

Research Article

Identification of Continuous-Time Systems with Multiple Unknown Time Delays from Sampled Data

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The purpose of this work is to identify a multiple-input single-output (MISO) continuous time (CT) system with various unknown time delays using sampled data. A novel technique for simultaneously estimating system parameters and various time delays may be developed using a gradient algorithm. Indeed, we propose a new formulation of the identification problem which permits defining the generalized vector containing the time delays and the dynamic parameters. The gradient optimization algorithm with exact derivatives of the objective function with respect to the adjustable parameters has been suggested as an alternative for solving this optimization problem. The identification method can be easily implemented by a filtering procedure to generate the input/output time derivatives. In this sense, the proposed method, which is based on a state variable filter, intrinsically allows data filtering and thus simplifies the use of the method in a practical case. An analysis is then presented which proves the convergence of the algorithm, and robustness issues are discussed as well. Finally, the suggested technique is evaluated on simulated instances to ensure its validity and performance.

1. Introduction

Time delays are unavoidable in the operation of many industrial process models [1, 2]. Indeed, ignoring its presence might result in complicated behavior (oscillations, disregarded data, instability of the closed loop, and so on) [3, 4]. As a result, the identification of time delay models is indeed a problem of considerable importance that deserves special attention in virtually all disciplines of science and engineering. Two major issues must be addressed: the first step is to identify the dynamics parameters, and the second step is to identify the temporal delays. Indeed, many techniques take into account both the identification of CT systems [5, 6] and the identification of DT systems [7, 8].

In fact, the identification of physical systems was usually conducted in discrete time. This was mainly due to the “go completely discrete-time” trend that was spurred by parallel developments in digital computers.

Indeed, historically discrete-time models have received more attention than continuous-time models. Therefore, many researchers in the system identification field have concentrated their efforts on developing DT identification methods [7, 9–11]. In fact, for some reasons, discrete-time approaches have been recognized. First, the stochastic theory of DT approaches is more easily compared with CT approaches which need to calculate time derivatives of input/output data. Second, estimating the physical parameters of the continuous-time physical model is not important if the model is to design controllers.

However, continuous-time approaches are preferable under certain circumstances. First, when we have to estimate physical signal process inputs such as pressure and temperature, we should use continuous-time approaches because the physical signal is always a continuous-time signal rather than a discrete-time signal from a zero-order holder. Second, when we need to estimate physical parameters for

the purposes of analysis and system design, continuous-time approaches are recommended.

Indeed, many continuous-time system identification methods dealing with CT systems with time delays have been developed [5, 12–15].

In reality, the challenge of identifying SISO time-delay systems is addressed by the majority of these approaches. Nonetheless, one of the most challenging issues is identifying MISO time-delay systems that is a field of study where few works have been published. Nevertheless, for multi-input single-output (MISO) systems with multiple unknown time delays, the problem is much more difficult. Such systems are often more challenging to model. In particular, systems with several inputs can be difficult. A basic reason for the difficulties is that the couplings between several inputs and an output lead to more complex models. A second reason for studying the MISO systems with multiple unknown time delays is the presence of multiple unknown delays that is generally considered as a great challenge where few works have been developed. This decision is also supported by the fact that these models are often employed in most process sectors, including vehicle traffic [16], thermal agitation [17], network management [18], and variable pitch milling [19]. In summary, it is worth noting that the identification issue of MISO CT systems with multiple unknown time delays is, therefore, both of theoretical and practical importance.

However, several studies have previously been conducted on the issue. In [20], a technique based on a modulating function method was proposed. Paplinski employs an ant colony optimization technique in [21]. This method, in reality, has a considerable execution time. The evolutionary algorithm [22], which estimates the optimal parameters, is another approach for estimating systems with many unknown time delays. This scheme is said to be an offline technique that takes too long to compute at times. Another technique is described in [23], which permits the detection of MISO CT systems with unknown time delays. There are two versions: one is based on least squares, while the other is not. The second version, on the other hand, is based on the instrumental variable estimator, which is utilized to provide unbiased estimates. GSEPNLS and GSEPNIV are the acronyms for global separable nonlinear least-squares and global separable nonlinear instrumental variable, respectively. In [24], the problem of MISO CT model identification with multiple unknown time delays from sampled data has been treated. The proposed method estimated the plant and the time delays in a separable way. More precisely, the plant has been estimated by the standard recursive least-square algorithm while the time delays have been explicitly estimated by the Gauss–Newton algorithm. In [25], we have estimated the parameters applying a Levenberg–Marquardt algorithm while the time delays have been estimated by the Gauss–Newton algorithm. In [26, 27], the separable estimation of continuous-time hybrid “Box–Jenkins” systems with multiple unknown time delays has been considered. It can be seen that the developed methods [24–27] require a separable identification strategy of parameters and time delays. This strategy

is considered to be very slow and sometimes requires a large computational effort. However, successful implementations of identification methods in terms of computation time require simultaneous identification of the unknown time delays and the linear parameters of the system. An interesting issue, in this regard, includes simultaneously identifying the model parameters and time delays.

