f(x) \psi_{j,k}(x) dx$.
We have, for any $x \in [0, 1]$,

$$\begin{aligned} \hat{f}^H(x) - f(x) &= \sum_{k=0}^{2^{\tau-1}} (\hat{\alpha}_{\tau,k} - \alpha_{\tau,k}) \phi_{\tau,k}(x) + \sum_{j=\tau}^{j_1} \sum_{k=0}^{2^j-1} (\hat{\beta}_{j,k} \mathbb{I}_{\{|\hat{\beta}_{j,k}| \geq \kappa \lambda_n\}} - \beta_{j,k}) \psi_{j,k}(x) \\ &\quad - \sum_{j=j_1+1}^{\infty} \sum_{k=0}^{2^j-1} \beta_{j,k} \psi_{j,k}(x). \end{aligned} \tag{6.54}$$

Since \mathcal{B} is an orthonormal basis of $\mathbb{L}^2([0, 1])$, we have

$$\mathbb{E} \left(\int_0^1 (\hat{f}^H(x) - f(x))^2 dx \right) = R + S + T, \tag{6.55}$$

where

$$\begin{aligned}
 R &= \sum_{k=0}^{2^{\tau}-1} \mathbb{E} \left((\hat{\alpha}_{\tau,k} - \alpha_{\tau,k})^2 \right), & S &= \sum_{j=\tau}^{j_1} \sum_{k=0}^{2^j-1} \mathbb{E} \left(\left(\hat{\beta}_{j,k} \mathbb{I}_{\{\|\hat{\beta}_{j,k}\| \geq \kappa \lambda_n\}} - \beta_{j,k} \right)^2 \right), \\
 T &= \sum_{j=j_1+1}^{\infty} \sum_{k=0}^{2^j-1} \beta_{j,k}^2.
 \end{aligned} \tag{6.56}$$

Let us bound R , T , and S , in turn.

Upper Bound for R

Using Proposition 6.3 and $2s/(2s+1) < 1$, we obtain

$$R \leq C 2^{\tau} \frac{1}{n} \leq C \left(\frac{(\ln n)^{1+1/\theta}}{n} \right)^{2s/(2s+1)}. \tag{6.57}$$

Upper Bound for T

For $r \geq 1$ and $p \geq 2$, we have $B_{p,r}^s(M) \subseteq B_{2,\infty}^s(M)$. Since $2s/(2s+1) < 2s$, we have

$$T \leq C \sum_{j=j_1+1}^{\infty} 2^{-2js} \leq C 2^{-2j_1 s} \leq C \left(\frac{(\ln n)^{1+1/\theta}}{n} \right)^{2s} \leq C \left(\frac{(\ln n)^{1+1/\theta}}{n} \right)^{2s/(2s+1)}. \tag{6.58}$$

For $r \geq 1$ and $p \in [1, 2)$, we have $B_{p,r}^s(M) \subseteq B_{2,\infty}^{s+1/2-1/p}(M)$. Since $s > 1/p$, we have $s + 1/2 - 1/p > s/(2s+1)$. So

$$\begin{aligned}
 T &\leq C \sum_{j=j_1+1}^{\infty} 2^{-2j(s+1/2-1/p)} \leq C 2^{-2j_1(s+1/2-1/p)} \\
 &\leq C \left(\frac{(\ln n)^{1+1/\theta}}{n} \right)^{2(s+1/2-1/p)} \leq C \left(\frac{(\ln n)^{1+1/\theta}}{n} \right)^{2s/(2s+1)}.
 \end{aligned} \tag{6.59}$$

Hence, for $r \geq 1$, $\{p \geq 2 \text{ and } s > 0\}$ or $\{p \in [1, 2) \text{ and } s > 1/p\}$, we have

$$T \leq C \left(\frac{(\ln n)^{1+1/\theta}}{n} \right)^{2s/(2s+1)}. \tag{6.60}$$

