# LONG-RUN AVAILABILITY OF A PRIORITY SYSTEM: A NUMERICAL APPROACH 

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We consider a two-unit cold standby system attended by two repairmen and subjected to a priority rule. In order to describe the random behavior of the twin system, we employ a stochastic process endowed with state probability functions satisfying coupled Hokstadtype differential equations. An explicit evaluation of the exact solution is in general quite intricate. Therefore, we propose a numerical solution of the equations. Finally, particular but important repair time distributions are involved to analyze the long-run availability of the T-system. Numerical results are illustrated by adequate computer-plotted graphs.

## 1. Introduction

Standby provides a powerful tool to enhance the reliability, availability, quality, and safety of operational plants, see, for example, [3, 7, 14]. However, in practice, standby systems are often subjected to an appropriate priority rule. For instance, the external power supply station of a technical plant has usually overall priority in operation with regard to an internal (local) power generator kept in cold or warm standby [3]. The local generator is only deployed if the external power station is down.

Cold or warm standby systems subjected to a priority rule and attended by a repair facility have received considerable attention in the current literature, see, for example, $[1,2,4,5,8,9,10,11,12,13,16,17,18,19,20,21]$.

As a variant, we consider a twin system composed of a priority unit (the p-unit) and a nonpriority unit (the $\mathbf{n}$-unit) kept in cold standby. The $\mathbf{p}$-unit has overall (break-in) priority in operation with regard to the $\mathbf{n}$-unit, that is, the $\mathbf{n}$-unit is only used when the p-unit is down. In order to avoid undesirable delays in repairing failed units, we suppose that the entire system (henceforth called the $\mathbf{T}$-system) is attended by two different repairmen. The $\mathbf{T}$-system satisfies the usual conditions, that is, independent identically distributed random variables, instantaneous and perfect switch [3], and perfect repair [6]. Each repairman has his own particular task. Repairman $\mathcal{N}$ is skilled in repairing the $\mathbf{n}$-unit, whereas repairman $\mathscr{P}$ is an expert in repairing the $\mathbf{p}$-unit. Both repairmen are jointly busy, if and only if, both units ( $\mathbf{p}$-unit and $\mathbf{n}$-unit) are down. In any other case, at least one repairman is idle.

In order to describe the random behavior of the T-system, we employ a stochastic process endowed with transition probability functions satisfying steady-state Hokstadtype differential equations. Unfortunately, the exact solution procedure is quite intricate (see, [21, page 359] and Remark 4.1). Therefore, we propose a numerical solution of the equations.

Finally, current repair time distributions (such as the Weibull-Gnedenko distribution) are involved to compute the long-run availability of the $\mathbf{T}$-system. The results are illustrated by adequate computer-plotted graphs.

## 2. Formulation

Consider a T-system satisfying the usual conditions. The p-unit has a constant failure rate [15] $\lambda>0$ and a general repair time distribution $R(\cdot), R(0)=0$, with mean $\rho$. The operative $\mathbf{n}$-unit has a constant failure rate $\lambda_{s}>0$, but a zero failure rate in standby (the so-called cold standby state) and a general repair time distribution $R_{S}(\cdot), R_{S}(0)=0$, with mean $\rho_{s}$. In order to describe the random behavior of the $\mathbf{T}$-system, we introduce a stochastic process $\left\{N_{t}, t \geq 0\right\}$ with arbitrary discrete state space $\{A, B, C, D\} \subset[0, \infty)$, characterized by the following mutually exclusive events:
(i) $\left\{N_{t}=A\right\}$ : "the $\mathbf{p}$-unit is operative and the $\mathbf{n}$-unit is in cold standby at time $t$,"
(ii) $\left\{N_{t}=B\right\}$ : "the $\mathbf{n}$-unit is operative and the $\mathbf{p}$-unit is under repair at time $t$,"
(iii) $\left\{N_{t}=C\right\}$ : "the $\mathbf{p}$-unit is operative and the $\mathbf{n}$-unit is under repair at time $t$,"
(iv) $\left\{N_{t}=D\right\}$ : "both units are simultaneously down at time $t$."