The majority of the developed approaches require a separable identification strategy of parameters and time delays. This technique is focused on using a specific algorithm to identify temporal delays. The linear parameters are then calculated using a different approach. In truth, this method is quite sluggish and frequently necessitates a significant amount of computing work. However, in order to reduce the computation time, successful identification methods require simultaneous identification of the unknown time delays and the linear parameters of the system. An interesting issue, in this regard, includes simultaneously identifying the model parameters and time delays.

In order to overcome the underlying problems, this study addresses the topic of simultaneous identification of the MISO CT system with various unknown time delays to address the underlying issue. Indeed, we provide a novel formulation of the issue that permits the temporal delays and dynamic parameters to be defined in the same estimated vector. Then, based on this formulation, we offer a novel gradient algorithm-based approach. The developed algorithm’s convergence criteria are also developed.

The remainder of the present paper is laid out as follows. The problem in this study is introduced in Section 2. The developed gradient algorithm is shown in Section 3. In Section 4, the suggested algorithm’s convergence is tested. We provide two instances in Section 5 to demonstrate the efficacy of the proposed approach. Finally, in Section 6, final observations are offered to bring the study to a close.

2. Problem Statement

In this research work, the following process is considered:

$$x_u(t) = \frac{G_1(s)}{F(s)}e^{-d_1s}u_1(t) + \dots + \frac{G_r(s)}{F(s)}e^{-d_rs}u_r(t). \quad (1)$$

e^{-d_js} is the time delay operator defined as follows:

$$e^{-d_js}u_j(t) = u_j(t - d_j). \quad (2)$$

So,

$$x_u(t) = \frac{G_1(s)}{F(s)}u_1(t - d_1) + \dots + \frac{G_r(s)}{F(s)}u_r(t - d_r). \quad (3)$$

The polynomials $G_j(s)$ and $F(s)$ are defined, respectively, by

$$\begin{aligned} G_j(s) &= g_{j0} + g_{j1}s + \dots + g_{jq_j}s^{q_j}, \\ F(s) &= 1 + f_1s + \dots + f_ps^p. \end{aligned} \quad (4)$$

The considered system can be described as

$$\sum_{i=0}^p f_i s^{p-i} x_u(t) = \sum_{j=1}^r \sum_{i=1}^{q_j} g_{ji} s^{q_j-i} u_j(t-d_j). \quad (5)$$

To sum up, the considered system is depicted in Figure 1.

3. Suggested Technique

Let us introduce the low-pass filter $K(s)$ [28] where α and p are the cutoff frequency and the order.

$$K(s) = \frac{1}{(\alpha s + 1)^p}. \quad (6)$$

Using the bilinear transformation described in [29], we obtain

$$\bar{x}_0(k) + \sum_{i=1}^p f_i \xi_{i\bar{x}}(k) = \sum_{j=1}^r \sum_{i=1}^{q_j} g_{ji} \xi_{(p-q_j+i)\bar{u}_j}(k - \tilde{d}_j) + v(k). \quad (7)$$

Such that

$$v(k) = \sum_{i=0}^p f_i \xi_{i\bar{v}}(k), \quad (8)$$

$$\xi_{i\bar{v}}(k) = \bar{v}(k) K_i(q^{-1}),$$

and

$$\xi_{i\bar{u}_j}(k) = \bar{u}_j(k) K_i(q^{-1}),$$

$$\xi_{i\bar{x}}(k) = \bar{y}(k) K_i(q^{-1}), \quad (9)$$

$$K_i(q^{-1}) = \frac{(T/2)^i (1 + q^{-1})^i (1 - q^{-1})^{p-i}}{[\alpha(1 - q^{-1}) + (T/2)(1 + q^{-1})]^p}$$

3.1. Definition. \tilde{d}_j is computed as

$$\tilde{d}_j = \frac{d_j}{T} = k_j + \frac{\Delta_j}{T}. \quad (10)$$

The approximated DT identification model described in (7) is rewritten as follows:

$$\bar{x}_0(k) = \varphi^T(k, \hat{d}) \hat{\theta} + v(k), \quad (11)$$

where $\hat{\theta}$ and \hat{d} are described as

$$\hat{\theta}^T = [\hat{f}^T, \hat{g}_1^T, \dots, \hat{g}_r^T], \quad (12)$$

with

$$\begin{aligned} \hat{f}^T &= [\hat{f}_1, \dots, \hat{f}_p], \\ \hat{g}_j^T &= [\hat{g}_{j1}, \dots, \hat{g}_{jq_j}], \\ \hat{d}^T &= [\hat{d}_1, \dots, \hat{d}_r]. \end{aligned} \quad (13)$$