Upper Bound for S

Note that we can write the term S as

$$S = S_1 + S_2 + S_3 + S_4, \tag{6.61}$$

where

$$\begin{aligned}
 S_1 &= \sum_{j=\tau}^{j_1} \sum_{k=0}^{2^j-1} \mathbb{E} \left(\left(\widehat{\beta}_{j,k} - \beta_{j,k} \right)^2 \mathbb{I}_{\{|\widehat{\beta}_{j,k}| \geq \kappa \lambda_n\}} \mathbb{I}_{\{|\beta_{j,k}| < \kappa \lambda_n / 2\}} \right), \\
 S_2 &= \sum_{j=\tau}^{j_1} \sum_{k=0}^{2^j-1} \mathbb{E} \left(\left(\widehat{\beta}_{j,k} - \beta_{j,k} \right)^2 \mathbb{I}_{\{|\widehat{\beta}_{j,k}| \geq \kappa \lambda_n\}} \mathbb{I}_{\{|\beta_{j,k}| \geq \kappa \lambda_n / 2\}} \right), \\
 S_3 &= \sum_{j=\tau}^{j_1} \sum_{k=0}^{2^j-1} \mathbb{E} \left(\beta_{j,k}^2 \mathbb{I}_{\{|\widehat{\beta}_{j,k}| < \kappa \lambda_n\}} \mathbb{I}_{\{|\beta_{j,k}| \geq 2\kappa \lambda_n\}} \right), \\
 S_4 &= \sum_{j=\tau}^{j_1} \sum_{k=0}^{2^j-1} \mathbb{E} \left(\beta_{j,k}^2 \mathbb{I}_{\{|\widehat{\beta}_{j,k}| < \kappa \lambda_n\}} \mathbb{I}_{\{|\beta_{j,k}| < 2\kappa \lambda_n\}} \right).
 \end{aligned}
 \tag{6.62}$$

Let us investigate the bounds of $S_1, S_2, S_3,$ and S_4 in turn.

Upper Bounds for S_1 and S_3

We have

$$\begin{aligned}
 \left\{ \left| \widehat{\beta}_{j,k} \right| < \kappa \lambda_n, \left| \beta_{j,k} \right| \geq 2\kappa \lambda_n \right\} &\subseteq \left\{ \left| \widehat{\beta}_{j,k} - \beta_{j,k} \right| > \frac{\kappa \lambda_n}{2} \right\}, \\
 \left\{ \left| \widehat{\beta}_{j,k} \right| \geq \kappa \lambda_n, \left| \beta_{j,k} \right| < \frac{\kappa \lambda_n}{2} \right\} &\subseteq \left\{ \left| \widehat{\beta}_{j,k} - \beta_{j,k} \right| > \frac{\kappa \lambda_n}{2} \right\}, \\
 \left\{ \left| \widehat{\beta}_{j,k} \right| < \kappa \lambda_n, \left| \beta_{j,k} \right| \geq 2\kappa \lambda_n \right\} &\subseteq \left\{ \left| \beta_{j,k} \right| \leq 2 \left| \widehat{\beta}_{j,k} - \beta_{j,k} \right| \right\}.
 \end{aligned}
 \tag{6.63}$$

So,

$$\max(S_1, S_3) \leq C \sum_{j=\tau}^{j_1} \sum_{k=0}^{2^j-1} \mathbb{E} \left(\left(\widehat{\beta}_{j,k} - \beta_{j,k} \right)^2 \mathbb{I}_{\{|\widehat{\beta}_{j,k} - \beta_{j,k}| > \kappa \lambda_n / 2\}} \right).
 \tag{6.64}$$

It follows from the Cauchy-Schwarz inequality, Propositions 6.3 and 6.4, and $2^j \leq 2^{j_1} \leq n$ that

$$\begin{aligned}
 \mathbb{E} \left(\left(\widehat{\beta}_{j,k} - \beta_{j,k} \right)^2 \mathbb{I}_{\{|\widehat{\beta}_{j,k} - \beta_{j,k}| > \kappa \lambda_n / 2\}} \right) &\leq \left(\mathbb{E} \left(\left(\widehat{\beta}_{j,k} - \beta_{j,k} \right)^4 \right) \right)^{1/2} \left(\mathbb{P} \left(\left| \widehat{\beta}_{j,k} - \beta_{j,k} \right| > \frac{\kappa \lambda_n}{2} \right) \right)^{1/2} \\
 &\leq C \left(2^j \frac{1}{n} \right)^{1/2} \left(\frac{1}{n^4} \right)^{1/2} \leq C \frac{1}{n^2}.
 \end{aligned}
 \tag{6.65}$$

Since $2s/(2s+1) < 1$, we have

$$\max(S_1, S_3) \leq C \frac{1}{n^2} \sum_{j=\tau}^{j_1} 2^j \leq C \frac{1}{n^2} 2^{j_1} \leq C \frac{1}{n} \leq C \left(\frac{(\ln n)^{1+1/\theta}}{n} \right)^{2s/(2s+1)}. \quad (6.66)$$

Upper Bound for S_2

Using again Proposition 6.3, we obtain

$$\mathbb{E} \left(\left(\hat{\beta}_{j,k} - \beta_{j,k} \right)^2 \right) \leq C \frac{1}{n} \leq C \frac{(\ln n)^{1+1/\theta}}{n}. \quad (6.67)$$