State $D$ is called the system-down state.
Figures 2.1, 2.2, 2.3, and 2.4 display a functional block diagram of the $\mathbf{T}$-system operating in states $A, B, C$, and $D$.

Observe that the process $\left\{N_{t}, t \geq 0\right\}$ is non-Markovian. A Markov characterization of the process is piecewise and conditionally defined by:
(i) $\left\{N_{t}\right\}$, if $N_{t}=A$ (i.e., if the event $\left\{N_{t}=A\right\}$ occurs),
(ii) $\left\{\left(N_{t}, X_{t}\right)\right\}$, if $N_{t}=B$, where $X_{t}$ denotes the remaining repair time of the $\mathbf{p}$-unit under progressive repair at time $t$,
(iii) $\left\{\left(N_{t}, Y_{t}\right)\right\}$, if $N_{t}=C$, where $Y_{t}$ denotes the remaining repair time of the $\mathbf{n}$-unit under progressive repair at time $t$,
(iv) $\left\{\left(N_{t}, X_{t}, Y_{t}\right)\right\}$, if $N_{t}=D$.

The state space of the underlying piecewise linear (vector) Markov process is given by

$$
\begin{equation*}
A \cup\{(B, x) ; x \geq 0\} \cup\{(C, y) ; y \geq 0\} \cup\{(D, x, y) ; x \geq 0 ; y \geq 0\} . \tag{2.1}
\end{equation*}
$$

Next, we consider the T-system in stationary state (the so-called ergodic state) with invariant measure $\left\{p_{K} ; K=A, B, C, D\right\}, \sum_{K} p_{K}=1$, where

$$
\begin{equation*}
p_{K}:=\lim _{t \rightarrow \infty} \mathbf{P}\left\{N_{t}=K \mid N_{0}=A\right\} . \tag{2.2}
\end{equation*}
$$



Figure 2.1. Functional block diagram of the $\mathbf{T}$-system operating in state $A$.


Figure 2.2. Functional block diagram of the $\mathbf{T}$-system operating in state $B$.


Figure 2.3. Functional block diagram of the $\mathbf{T}$-system operating in state $C$.


Figure 2.4. Functional block diagram of the $\mathbf{T}$-system in state $D$.

It can be demonstrated that the invariant measure exists for arbitrary $R$ and $R_{S}$ with finite mean. However, in order to keep the analysis as simple as possible, we henceforth assume that $R$ and $R_{S}$ have bounded densities on $[0, \infty)$, denoted by $r$ and $r_{s}$. Finally, we introduce the measures

$$
\begin{array}{r}
p_{B}(x) d x:=\lim _{t \rightarrow \infty} \mathbf{P}\left\{N_{t}=B, X_{t} \in(x, x+d x] \mid N_{0}=A\right\}, \\
p_{C}(y) d y:=\lim _{t \rightarrow \infty} \mathbf{P}\left\{N_{t}=C, Y_{t} \in(y, y+d y] \mid N_{0}=A\right\},  \tag{2.3}\\
p_{D}(x, y) d x d y:=\lim _{t \rightarrow \infty} \mathbf{P}\left\{N_{t}=D, X_{t} \in(x, x+d x], Y_{t} \in(y, y+d y] \mid N_{0}=A\right\} .
\end{array}
$$

Note that, for instance, $p_{D}=\int_{0}^{\infty} \int_{0}^{\infty} p_{D}(x, y) d x d y$.

## 3. Long-run availability

We recall that the $\mathbf{T}$-system is only available (functioning) in states $A, B$, and $C$. Therefore, the long-run availability of the operational plant, denoted by $\mathscr{A}$, is given by $\mathscr{A}=1-p_{D}$. Invoking the substitutions $p_{B}(\cdot)=p_{A} \varphi_{B}(\cdot), p_{C}(\cdot)=p_{A} \varphi_{C}(\cdot), p_{D}(\cdot, \cdot)=p_{A} \varphi_{D}(\cdot, \cdot)$ and the law $\sum_{K} p_{K}=1$ entails that $p_{A}=1 /\left(1+\Phi_{B}+\Phi_{C}+\Phi_{D}\right)$, where $\Phi_{B}:=\int_{0}^{\infty} \varphi_{B}(x) d x$, $\Phi_{C}:=\int_{0}^{\infty} \varphi_{C}(y) d y$ and $\Phi_{D}:=\int_{0}^{\infty} \int_{0}^{\infty} \varphi_{D}(x, y) d x d y$. Hence,

$$
\begin{equation*}
\mathscr{A}=\frac{1+\Phi_{B}+\Phi_{C}}{1+\Phi_{B}+\Phi_{C}+\Phi_{D}} . \tag{3.1}
\end{equation*}
$$