Consider the following prediction error denoted by

$$e(k) = \bar{x}_0(k) - \hat{\theta} \varphi^T(k, \hat{d}). \quad (14)$$

$\varphi(k, \hat{d})$ is defined as

$$\varphi^T(k, \hat{d}) = \begin{bmatrix} -\varphi_{\bar{x}}^T(k) \\ \varphi_{\bar{u}_1}^T(k - \tilde{d}_1) \\ \varphi_{\bar{u}_r}^T(k - \tilde{d}_r) \\ \vdots \\ \varphi_{\bar{u}_r}^T(k - \tilde{d}_r) \end{bmatrix}, \quad (15)$$

with

$$\begin{aligned} \varphi_{\bar{x}}^T(k) &= [\xi_{1\bar{x}}(k), \dots, \xi_{p\bar{x}}(k)], \\ \varphi_{\bar{u}_j}^T(k - \tilde{d}_j) &= [\xi_{(p-q_j+1)\bar{u}_j}(k - \tilde{d}_j), \dots, \xi_{p\bar{u}_j}(k - \tilde{d}_j)]. \end{aligned} \quad (16)$$

We consider (17) which describe the generalized vector:

$$\hat{\Theta}^T = [\hat{\theta}^T, \hat{d}^T]^T. \quad (17)$$

The following criteria is adopted:

$$V(\hat{\Theta}) = \frac{1}{N - h_s} \sum_{k=h_s+1}^N \frac{1}{2} e^2(k). \quad (18)$$

$(N - h_s)$ is the data length such that $h_s \geq \max(\tilde{d}_1, \dots, \tilde{d}_r)$.

The generalized observation vector is the following:

$$\Phi(k, \hat{\Theta}) = \frac{-\partial e(k)}{\partial \hat{\Theta}} = \left[\varphi(k, \hat{d}), \frac{-\partial e(k)}{\partial \hat{d}} \right]. \quad (19)$$

The gradient of the error is computed as in [30]:

$$\begin{aligned}
\frac{\partial \varepsilon(k)}{\partial \hat{d}_j} &= - \sum_{i=1}^{q_j} \hat{g}_{ji} \frac{\partial}{\partial d_j} \bar{u}_j(p-q_j+i)(k-\hat{d}_j)|_{t=kT} = - \sum_{i=1}^{q_j} \hat{g}_{ji} \frac{d}{dt} \bar{u}_j(p-q_j+i)(k-\hat{d}_j)|_{t=kT} \\
&= - \sum_{i=1}^{q_j} \hat{g}_{ji} A_d(q^{-1})(T/2)^{p-q_j+i-1} \times (1-q^{-1})^{p-q_j+i-1} (1-q^{-1})^{p_j+i-1} \bar{u}_j(p-q_j+i)(k-\hat{d}_j) \\
&= - \sum_{i=1}^{q_j} \hat{g}_{ji} \bar{u}_j(p-q_j+i-1)(k-\hat{d}_j).
\end{aligned} \tag{20}$$

Finally, we obtain

$$\Phi(k, \hat{\Theta}) = \frac{-\partial e(k)}{\partial \hat{\Theta}} = \begin{bmatrix} -\varphi_x^T(k) \\ \varphi_{u_1}^T(k-\tilde{d}_1) \\ \vdots \\ \varphi_{u_r}^T(k-\tilde{d}_r) \\ -\sum_{i=1}^{q_1} \hat{g}_{1i} \bar{u}_j(p-q_j+i-1)(k-\tilde{d}_j) \\ \vdots \\ -\sum_{i=1}^{q_r} \hat{g}_{ri} \bar{u}_j(p-q_r+i-1)(k-\tilde{d}_r) \end{bmatrix}. \tag{21}$$

The following gradient algorithm is used:

$$\hat{\Theta}^{(i+1)} = \hat{\Theta}^{(i)} - \gamma^{(i)} V'(\hat{\Theta}^{(i)}). \tag{22}$$

So,

$$V'(\hat{\Theta}) = \frac{\partial V(\hat{\Theta})}{\partial \hat{\Theta}} = \frac{\partial(1/N - h_s \sum_{k=h_s+1}^N 1/2e^2(k))}{\partial \hat{\Theta}} = \frac{1}{N-h_s} \sum_{k=h_s+1}^N \frac{\partial e(k)}{\partial \hat{\Theta}} e(k). \tag{23}$$

(24) can be derived from (23):

$$V'(\hat{\Theta}^{(i)}) = \frac{1}{N-h_s} \sum_{k=h_s+1}^N e(k) \Phi(k, \hat{\Theta}). \tag{24}$$

$$\text{Compute } \hat{\Theta}^{(i+1)} = \hat{\Theta}^{(i)} - \gamma^{(i)} V'(\hat{\Theta}^{(i)})$$

Step 4: stopping criteria

If $\hat{\Theta}^{(i+1)}$ has converged or the maximum number of iterations is reached, then stop, else increase i by 1 such that $i = i + 1$ and go back to Step 2

End

4. Summary of Proposed Algorithm

The above approach can be summarized by the following step-by-step procedure:

Step 1: initialization

- (1) Let $i = 0$
- (2) Set the initial value of the tuning parameters $\gamma^{(0)}$, and the initial value of the cost function $V^{(0)}$