Hence,

$$S_2 \leq C \frac{(\ln n)^{1+1/\theta}}{n} \sum_{j=\tau}^{j_1} \sum_{k=0}^{2^j-1} \mathbb{I}_{\{|\beta_{j,k}| > \kappa \lambda_n / 2\}}. \quad (6.68)$$

Let j_2 be the integer defined by

$$\frac{1}{2} \left(\frac{n}{(\ln n)^{1+1/\theta}} \right)^{1/(2s+1)} < 2^{j_2} \leq \left(\frac{n}{(\ln n)^{1+1/\theta}} \right)^{1/(2s+1)}. \quad (6.69)$$

We have

$$S_2 \leq S_{2,1} + S_{2,2}, \quad (6.70)$$

where

$$\begin{aligned} S_{2,1} &= C \frac{(\ln n)^{1+1/\theta}}{n} \sum_{j=\tau}^{j_2} \sum_{k=0}^{2^j-1} \mathbb{I}_{\{|\beta_{j,k}| > \kappa \lambda_n / 2\}}, \\ S_{2,2} &= C \frac{(\ln n)^{1+1/\theta}}{n} \sum_{j=j_2+1}^{j_1} \sum_{k=0}^{2^j-1} \mathbb{I}_{\{|\beta_{j,k}| > \kappa \lambda_n / 2\}}. \end{aligned} \quad (6.71)$$

We have

$$S_{2,1} \leq C \frac{(\ln n)^{1+1/\theta}}{n} \sum_{j=\tau}^{j_2} 2^j \leq C \frac{(\ln n)^{1+1/\theta}}{n} 2^{j_2} \leq C \left(\frac{(\ln n)^{1+1/\theta}}{n} \right)^{2s/(2s+1)}. \quad (6.72)$$

For $r \geq 1$ and $p \geq 2$, since $B_{p,r}^s(M) \subseteq B_{2,\infty}^s(M)$,

$$\begin{aligned} S_{2,2} &\leq C \frac{(\ln n)^{1+1/\theta}}{n\lambda_n^2} \sum_{j=j_2+1}^{j_1} \sum_{k=0}^{2^j-1} \beta_{j,k}^2 \leq C \sum_{j=j_2+1}^{\infty} \sum_{k=0}^{2^j-1} \beta_{j,k}^2 \leq C 2^{-2j_2s} \\ &\leq C \left(\frac{(\ln n)^{1+1/\theta}}{n} \right)^{2s/(2s+1)}. \end{aligned} \tag{6.73}$$

For $r \geq 1, p \in [1, 2)$ and $s > 1/p$, using $\mathbb{I}_{\{|\beta_{j,k}| > \kappa\lambda_n/2\}} \leq C|\beta_{j,k}|^p/\lambda_n^p, B_{p,r}^s(M) \subseteq B_{2,\infty}^{s+1/2-1/p}(M)$ and $(2s+1)(2-p)/2 + (s+1/2-1/p)p = 2s$, we have

$$\begin{aligned} S_{2,2} &\leq C \frac{(\ln n)^{1+1/\theta}}{n\lambda_n^p} \sum_{j=j_2+1}^{j_1} \sum_{k=0}^{2^j-1} |\beta_{j,k}|^p \leq C \left(\frac{(\ln n)^{1+1/\theta}}{n} \right)^{(2-p)/2} \sum_{j=j_2+1}^{\infty} 2^{-j(s+1/2-1/p)p} \\ &\leq C \left(\frac{(\ln n)^{1+1/\theta}}{n} \right)^{(2-p)/2} 2^{-j_2(s+1/2-1/p)p} \leq C \left(\frac{(\ln n)^{1+1/\theta}}{n} \right)^{2s/(2s+1)}. \end{aligned} \tag{6.74}$$

So, for $r \geq 1, \{p \geq 2 \text{ and } s > 0\}$ or $\{p \in [1, 2) \text{ and } s > 1/p\}$, we have

$$S_2 \leq C \left(\frac{(\ln n)^{1+1/\theta}}{n} \right)^{2s/(2s+1)}. \tag{6.75}$$

Upper Bound for S_4

We have

$$S_4 \leq \sum_{j=\tau}^{j_1} \sum_{k=0}^{2^j-1} \beta_{j,k}^2 \mathbb{I}_{\{|\beta_{j,k}| < 2\kappa\lambda_n\}}. \tag{6.76}$$