## 4. Differential equations

In order to determine the $\varphi$-functions, we first construct a system of coupled steady statetype differential equations based on a time-independent version of Hokstad's supplementary variable technique (see, e.g., [22, page 526] for further details). For $x>0, y>0$, we obtain

$$
\begin{gather*}
\lambda=\varphi_{B}(0)+\varphi_{C}(0)  \tag{4.1}\\
\left(\lambda_{s}-\frac{d}{d x}\right) \varphi_{B}(x)=\varphi_{D}(x, 0)+\lambda r(x),  \tag{4.2}\\
\left(\lambda-\frac{d}{d y}\right) \varphi_{C}(y)=\varphi_{D}(0, y),  \tag{4.3}\\
\left(-\frac{\partial}{\partial x}-\frac{\partial}{\partial y}\right) \varphi_{D}(x, y)=\lambda_{s} \varphi_{B}(x) r_{s}(y)+\lambda \varphi_{C}(y) r(x) . \tag{4.4}
\end{gather*}
$$

Remark 4.1. A particular but important family $\mathscr{F}$ of current repair time distributions with nonrational characteristic functions, such as the Weibull-Gnedenko and Lognormal distributions, are fairly suitable to model repair times. Unfortunately, if both $R$ and $R_{S}$ belong to $\mathscr{F}$, an explicit evaluation of the exact solution of (4.1), (4.2), (4.3), and (4.4) in terms of finite linear combinations of known algebraic and/or transcendental functions is as good as excluded (see [21, page 361] for further details). Therefore, we propose a numerical solution of the equations.

## 5. Numerical scheme

In order to construct an appropriate numerical procedure, we first remark that the $\varphi$ functions are vanishing at infinity irrespective of the asymptotic behavior of the repair time density functions! Therefore, a numerical procedure to solve the equations in the region $(0, \infty) \times(0, \infty)$ may be converted into a numerical solution procedure in the truncated region $(0, L) \times(0, L)$, for some $L>0$, with prescribed boundary conditions $\varphi_{B}(L)=$ $\varphi_{C}(L)=\varphi_{D}(L, \cdot)=\varphi_{D}(\cdot, L)=0$. Let $\varphi_{B, i}:=\varphi_{B}\left(x_{i}\right), \varphi_{C, j}:=\varphi_{C}\left(y_{j}\right), \varphi_{D, i, j}:=\varphi_{D}\left(x_{i}, y_{i}\right)$, where $x_{i}:=i \Delta, y_{j}:=j \Delta, i=0, \ldots, N+1 ; j=0, \ldots, N+1 ; \Delta:=L / N$. We propose the following numerical scheme. Let $k$ be the iteration number. Given $\varphi_{D, i, N+1}^{k+1}=0, \varphi_{D, N+1, j}^{k+1}=0$, $\varphi_{B, N+1}^{k+1}=0, \varphi_{C, N+1}^{k+1}=0$, and the values of $\varphi_{B, i}^{k}$ and $\varphi_{C, j}^{k}$, we compute $\varphi_{D, i, j}^{k+1}$ by means of the two-point first-order approximation of (4.4), namely,

$$
\begin{equation*}
\varphi_{D, i, j}^{k+1}=\frac{1}{2}\left(\varphi_{D, i, j+1}^{k+1}+\varphi_{D, i+1, j}^{k+1}\right)+\frac{\Delta}{2}\left(\lambda_{s} \varphi_{B, i}^{k} r_{s, j}+\lambda \varphi_{C, j}^{k} r_{i}\right), \tag{5.1}
\end{equation*}
$$