Step 2: gradient computation

Compute the gradient $V'(\hat{\Theta}^{(i)})$ using (24)

Step 3: gradient optimization algorithm

5. Convergence Analysis

Theorem 1. The following condition should be verified:

$$0 < \gamma^{(i)} < \frac{2}{N-h_s} \times \frac{1}{\sum_{h=h_s+1}^N \Phi(k, \hat{\Theta}) \Phi^T(k, \hat{\Theta})}. \tag{25}$$

Proof. Replacing $V'(\hat{\Theta}^{(i)})$ in (22) with its expression (29), we obtain

$$\widehat{\Theta}^{(i+1)} = \widehat{\Theta}^{(i)} + \frac{\gamma^{(i)}}{N - h_s} \sum_{k=h_s+1}^N e(k) \Phi(k, \widehat{\Theta}). \quad (26)$$

Such that from (14), the prediction error is given as

$$e(k) = \bar{x}_0(k) - \Phi^T(k, \widehat{\Theta}) \widehat{\Theta}. \quad (27)$$

By replacing $e(k)$ in (26) with its expression given in (27), we lead to

$$\widehat{\Theta}^{(i+1)} = \widehat{\Theta}^{(i)} + \frac{\gamma^{(i)}}{N - h_s} \sum_{k=h_s+1}^N \Phi(k, \widehat{\Theta}) (\bar{x}_0(k) - \Phi^T(k, \widehat{\Theta}) \widehat{\Theta}). \quad (28)$$

Let denotes $\widehat{\Theta}_{opt}$ the optimal value of $\widehat{\Theta}$. We can then rewrite (32) as

$$\widehat{\Theta}_{op}^{(i+1)} = \widehat{\Theta}_{op}^{(i)} + \frac{\gamma^{(i)}}{N - h_s} \sum_{k=h_s+1}^N \Phi(k, \widehat{\Theta}) (\bar{x}_0(k) - \Phi^T(k, \widehat{\Theta}) \widehat{\Theta}). \quad (29)$$

$$\widehat{\Theta}^{(i+1)} = \widehat{\Theta}^{(i)} + \frac{\gamma^{(i)}}{N - h_s} \sum_{k=h_s+1}^N \Phi(k, \widehat{\Theta}) \Phi^T(k, \widehat{\Theta}) (\widehat{\Theta}_{opt} - \widehat{\Theta}^{(i)}). \quad (34)$$

Now, let define

$$\chi^{(i)} = \widehat{\Theta}_{opt} - \widehat{\Theta}^{(i)}. \quad (35)$$

Then,

$$\widehat{\Theta}^{(i+1)} - \widehat{\Theta}_{opt} = \widehat{\Theta}^{(i)} - \widehat{\Theta}_{opt} + \frac{\gamma^{(i)}}{N - h_s} \sum_{k=h_s+1}^N \Phi(k, \widehat{\Theta}) \Phi^T(k, \widehat{\Theta}) \chi^{(i)}. \quad (37)$$

Taking into account the expression of $\chi^{(i)}$ in (35), it follows:

$$-\chi^{(i+1)} = -\chi^{(i)} + \frac{\gamma^{(i)}}{N - h_s} \sum_{k=h_s+1}^N \Phi(k, \widehat{\Theta}) \Phi^T(k, \widehat{\Theta}) \chi^{(i)}. \quad (38)$$

Multiplying (38) by (-1) , we obtain

$$\chi^{(i+1)} = \chi^{(i)} - \frac{\gamma^{(i)}}{N - h_s} \sum_{k=h_s+1}^N \Phi(k, \widehat{\Theta}) \Phi^T(k, \widehat{\Theta}) \chi^{(i)}. \quad (39)$$

We get

$$\chi^{(i+1)} = \chi^{(i)} \left(1 - \frac{\gamma^{(i)}}{N - h_s} \sum_{k=h_s+1}^N \Phi(k, \widehat{\Theta}) \Phi^T(k, \widehat{\Theta}) \right). \quad (40)$$

Or at the convergence, we have

$$\widehat{\Theta}^{(i+1)} = \widehat{\Theta}^{(i)}. \quad (30)$$

Consequently, we obtain

$$\frac{\gamma^{(i)}}{N - h_s} \sum_{k=h_s+1}^N (\bar{x}_0(k) - \Phi^T(k, \widehat{\Theta}) \widehat{\Theta}_{opt}) \Phi(k, \widehat{\Theta}) = 0. \quad (31)$$

It derives then

$$(\bar{x}_0(k) - \Phi^T(k, \widehat{\Theta}) \widehat{\Theta}_{opt}) \Phi(k, \widehat{\Theta}) = 0. \quad (32)$$

Then,

$$\Phi(k, \widehat{\Theta}) \bar{x}_0(k) = \Phi(k, \widehat{\Theta}) \Phi^T(k, \widehat{\Theta}) \widehat{\Theta}_{opt}. \quad (33)$$