Let j_2 be the integer (6.69). Then

$$S_4 \leq S_{4,1} + S_{4,2}, \tag{6.77}$$

where

$$S_{4,1} = \sum_{j=\tau}^{j_2} \sum_{k=0}^{2^j-1} \beta_{j,k}^2 \mathbb{I}_{\{|\beta_{j,k}| < 2\kappa\lambda_n\}}, \quad S_{4,2} = \sum_{j=j_2+1}^{j_1} \sum_{k=0}^{2^j-1} \beta_{j,k}^2 \mathbb{I}_{\{|\beta_{j,k}| < 2\kappa\lambda_n\}}. \tag{6.78}$$

We have

$$S_{4,1} \leq C \sum_{j=\tau}^{j_2} 2^j \lambda_n^2 = C \frac{(\ln n)^{1+1/\theta}}{n} \sum_{j=\tau}^{j_2} 2^j \leq C \frac{(\ln n)^{1+1/\theta}}{n} 2^{j_2} \leq C \left(\frac{(\ln n)^{1+1/\theta}}{n} \right)^{2s/(2s+1)}. \tag{6.79}$$

For $r \geq 1$ and $p \geq 2$, since $B_{p,r}^s(M) \subseteq B_{2,\infty}^s(M)$, we have

$$S_{4,2} \leq \sum_{j=j_2+1}^{\infty} \sum_{k=0}^{2^j-1} \beta_{j,k}^2 \leq C 2^{-2j_2 s} \leq C \left(\frac{(\ln n)^{1+1/\theta}}{n} \right)^{2s/(2s+1)}. \quad (6.80)$$

For $r \geq 1$, $p \in [1, 2)$ and $s > 1/p$, using $\beta_{j,k}^2 \mathbb{I}_{\{|\beta_{j,k}| < 2\kappa\lambda_n\}} \leq C\lambda_n^{2-p} |\beta_{j,k}|^p$, $B_{p,r}^s(M) \subseteq B_{2,\infty}^{s+1/2-1/p}(M)$ and $(2s+1)(2-p)/2 + (s+1/2-1/p)p = 2s$, we have

$$\begin{aligned} S_{4,2} &\leq C\lambda_n^{2-p} \sum_{j=j_2+1}^{j_1} \sum_{k=0}^{2^j-1} |\beta_{j,k}|^p = C \left(\frac{(\ln n)^{1+1/\theta}}{n} \right)^{(2-p)/2} \sum_{j=j_2+1}^{j_1} \sum_{k=0}^{2^j-1} |\beta_{j,k}|^p \\ &\leq C \left(\frac{(\ln n)^{1+1/\theta}}{n} \right)^{(2-p)/2} \sum_{j=j_2+1}^{\infty} 2^{-j(s+1/2-1/p)p} \\ &\leq C \left(\frac{(\ln n)^{1+1/\theta}}{n} \right)^{(2-p)/2} 2^{-j_2(s+1/2-1/p)p} \leq C \left(\frac{(\ln n)^{1+1/\theta}}{n} \right)^{2s/(2s+1)}. \end{aligned} \quad (6.81)$$

So, for $r \geq 1$, $\{p \geq 2 \text{ and } s > 0\}$ or $\{p \in [1, 2) \text{ and } s > 1/p\}$, we have

$$S_4 \leq C \left(\frac{(\ln n)^{1+1/\theta}}{n} \right)^{2s/(2s+1)}. \quad (6.82)$$

It follows from (6.61), (6.66), (6.75), and (6.82) that

$$S \leq C \left(\frac{(\ln n)^{1+1/\theta}}{n} \right)^{2s/(2s+1)}. \quad (6.83)$$

Combining (6.55), (6.57), (6.60), and (6.83), we have, for $r \geq 1$, $\{p \geq 2 \text{ and } s > 0\}$ or $\{p \in [1, 2) \text{ and } s > 1/p\}$,

$$\mathbb{E} \left(\int_0^1 (\hat{f}^H(x) - f(x))^2 dx \right) \leq C \left(\frac{(\ln n)^{1+1/\theta}}{n} \right)^{2s/(2s+1)}. \quad (6.84)$$