$i=N, N-1, \ldots, 0$ and $j=N, N-1, \ldots, 0$.
Next, we calculate $\varphi_{B, i}^{k+1}$ and $\varphi_{C, j}^{k+1}$ by means of the first-order approximations of (4.2) and (4.3) given by

$$
\begin{gather*}
\varphi_{B, i}^{k+1}=\frac{1}{\gamma_{B}}\left(\frac{\varphi_{B, i+1}^{k+1}}{\Delta}+\varphi_{D, i, 0}^{k+1}+\lambda r_{i}\right),  \tag{5.2}\\
\varphi_{C, j}^{k+1}=\frac{1}{\gamma_{C}}\left(\frac{\varphi_{C, j+1}^{k+1}}{\Delta}+\varphi_{D, 0, j}^{k+1}\right)
\end{gather*}
$$

where $\gamma_{B}:=\lambda_{s}+1 / \Delta$ and $\gamma_{C}:=\lambda+1 / \Delta$. Finally, in order to satisfy (4.1) we use the normalizing procedure

$$
\begin{align*}
& \varphi_{C, j}^{k+1, \text { new }}=\lambda \frac{\varphi_{C, j}^{k+1}}{\varphi_{C, 0}^{k+1}+\varphi_{B, 0}^{k+1}},  \tag{5.3}\\
& \varphi_{B, i}^{k+1, \text { new }}=\lambda \frac{\varphi_{B, i}^{k+1}}{\varphi_{C, 0}^{k+1}+\varphi_{B, 0}^{k+1}} .
\end{align*}
$$

Remarks 5.1. Let $\varphi_{\Delta}$ denote a numerical solution obtained with the space-step $\Delta$. The relevant numerical error is then evaluated on a nested grid by $\varepsilon:=\left|\varphi_{\Delta}-\varphi_{\Delta / 2}\right|$. However, such an estimate is only accurate if $L$ is large enough. Roughly speaking, if $\max \left(r(x), r_{s}(x)\right)$ at $x=L$ is small, then (most likely) this particular $L$ is appropriate. However, such a "brutal force" approach may require a large number of grid points and is therefore rarely applicable. We illustrate the phenomenon by comparing the exact and the numerical solution in the most simple case, that is, let $R(x)=1-e^{-x}, R_{S}(y)=1-e^{-y}$. Then, $\varphi_{D}(x, y)=$ $l_{D} e^{-(x+y)}, \varphi_{C}(y)=l_{C} e^{-y}, \varphi_{B}(x)=l_{B} e^{-x}$, where $l_{D}:=\lambda_{s}(\lambda+1) \lambda /\left(\lambda_{s}+\lambda+2\right), l_{C}:=\lambda_{s} \lambda /\left(\lambda_{s}+\right.$ $\lambda+2), l_{B}:=\lambda(\lambda+2) /\left(\lambda_{s}+\lambda+2\right)$.

Figure 5.1 shows the numerical error

$$
\begin{equation*}
\varepsilon_{M}:=\max \left\{\left|\varphi_{D}^{\text {exact }}-\varphi_{D}\right|,\left|\varphi_{C}^{\text {exact }}-\varphi_{C}\right|,\left|\varphi_{B}^{\text {exact }}-\varphi_{B}\right|\right\} \tag{5.4}
\end{equation*}
$$

versus the grid size for various $L$.


Figure 5.1. The horizontal axis denotes the logarithm of the numerical error, the vertical axis denotes the number of the grid points, (1) $L=0.4$; (2) $L=1.0$; (3) $L=1.5$; (4) $L=2$; (5) $L=4$; (6) $L=6$; (7) $L=10$; (8) $L=50$.


Figure 5.2. Spatial distribution of $\varepsilon_{D}$, (1) $L=1.5$; (2) $L=3$; (3) $L=6$.

Observe that, if $L$ is not large enough, $\varepsilon_{M}$ does not decrease as $\Delta$ decreases (see Figure 5.1). On the other hand, too large $L$ (consequently, too large $\Delta$ ) lead to large numerical errors. For instance, the error with $L=30$ is larger than $2.5 \cdot 10^{-2}$ for any $N \in[20,100]$, whereas the error with $L=4$ is less than $2.5 \cdot 10^{-2}$. There could be multiple options too. For instance, an error less than $2.5 \cdot 10^{-2}$ is achieved either with $L=4$, $N=15$, or $L=6, N=22$, or $L=10, N=38$.