So, we get

$$\widehat{\Theta}^{(i+1)} = \widehat{\Theta}^{(i)} + \frac{\gamma^{(i)}}{N - h_s} \sum_{k=h_s+1}^N \Phi(k, \widehat{\Theta}) \Phi^T(k, \widehat{\Theta}) \chi^{(i)}. \quad (36)$$

We subtract the term $\widehat{\Theta}_{opt}$ from (36):

The condition described in (41) should be verified to guarantee the convergence:

$$\left\| 1 - \frac{\gamma^{(i)}}{N - h_s} \sum_{k=h_s+1}^N \Phi(k, \widehat{\Theta}) \Phi^T(k, \widehat{\Theta}) \right\|_2 < 1. \quad (41)$$

We obtain, finally,

$$0 < \gamma^{(i)} < \frac{2}{N - h_s} \times \frac{1}{\sum_{k=h_s+1}^N \Phi(k, \widehat{\Theta}) \Phi^T(k, \widehat{\Theta})}. \quad (42)$$

End of proof. \square

Theorem 2. The following theorem is verified:

$$E[(\widehat{\Theta} - \Theta)(\widehat{\Theta} - \Theta)^T] = \left(\frac{1}{N - h_s} \right)^2 \sum_{k=h_s+1}^N \Phi(k, \Theta) \Phi^T(k, \Theta) \delta^2. \quad (43)$$

Proof. By the use of the first-order Taylor series expansion around the real parameter of $\Theta^{(i)}$, we obtain

$$V'(\hat{\Theta}^{(i)}) = V'(\Theta^{(i)}) + V''(\Theta^{(i)}) \times [\hat{\Theta}^{(i)} - \Theta^{(i)}]. \quad (44)$$

The proposed algorithm estimates $\Theta^{(i)}$ satisfying $V'(\hat{\Theta}^{(i)}) = 0$, so it can be derived from (44) that

$$(\hat{\Theta}^{(i)} - \Theta^{(i)}) (\hat{\Theta}^{(i)} - \Theta^{(i)})^T = (V''(\Theta^{(i)}))^{-1} V'(\Theta^{(i)}) [V'(\Theta^{(i)})]^T [(V''(\Theta^{(i)}))^{-1}]^T. \quad (45)$$

It follows:

$$E \left\{ (\hat{\Theta}^{(i)} - \Theta^{(i)}) (\hat{\Theta}^{(i)} - \Theta^{(i)})^T \right\} = E \left\{ (V''(\Theta^{(i)}))^{-1} V'(\Theta^{(i)}) [V'(\Theta^{(i)})]^T [(V''(\Theta^{(i)}))^{-1}]^T \right\}. \quad (46)$$

Or:

$$V''(\Theta^{(i)}) = e(k)e''(k) - e'(k)\Phi(k, \Theta). \quad (47)$$

Then,

$$V''(\Theta^{(i)}) = e(k)e''(k) + \Phi^T(k, \Theta)\Phi(k, \Theta). \quad (48)$$

Using the small residual algorithms, the term $e(k)e''(k)$ can be neglected:

$$V''(\Theta^{(i)}) \approx \Phi^T(k, \Theta)\Phi(k, \Theta). \quad (49)$$

The mean value of $V'(\Theta)[V'(\Theta)]^T$ permits to get

$$E[V'(\Theta)[V'(\Theta)]^T] = \left(\frac{1}{N - h_s} \right)^2 E \left\{ \sum_{k=h_s+1}^N \Phi(k, \hat{\Theta})\Phi^T(k, \hat{\Theta}) \sum_{k=h_s+1}^N e(k)e^T(k) \right\}. \quad (50)$$

Or,

$$E \left\{ \sum_{k=h_s+1}^N e(k)e^T(k) \right\} = \delta^2. \quad (51)$$

Finally, (52) can be simplified as follows:

$$E[(\hat{\Theta} - \Theta)(\hat{\Theta} - \Theta)^T] = \left(\frac{1}{N - h_s} \right)^2 \sum_{k=h_s+1}^N \Phi(k, \hat{\Theta})\Phi^T(k, \hat{\Theta})\delta^2. \quad (52)$$

□

6. Numerical Example

Two examples are presented to confirm the performance of the developed scheme.