The proof of Theorem 5.2 is complete. \square

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References

- [1] S. T. Buckland, D. R. Anderson, K. P. Burnham, and J. L. Laake, *Distance Sampling: Estimating Abundance of Biological Populations*, Chapman & Hall, London, UK, 1993.
- [2] D. Cox, "Some sampling problems in technology," in *New Developments in Survey Sampling*, N. L. Johnson and H. Smith Jr., Eds., pp. 506–527, John Wiley & Sons, New York, NY, USA, 1969.
- [3] J. Heckman, "Selection bias and self-selection," in *The New Palgrave : A Dictionary of Economics*, pp. 287–296, MacMillan Press, New York, NY, USA, 1985.
- [4] G. P. Patil and C. R. Rao, "The weighted distributions: a survey of their applications," in *Applications of Statistics*, P. R. Krishnaiah, Ed., pp. 383–405, North-Holland, Amsterdam, The Netherlands, 1977.
- [5] H. El Barmi and J. S. Simonoff, "Transformation-based density estimation for weighted distributions," *Journal of Nonparametric Statistics*, vol. 12, no. 6, pp. 861–878, 2000.
- [6] S. Efromovich, "Density estimation for biased data," *The Annals of Statistics*, vol. 32, no. 3, pp. 1137–1161, 2004.
- [7] E. Brunel, F. Comte, and A. Guilloux, "Nonparametric density estimation in presence of bias and censoring," *Test*, vol. 18, no. 1, pp. 166–194, 2009.
- [8] C. Chesneau, "Wavelet block thresholding for density estimation in the presence of bias," *Journal of the Korean Statistical Society*, vol. 39, no. 1, pp. 43–53, 2010.
- [9] P. Ramirez and B. Vidakovic, "Wavelet density estimation for stratified size-biased sample," *Journal of Statistical Planning and Inference*, vol. 140, no. 22, pp. 419–432, 2010.
- [10] H. Doosti and I. Dewan, "Wavelet linear density estimation for associated stratified size-biased sample," *Statistics & Mathematics Unit*. In press.
- [11] P. Doukhan, *Mixing. Properties and Examples*, vol. 85 of *Lecture Notes in Statistics*, Springer, New York, NY, USA, 1994.
- [12] M. Carrasco and X. Chen, "Mixing and moment properties of various GARCH and stochastic volatility models," *Econometric Theory*, vol. 18, no. 1, pp. 17–39, 2002.
- [13] A. Antoniadis, "Wavelets in statistics: a review (with discussion)," *Journal of the Italian Statistical Society, Series B*, vol. 6, pp. 97–144, 1997.
- [14] W. Härdle, G. Kerkycharian, D. Picard, and A. Tsybakov, *Wavelets, Approximation, and Statistical Applications*, vol. 129 of *Lecture Notes in Statistics*, Springer, New York, NY, USA, 1998.
- [15] D. L. Donoho, I. M. Johnstone, G. Kerkycharian, and D. Picard, "Density estimation by wavelet thresholding," *The Annals of Statistics*, vol. 24, no. 2, pp. 508–539, 1996.
- [16] C. S. Withers, "Conditions for linear processes to be strong-mixing," *Zeitschrift für Wahrscheinlichkeitstheorie und Verwandte Gebiete*, vol. 57, no. 4, pp. 477–480, 1981.
- [17] D. S. Modha and E. Masry, "Minimum complexity regression estimation with weakly dependent observations," *IEEE Transactions on Information Theory*, vol. 42, no. 6, part 2, pp. 2133–2145, 1996.
- [18] A. Cohen, I. Daubechies, and P. Vial, "Wavelets on the interval and fast wavelet transforms," *Applied and Computational Harmonic Analysis*, vol. 1, no. 1, pp. 54–81, 1993.
- [19] Y. Meyer, *Wavelets and Operators*, vol. 37 of *Cambridge Studies in Advanced Mathematics*, Cambridge University Press, Cambridge, UK, 1992.
- [20] Y.P. Chaubey, C. Chesneau, and H. Doosti, "On linear wavelet density estimation: some recent developments," *Journal of the Indian Society of Agricultural Statistics*. In press.
- [21] F. Leblanc, "Wavelet linear density estimator for a discrete-time stochastic process: L_p -losses," *Statistics & Probability Letters*, vol. 27, no. 1, pp. 71–84, 1996.
- [22] E. Masry, "Probability density estimation from dependent observations using wavelets orthonormal bases," *Statistics & Probability Letters*, vol. 21, no. 3, pp. 181–194, 1994.
- [23] B. Delyon and A. Juditsky, "On minimax wavelet estimators," *Applied and Computational Harmonic Analysis*, vol. 3, no. 3, pp. 215–228, 1996.
- [24] J. A. Davydov, "The invariance principle for stationary processes," *Theory of Probability and Its Applications*, vol. 15, pp. 498–509, 1970.
- [25] E. Rio, "The functional law of the iterated logarithm for stationary strongly mixing sequences," *The Annals of Probability*, vol. 23, no. 3, pp. 1188–1203, 1995.
- [26] E. Liebscher, "Strong convergence of sums of α -mixing random variables with applications to density estimation," *Stochastic Processes and their Applications*, vol. 65, no. 1, pp. 69–80, 1996.



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