Figure 5.2 shows a two-dimensional spatial distribution of the error $\varepsilon_{D}:=\left|\varphi_{D}^{\text {exact }}-\varphi_{D}\right|$ for various $L$. Clearly, $\varepsilon_{D}$ could be increasing near the origin as $L$ increases. However, the error decreases for large $x$ and $y$.

## 6. Trial-and-error procedure

The complicated behavior of the numerical error requires an adaptive choice of $\Delta$ and $L$. Therefore, we introduce the subordinate errors $\varepsilon_{1}:=\left|\varphi_{\Delta, L}-\varphi_{\Delta, L / 2}\right|$ and $\varepsilon_{2}:=\mid \varphi_{\Delta, L}-$ $\varphi_{\Delta / 2, L}$ l, where $\varepsilon_{1}$ characterizes the numerical error caused by truncation of the infinite region and $\varepsilon_{2}$ the numerical error related to the first-order approximants. In order to find the optimal pair $(L, \Delta)$, we first specify the required accuracy $\varepsilon$. Next, we propose the following trial-and-error procedure: we vary $L$ until $\varepsilon_{1}<\varepsilon$ and then $\Delta$ until $\varepsilon_{2}<\varepsilon$. Finally, we introduce the following.

## 7. Application. The Weibull-Gnedenko distribution

We consider the particular but important case of Weibull-Gnedenko repair time distributions, that is, let $R(x)=1-e^{-x^{\beta_{1}}}, R_{S}(y)=1-e^{-y^{\beta_{2}}}$. Obviously, the optimal pair $(L, \Delta)$ depends on $\lambda, \lambda_{s}, \beta_{1}$, and $\beta_{2}$. We demonstrate the trial-and-error procedure applied to the particular case $\lambda=1, \lambda_{s}=0.1, \beta_{1}=2, \beta_{2}=3$. However, no restrictions are imposed on the analysis of $\mathscr{A}$ for arbitrary values of $\lambda, \lambda_{s}, \beta_{1}$ and $\beta_{2}$. Let $\varepsilon=10^{-3}$.

First, we vary $L$, as shown in Table 7.1, until $\varepsilon_{1}<\varepsilon$. Next, we vary $\Delta$, as shown in Table 7.2, until $\varepsilon_{2}<\varepsilon$. A spatial distribution of $\varepsilon_{1}$ and $\varepsilon_{2}$ is depicted in Figures 7.1 and 7.2.

Table 7.1. The $L$ trials.

| $L$ | $N$ | $\Delta$ | $\varepsilon_{1}$ |
| :---: | :---: | :---: | :---: |
| 3 | 40 | $3 / 40$ | $1.9 \cdot 10^{-2}$ |
| 6 | 80 | $3 / 40$ | $7.4 \cdot 10^{-4}$ |

Table 7.2. The $\Delta$ trials.

| $L$ | $N$ | $\Delta$ | $\varepsilon_{2}$ |
| :---: | :---: | :---: | :---: |
| 6 | 80 | $3 / 40$ | $6.8 \cdot 10^{-3}$ |
| 6 | 160 | $3 / 80$ | $3.3 \cdot 10^{-3}$ |
| 6 | 320 | $3 / 160$ | $9.2 \cdot 10^{-4}$ |



Figure 7.1. Spatial distribution of $\varepsilon_{1}$, (1) $L=3$; (2) $L=6$.


Figure 7.2. Spatial distribution of $\varepsilon_{2}$ for $N=320$.


Figure 7.3. Numerically generated: (1) $p_{B}(x) / 1.5$, (2) $p_{C}(x)$, (3) $p_{D}(x, y), \lambda_{s}=0.3 \lambda=1.0$. Note that $p_{B}$ is divided by 1.5 due to scaling.


Figure 7.4. Numerically generated: (1) $p_{B}(x) / 1.5$, (2) $p_{C}(x)$, (3) $p_{D}(x, y), \lambda_{s}=0.7 \lambda=1.0$. Note that $p_{B}$ is divided by 1.5 due to scaling.