The vector of parameters to be estimated is the following:

$$\Theta = [f_1, f_2, g_{11}, g_{12}, g_{21}, g_{22}, d_1, d_2]. \quad (53)$$

6.1. Example 1. The first simulated example is described as

$$\ddot{x}(t) + f_1\dot{x}(t) + f_2x(t) = g_{11}\dot{u}_1(t - d_1) + g_{12}u_1(t - d_1) + g_{21}\dot{u}_2(t - d_2) + g_{22}u_2(t - d_2). \quad (54)$$

In the Laplace domain, the above equation becomes

$$x(s) = \frac{e^{-d_1s}}{s+1}u_1(s) + \frac{2e^{-d_2s}}{s+2}u_2(s) = \frac{e^{-d_1s}(s+2)u_1(s) + e^{-d_2s}(2s+2)u_2(s)}{s^2 + 3s + 2}. \quad (55)$$

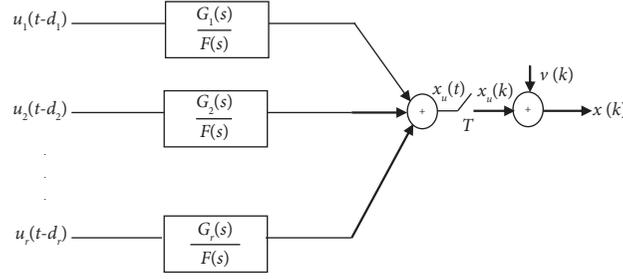


FIGURE 1: The considered process.

TABLE 1: Simulation results with noise-free output.

	True	Estimates
f_1	3	2.9818 ± 0.0182
f_2	2	1.9844 ± 0.0156
g_{11}	1	0.9912 ± 0.0088
g_{12}	2	1.9876 ± 0.0124
g_{21}	2	1.9941 ± 0.0059
g_{22}	2	1.9830 ± 0.017
d_1	8.83	8.8301 ± 0.0001
d_2	2.32	2.3255 ± 0.0055

TABLE 2: Results simulation with SNR = 20dB.

	True	Estimates
f_1	3	2.9813 ± 0.0187
f_2	2	1.9840 ± 0.016
g_{11}	1	0.9931 ± 0.0069
g_{12}	2	1.9887 ± 0.0113
g_{21}	2	1.9955 ± 0.0045
g_{22}	2	1.9810 ± 0.019
d_1	8.83	8.8276 ± 0.0024
d_2	2.32	2.3240 ± 0.0032

The real parameters are the following:

$$\begin{aligned} f_1 = 3, f_2 = 2, g_{11} = 1, g_{12} = 2, g_{21} = 2, g_{22} = 2. \\ d_1 = 8.83s, d_2 = 2.32s. \end{aligned} \quad (56)$$

6.1.1. Simulation Results. We obtain the following results illustrated in Table 1.

With a SNR = 20dB, the parameters estimated are shown in Table 2.

6.1.2. Model Validation. Validations of the obtained model by using the gradient algorithm are plotted, respectively, in the following Figures 2 and 3.

6.1.3. Bode Diagrams. The bode diagrams are plotted, respectively, in Figures 4 and 5.

The accuracy of the developed prediction error method is gradually improved in the frequency domain.

6.1.4. Computational Time. A comparison is done in Table 3 with the computational time in the second simulation case. For comparison, the identification algorithm GSEPNIV for the global separable nonlinear instrumental variable given in reference [23] is also adopted to estimate the above stable-type model parameters defined by (53) by taking the same initial parameter setting.

The computational time is measured using the tic and toc Matlab function.

Table 3 shows that the GSEPNIV method needs a long computational time. However, it can be seen that the computational time is smaller by using the proposed prediction error method.

6.1.5. Noise Effects. Table 4 addresses another comparison between the proposed method and the GSEPNIV method developed in [23] in terms of robustness with regard to noise in the case of a noisy output.

Judging from the simulation results tabulated in Table 4 concerning computational error values for each parameter, it can be seen that the proposed method outperforms the GSEPNIV method.

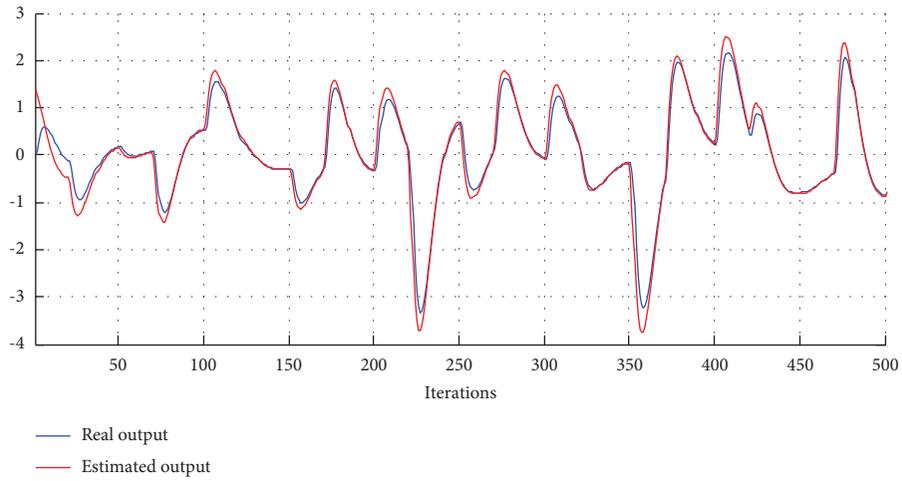


FIGURE 2: Model validation of the proposed scheme with noise-free output.

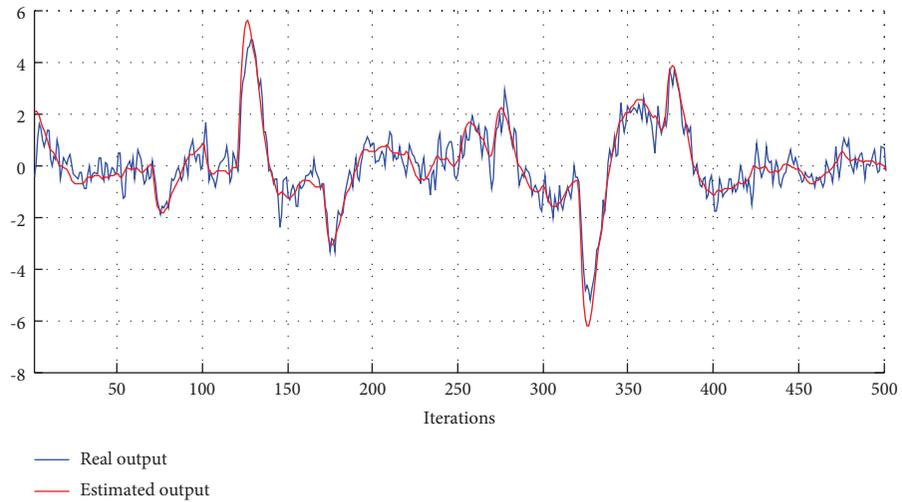


FIGURE 3: Model validation in the case of SNR = 20dB.

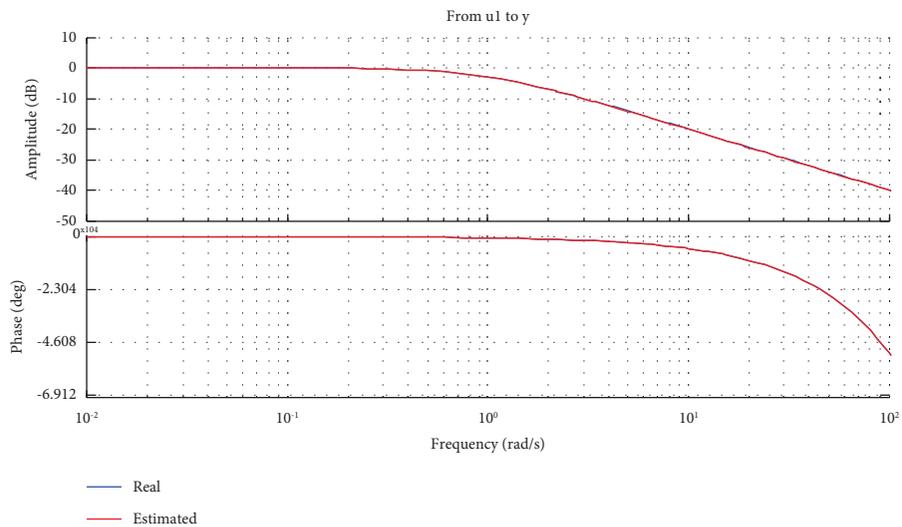


FIGURE 4: Bode diagrams of the true system model with the estimated model: noise-free case.

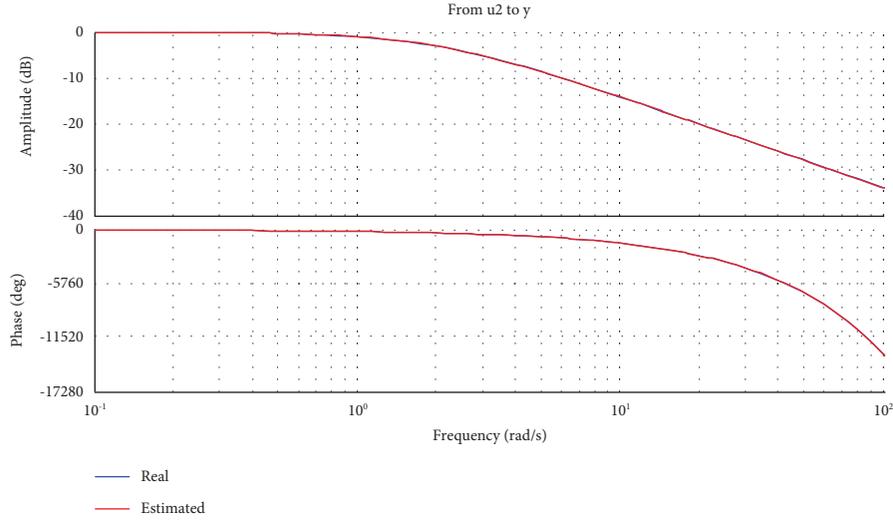


FIGURE 5: Bode diagrams of the true system model with the estimated model SNR = 20dB.

6.2. Example 2. We consider the following process described in Figure 6, where u_1 is the steam flow, u_2 is the water flow u_2 , and x is the water temperature.