Figure 7.3 displays $p_{B}(\cdot), p_{C}(\cdot)$, and $p_{D}(\cdot, \cdot)$ for $\lambda=1, \lambda_{s}=0.3$ and Figure 7.4 for $\lambda=$ $1, \lambda_{s}=0.5$. Figure 7.5 shows $p_{D}(x, y)$ for various $\lambda_{s}$. Let $\mathcal{A}_{\beta_{1}, \beta_{2}}\left(\lambda, \lambda_{s}\right)$ denote the long-run availability as a function of $\lambda$ and $\lambda_{s}$.


Figure 7.5. Numerically generated $p_{D}(x, y), \lambda=1.0$, (1) $\lambda_{s}=0.1$, (2) $\lambda_{s}=0.3$, (3) $\lambda_{s}=0.7$.


Figure 7.6. Numerically generated long-run availability.

Figure 7.6 shows that the long-run availability exhibits a nonlinear behavior for sufficiently large $\lambda$ and $\lambda_{s}$ (see also Table 7.3). Finally, Figure 7.7 displays the deviations $d_{1}:=\left|\mathscr{A}_{2,2}-\mathscr{A}_{2,4}\right|, d_{2}:=\left|\mathscr{A}_{2,2}-\mathscr{A}_{4,2}\right|, d_{3}:=\left|\mathscr{A}_{2,2}-\mathscr{A}_{4,4}\right|$. The plot reveals that $\mathscr{A}$ is fairly insensitive for $\beta$-variations.

Table 7.3. Long-run availability $\mathscr{A}_{2,2}\left(\lambda, \lambda_{s}\right)$.

| $\lambda / \lambda_{s}$ | 0.1 | 0.2 | 0.3 | 0.4 | 0.5 | 0.6 | 0.7 | 0.8 | 0.9 | 1.0 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0.1 | 0.9950 | 0.9895 | 0.9837 | 0.9775 | 0.9709 | 0.9640 | 0.9566 | 0.9489 | 0.9408 | 0.9324 |
| 0.2 | 0.9908 | 0.9811 | 0.9706 | 0.9596 | 0.9480 | 0.9358 | 0.9231 | 0.9098 | 0.8962 | 0.8821 |
| 0.3 | 0.9874 | 0.9740 | 0.9599 | 0.9450 | 0.9295 | 0.9133 | 0.8965 | 0.8792 | 0.8615 | 0.8434 |
| 0.4 | 0.9845 | 0.9681 | 0.9509 | 0.9330 | 0.9143 | 0.8949 | 0.8750 | 0.8546 | 0.8338 | 0.8128 |
| 0.5 | 0.9820 | 0.9631 | 0.9434 | 0.9229 | 0.9016 | 0.8797 | 0.8573 | 0.8345 | 0.8114 | 0.7882 |
| 0.6 | 0.9798 | 0.9588 | 0.9369 | 0.9143 | 0.8909 | 0.8670 | 0.8426 | 0.8179 | 0.7930 | 0.7680 |
| 0.7 | 0.9779 | 0.9550 | 0.9313 | 0.9069 | 0.8818 | 0.8562 | 0.8301 | 0.8039 | 0.7775 | 0.7512 |
| 0.8 | 0.9762 | 0.9517 | 0.9265 | 0.9005 | 0.8740 | 0.8469 | 0.8195 | 0.7920 | 0.7644 | 0.7370 |
| 0.9 | 0.9748 | 0.9488 | 0.9222 | 0.8949 | 0.8671 | 0.8389 | 0.8103 | 0.7817 | 0.7532 | 0.7249 |
| 1.0 | 0.9734 | 0.9462 | 0.9184 | 0.8900 | 0.8611 | 0.8318 | 0.8024 | 0.7729 | 0.7435 | 0.7145 |



Figure 7.7. Spatial deviations $d_{i}, i=1,2,3$.

## 8. Conclusion

An effective statistical analysis of the T-system requires the solution of coupled Hokstadtype differential equations. Our numerical solution procedure, endowed with a simple and robust algorithm, allows to compute and to analyze the long-run availability for a general class of current repair time distributions with tangible engineering applications.

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