The equation of the above process is

$$\ddot{x}(t) + f_1\dot{x}(t) + f_2x(t) = g_{11}\dot{u}_1(t - d_1) + g_{12}u_1(t - d_1) + g_{21}\dot{u}_2(t - d_2) + g_{22}u_2(t - d_2). \quad (57)$$

Then,

$$x(s) = \frac{e^{-d_1s}}{s + 2}u_1(s) + \frac{e^{-\tau_2s}}{s + 2}u_2(s) = \frac{e^{-\tau_1s}(g_{11}s + g_{12})u_1(s) + e^{-\tau_2s}(g_{21}s + g_{22})u_2(s)}{s^2 + f_1s + f_2}. \quad (58)$$

Such that

$$f_1 = 4, f_2 = 4, g_{11} = 1, g_{12} = 2, g_{21} = 1, g_{22} = 2, d_1 = 6.84s, d_2 = 1.87s. \quad (59)$$

6.2.1. Simulation Results. The simulation results in the case of a noise free or a noisy output are tabulated in Tables 5 and 6.

From Tables 5 and 6, we can see that the proposed method gives optimal estimates in the noise free and the noisy cases [31, 32].

6.2.2. Computational Time. To further demonstrate the effectiveness of the proposed identification method, a comparison is presented in Table 7 with the computational

TABLE 3: Computational time SNR = 20dB.

	GSEPNIV	Proposed method
Computational time (sec)	5.730758	1.756012

time, respectively, of the GSEPNIV and the proposed method in the case of a noisy output.

It can be seen that the proposed prediction error method outperforms the GSEPNIV method in terms of the computational time.

7. Discussions

From Tables 1–7 and Figures 2–5, we can draw the following features:

TABLE 4: Parameters estimates: simulation 2.

	True	GSEPNIV	Proposed method
f_1	3	3.0646 ± 0.0646	2.9813 ± 0.0187
f_2	2	2.0773 ± 0.0773	1.9840 ± 0.016
g_{11}	1	0.9804 ± 0.0196	0.9931 ± 0.0069
g_{12}	2	2.0759 ± 0.0759	1.9887 ± 0.0113
g_{21}	2	2.0269 ± 0.0269	1.9955 ± 0.0045
g_{22}	2	2.0758 ± 0.0758	1.9810 ± 0.019
d_1	8.84	8.8671 ± 0.0271	8.8276 ± 0.0024
d_2	2.32	2.3315 ± 0.0115	2.3240 ± 0.0032

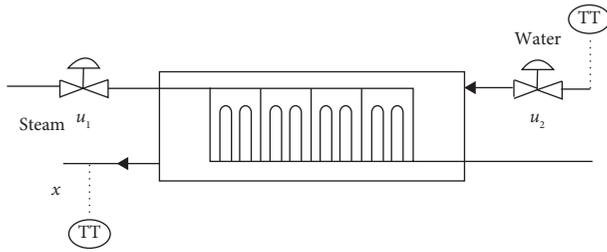


FIGURE 6: The steam-water heat exchanger.

TABLE 5: Simulation results with a noise-free output.

	True	Estimates
f_1	4	3.9954 ± 0.0046
f_1	4	4.0003 ± 0.0003
g_{11}	1	1.0043 ± 0.0043
g_{12}	2	1.9747 ± 0.0253
g_{21}	1	0.9838 ± 0.0162
g_{22}	2	1.9938 ± 0.0062
d_1	6.84	6.8314 ± 0.0086
d_2	1.87	1.6886 ± 0.0014

TABLE 6: Results simulation with SNR = 20dB.

	True	Estimates
f_1	4	4.0009 ± 0.0009
f_2	4	3.9976 ± 0.0024
g_{11}	1	0.9977 ± 0.0023
g_{12}	2	2.0007 ± 0.0007
g_{21}	1	0.9942 ± 0.0058
g_{22}	2	2.0017 ± 0.0017
d_1	6.84	6.8392 ± 0.0008
d_2	1.87	1.8735 ± 0.0035

TABLE 7: Computational time.

	GSEPNIV	Proposed method
Computational time (sec)	4.335081	1.713794

- (1) The proposed method insures a high accuracy in the noise-free case or in the noisy case
- (2) According to Figures 2–3, we can see clearly that the estimated outputs track fast the true outputs with acceptable speed

- (3) The simulation results tabulated in Tables 1–7 confirm the robustness of the method proposed in this paper in the case of noise-free output likewise in the case of noisy output
- (4) Considering the computational time, the proposed method converges faster than the GSEPNIV method proposed in [23]

8. Conclusion

In the present research work, we have addressed the problem of simultaneous identification of the MISO CT system with multiple unknown time delays from sampled data. Indeed, we have developed a new scheme based on a gradient algorithm. The proposed algorithm is then simulated on two examples to confirm its performance.

By using the presented scheme, the identification of the considered system operating on the closed loop is under investigation.

Data Availability

The data used to support the findings of this study are available from the corresponding author upon request.

Conflicts of Interest

The authors declare that they have no conflicts of interest.

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